THEORY AND METHODS OF PANEL DATA MODELS
WITH INTERACTIVE EFFECTS

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This paper considers the maximum likelihood estimation of panel data models with interactive effects. Motivated by applications in economics and other social sciences, a notable feature of the model is that the explanatory variables are correlated with the unobserved effects. The usual within-group estimator is inconsistent. Existing methods for consistent estimation are either designed for panel data with short time periods or are less efficient. The maximum likelihood estimator has desirable properties and is easy to implement, as illustrated by the Monte Carlo simulations. This paper develops the inferential theory for the maximum likelihood estimator, including consistency, rate of convergence and the limiting distributions. We further extend the model to include time-invariant regressors and common regressors (cross-section invariant). The regression coefficients for the time-invariant regressors are time-varying, and the coefficients for the common regressors are cross-sectionally varying.

1. Introduction. This paper studies the following panel data models with unobservable interactive effects:

\[ y_{it} = \alpha_i + x_{it1}\beta_1 + \cdots + x_{itK}\beta_K + \lambda_i'f_t + \epsilon_{it} \]

where \( y_{it} \) is the dependent variable; \( x_{it} = (x_{it1}, \ldots, x_{itK}) \) is a row vector of explanatory variables; \( \alpha_i \) is an intercept; the term \( \lambda_i'f_t + \epsilon_{it} \) is unobservable and has a factor structure, \( \lambda_i \) is an \( r \times 1 \) vector of factor loadings, \( f_t \) is a vector of factors, and \( \epsilon_{it} \) is the idiosyncratic error. The interactive effects (\( \lambda_i'f_t \)) generalize the usual additive individual and time effects, for example, if \( \lambda_i \equiv 1 \), then \( \alpha_i + \lambda_i'f_t = \alpha_i + f_t \).

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A key feature of the model is that the regressors \( x_{it} \) are allowed to be correlated with \((\alpha_i, \lambda_i, f_t)\). This situation is commonly encountered in economics and other social sciences, in which some of the regressors \( x_{it} \) are decision variables that are influenced by the unobserved individual heterogeneities. The practical relevance of the model will be further discussed below. The objective of this paper is to obtain consistent and efficient estimation of \( \beta \) in the presence of correlations between the regressors and the factor loadings and factors.

The usual pooled least squares estimator or even the within-group estimator is inconsistent for \( \beta \). One method to obtain a consistent estimator is to treat \((\alpha_i, \lambda_i, f_t)\) as parameters and estimate them jointly with \( \beta \). The idea is “controlling through estimating” (controlling the effects by estimating them). This is the approach used by [8], [23] and [31]. While there are some advantages, an undesirable consequence of this approach is the incidental parameters problem. There are too many parameters being estimated, and the incidental parameters bias arises (Neyman and Scott, 1948). [1], [2] and [17] consider the generalized method of moments (GMM) method. The GMM method is based on a nonlinear transformation known as quasi-differencing that eliminates the factor errors. Quasi-differencing increases the nonlinearity of the model especially with more than one factor. The GMM method works well with a small \( T \). When \( T \) is large, the number of moment equations will be large and the so called many-moment bias arises. [27] considers an alternative method by augmenting the model with additional regressors \( \bar{y}_t \) and \( \bar{x}_t \), which are the cross-sectional averages of \( y_{it} \) and \( x_{it} \). These averages provide an estimate for \( f_t \). A further approach to controlling the correlation between the regressors and factor errors is to use the Mundlak-Chamberlain projection ([24] and [14]). The latter method projects \( \alpha_i \) and \( \lambda_i \) onto the regressors such that \( \lambda_i = c_0 + c_1 x_{i1} + \cdots + c_T x_{iT} + \eta_i \), where \( c_s \) (\( s = 0, 1, \ldots, T \)) are parameters to be estimated and \( \eta_i \) is the projection residual (a similar projection is done for \( \alpha_i \)). The projection residuals are uncorrelated with the regressors so that a variety of approaches can be used to estimate the model. This framework is designed for small \( T \), and is studied by [9].

In this paper we consider the pseudo-Gaussian maximum likelihood method under large \( N \) and large \( T \). The theory does not depend on normality. In view of the importance of the MLE in the statistical literature, it is of both practical and theoretical interest to examine the MLE in this context. We develop a rigorous theory for the MLE. We show that there is no incidental parameters bias despite large \( N \) and large \( T \).

We allow time-invariant regressors such as education, race and gender in the model. The corresponding regression coefficients are time-dependent.
Similarly, we allow common regressors, which do not vary across individuals, such as prices and policy variables. The corresponding regression coefficients are individual-dependent so that individuals respond differently to policy or price changes. In our view, this is a sensible way to incorporate time-invariant and common regressors. For example, wages associated with education and with gender are more likely to change over time rather than remain constant. In our analysis, time invariant regressors are treated as the components of \( \lambda_i \) that are observable, and common regressors as the components of \( f_t \) that are observable. This view fits naturally into the factor framework in which part of the factor loadings and factors are observable, and the maximum likelihood method imposes the corresponding loadings and factors at their observed values.

While the theoretical analysis of MLE is demanding, the limiting distributions of the MLE are simple and have intuitive interpretations. The computation is also easy and can be implemented by adapting the ECM (expectation and constrained maximization) of [22]. In addition, the maximum likelihood method allows restrictions to be imposed on \( \lambda_i \) or on \( f_t \) to achieve more efficient estimation. These restrictions can take the form of known values, being either zeros, or other fixed values. Part of the rigorous analysis includes setting up the constrained maximization as a Lagrange multiplier problem. This approach provides insight on which kinds of restrictions are binding and which are not, shedding light on efficiency gain resulting from the restrictions.

Panel data models with interactive effects have wide applicability in economics. In macroeconomics, for example, \( y_{it} \) can be the output growth rate for country \( i \) in year \( t \); \( x_{it} \) represents production inputs, and \( f_t \) is a vector of common shocks (technological progress, financial crises); the common shocks have heterogenous impacts across countries through the different factor loadings \( \lambda_i \); \( e_{it} \) represents the country-specific unmeasured growth rates. In microeconomics, and especially in earnings studies, \( y_{it} \) is the wage rate for individual \( i \) for period \( t \) (or for cohort \( t \)), \( x_{it} \) is a vector of observable characteristics such as marital status and experience; \( \lambda_i \) is a vector of unobservable individual traits such as ability, perseverance, motivation and dedication; the payoff to these individual traits is not constant over time, but time varying through \( f_t \); and \( e_{it} \) is idiosyncratic variations in the wage rates. In finance, \( y_{it} \) is stock \( i \)'s return in period \( t \), \( x_{it} \) is a vector of observable factors, \( f_t \) is a vector of unobservable common factors (systematic risks) and \( \lambda_i \) is the exposure to the risks; \( e_{it} \) is the idiosyncratic returns. Factor error structures are also used as a flexible trend modeling as in [20]. Most of panel data analysis assumes cross-sectional independence, e.g., [6], [12],
and [18]. The factor structure is also capable of capturing the cross-sectional
dependence arising from the common shocks $f_t$.

Throughout the paper, the norm of a vector or matrix is that of Frobenius,
i.e., $\|A\| = [\text{tr}(A^tA)]^{1/2}$ for matrix $A$; diag($A$) is a column vector consisting
of the diagonal elements of $A$ when $A$ is matrix, but diag($A$) represents
a diagonal matrix when $A$ is a vector. In addition, we use $\dot{v}_t$ to denote
$v_t - \frac{1}{T} \sum_{t=1}^{T} v_t$ for any column vector $v_t$ and $M_{ww}$ to denote $\frac{1}{T} \sum_{t=1}^{T} \dot{w}_t \dot{v}_t^t$ for
any vectors $w_t, v_t$.

2. A common shock model. In the common-shock model, we assume
that both $y_{it}$ and $x_{it}$ are impacted by the common shocks $f_t$ so the model
takes the form

$$
\begin{align*}
y_{it} &= \alpha_i + x_{it1}\beta_1 + x_{it2}\beta_2 + \cdots + x_{itK}\beta_K + \lambda_i^t f_t + e_{it} \\
x_{itk} &= \mu_{ik} + \gamma_{ik}^t f_t + v_{itk}
\end{align*}
$$

for $k = 1, 2, ..., K$. In across-country output studies, for example, output $y_{it}$
and inputs $x_{it}$ (labor and capital) are both affected by the common shocks.

The parameter of interest is $\beta = (\beta_1, ..., \beta_K)'$. We also estimate $\alpha_i, \lambda_i, \mu_{ik}$
and $\gamma_{ik} (k = 1, 2, ..., K)$. By treating the latter as parameters, we also allow
arbitrary correlations between $(\alpha_i, \lambda_i)$ and $(\mu_{ik}, \gamma_{ik})$. Although we also treat
$f_t$ as fixed parameters, there is no need to estimate the individual $f_t$, but
only the sample covariance of $f_t$. This is an advantage of the maximum
likelihood method, which eliminates the incidental parameters problem in
the time dimension. This kind of the maximum likelihood method was used
for pure factor models in [3], [4], and [10]. By symmetry, we could also
estimate individuals $f_t$, but then we only estimate the sample covariance of
the factor loadings. The idea is that we do not simultaneously estimate the
factor loadings and the factors $f_t$ (which would be the case for the principal
components method). This reduces the number of parameters considerably.
If $N$ is much smaller than $T$ ($N \ll T$), treating factor loadings as parameters
is preferable since there are fewer number of parameters.

Because of the correlation between the regressors and regression errors
in the $y$ equation, the $y$ and $x$ equations form a simultaneous equation
system; the MLE jointly estimates the parameters in both equations. The
joint estimation avoids the Mundlak-Chamberlain projection and thus is
applicable for large $N$ and large $T$.

Throughout the paper, we assume the number of factors $r$ is fixed and
known. If not, the information criterions developed by [11] can be used to
determine it. So $\lambda_i$ and $f_t$ are $r \times 1$ vectors. Let $x_{it} = (x_{it1}, x_{it2}, \cdots, x_{itK}), \gamma_{ix} =
(\gamma_{i1}, \gamma_{i2}, \cdots, \gamma_{iK}), v_{itx} = (v_{it1}, v_{it2}, \cdots, v_{itK})'$ and $\mu_{ix} = (\mu_{i1}, \mu_{i2}, \cdots, \mu_{iK})'$. 
The second equation of (2.1) can be written in matrix form as
\[ x'_{it} = \mu_{ix} + \gamma'_{ix} f_t + v_{it} \]

Further let \( \Gamma_i = (\lambda_i, \gamma_{ix}) \), \( z_{it} = (y_{it}, x_{it})' \), \( \varepsilon_{it} = (e_{it}, v'_{itx})' \), \( \mu_i = (\alpha_i, \mu'_{ix})' \).

Then model (2.1) can be written as
\[ \begin{bmatrix} 1 \\ 0 \end{bmatrix} z_{it} = \mu_i + \Gamma_i f_t + \varepsilon_{it} \]

Let \( B \) denote the coefficient matrix of \( z_{it} \) in the preceding equation. Let \( z_t = (z'_{1t}, z'_{2t}, \cdots, z'_{Nt})' \), \( \Gamma = (\Gamma_1, \Gamma_2, \cdots, \Gamma_N)' \), \( \varepsilon_t = (\varepsilon'_{1t}, \varepsilon'_{2t}, \cdots, \varepsilon'_{Nt})' \) and \( \mu = (\mu'_{1t}, \mu'_{2t}, \cdots, \mu'_{Nt})' \). Stacking the equations over \( i \), we have
\[ (I_N \otimes B) z_t = \mu + \Gamma f_t + \varepsilon_t \]

To analyze this model, we impose the following assumptions.

2.1. Assumptions. **Assumption A:** The \( f_t \) is a sequence of constants. Let \( M_{ff} = T^{-1} \sum_{t=1}^{T} \hat{f}_t f_t' \), where \( \hat{f}_t = f_t - \frac{1}{T} \sum_{t=1}^{T} f_t \). We assume that \( M_{ff} = \lim_{T \to \infty} M_{ff} \) is a strictly positive definite matrix.

**Remark 2.1.** The non-randomness assumption for \( f_t \) is not crucial. In fact, \( f_t \) can be a sequence of random variables such that \( E(||f_t||^4) \leq C < \infty \) uniformly in \( t \) and \( f_t \) is independent of \( \varepsilon_s \) for all \( s \). The fixed \( f_t \) assumption conforms with the usual fixed effects assumption in panel data literature and, in certain sense, is more general than random \( f_t \).

**Assumption B:** The idiosyncratic error terms \( \varepsilon_{it} = (e_{it}, v'_{itx})' \) are assumed such that

B.1 The \( e_{it} \) is independent and identically distributed over \( t \) and uncorrelated over \( i \) with \( E(e_{it}) = 0 \) and \( E(e_{it}^2) \leq \infty \) for all \( i = 1, \cdots, N \) and \( t = 1, \cdots, T \). Let \( \Sigma_{eit} \) denote the variance of \( e_{it} \).

B.2 \( v_{itx} \) is also independent and identically distributed over \( t \) and uncorrelated over \( i \) with \( E(v_{itx}) = 0 \) and \( E(||v_{itx}||^4) \leq \infty \) for all \( i = 1, \cdots, N \) and \( t = 1, \cdots, T \). We use \( \Sigma_{vtx} \) to denote the variance matrix of \( v_{itx} \).

B.3 \( e_{it} \) is independent of \( v_{jtx} \) for all \( (i, j, t, s) \). Let \( \Sigma_{ei} \) denote the variance matrix \( e_{it} \). So we have \( \Sigma_{ei} = \text{diag}(\Sigma_{eit}, \Sigma_{eit}) \), a block-diagonal matrix.

**Remark 2.2.** Let \( \Sigma_{ee} \) denote the variance of \( \varepsilon_t = (\varepsilon'_{1t}, \cdots, \varepsilon'_{Nt})' \). Due to the uncorrelatedness of \( \varepsilon_{it} \) over \( i \), we have \( \Sigma_{ee} = \text{diag}(\Sigma_{e1}, \Sigma_{e2}, \cdots, \Sigma_{eN}) \),
a block-diagonal matrix. Assumption B is more general than the usual assumption in the factor analysis. In a traditional factor model, the variance of the idiosyncratic error terms are assumed to be a diagonal matrix. In the present setting, the variance of \( \varepsilon_t \) is a block-diagonal matrix. Even without explanatory variables, this generalization is of interest. The factor analysis literature has a long history to explore the block-diagonal idiosyncratic variance, known as multiple battery factor analysis, see [32]. The maximum likelihood estimation theory for high dimensional factor models with block diagonal covariance matrix has not been previously studied. The asymptotic theory developed in this paper not only provides a way of analyzing the coefficient \( \beta \), but also a way of analyzing the factors and loadings in the multiple battery factor models. This framework is of independent interest.

Assumption B allows cross-sectional heteroskedasticity. The maximum likelihood method will simultaneously estimate the heteroskedastic variances and other parameters. This assumption assumes the independence and homoskedasticity of the error terms over time and uncorrelatedness over the cross section. Extension to more general heteroscedasticity and correlation patterns can be considered by our method. The model with more general error covariance structure, known as approximate factor models in the sense of [15], has been extensively investigated by the recent literature, such as [11], [7], [30] among others. This literature largely focuses on the principal components method and for pure factor models without explanatory variables. The analysis of the maximum likelihood method for our model is already challenging, the extension to approximate factor models is not considered in this paper.

**Assumption C:** There exists a positive constant \( C \) sufficiently large such that

\[
C.1 \quad \|\Gamma_j\| \leq C \quad \text{for all} \quad j = 1, \cdots, N.
\]

\[
C.2 \quad C^{-1} \leq \tau_{\min}(\Sigma_{jj}) \leq \tau_{\max}(\Sigma_{jj}) \leq C \quad \text{for all} \quad j = 1, \cdots, N, \quad \text{where}
\]

\[
\tau_{\min}(\Sigma_{jj}) \quad \text{and} \quad \tau_{\max}(\Sigma_{jj}) \quad \text{denote the smallest and largest eigenvalues of the matrix} \quad \Sigma_{jj}, \quad \text{respectively.}
\]

\[
C.3 \quad \text{there exists an} \quad r \times r \quad \text{positive matrix} \quad Q \quad \text{such that} \quad Q = \lim_{N \to \infty} N^{-1} \Gamma' \Sigma_{\varepsilon\varepsilon}^{-1} \Gamma,
\]

where \( \Gamma \) is defined earlier.

**Assumption D:** The variances \( \Sigma_{ii} \) for all \( i \) and \( M_{ff} \) are estimated in a compact set, i.e. all the eigenvalues of \( \hat{\Sigma}_{ii} \) and \( \hat{M}_{ff} \) are in an interval \([C^{-1}, C]\) for a sufficiently large constant \( C \).

**Remark 2.3.** Assumption D requires that part of the estimators be estimated in a compact set. This assumption is usually made for theoretical
analysis, especially when dealing with nonlinear objective functions, e.g., [19], [25], and [33]. The objective function considered in this paper exhibits high nonlinearity.

2.2. Identification restrictions. It is a well-known result in factor analysis that the factors and loadings can only be identified up to a rotation. The models considered in this paper can be viewed as extensions of the factor models. As such they inherit the same identification problem. We show that identification conditions can be imposed on the factors and loadings without loss of generality. To see this, model (2.2) can be rewritten as

\[(I_N \otimes B)z_t = \mu + \Gamma f_t + \epsilon_t\]

(2.3)

where \(R\) is an orthogonal matrix, which we choose to be the matrix consisting of the eigenvectors of \(M_{ff}^{1/2} \Gamma' \Sigma_{ee}^{-1} \Gamma M_{ff}^{1/2}\) associated with the eigenvalues arranged in descending order. Treating \(\mu + \Gamma \bar{f}\) as the new \(\mu^*\), \(\Gamma M_{ff}^{1/2} R\) as the new \(\Gamma^*\) and \(R'M_{ff}^{-1/2}(f_t - \bar{f})\) as the new \(f_t^*\), we have

\[(I_N \otimes B)z_t = \mu^* + \Gamma^* f_t^* + \epsilon_t\]

with \(\frac{1}{T} \sum_{t=1}^{T} f_t^* = 0\), \(\frac{1}{T} \sum_{t=1}^{T} f_t^* f_t^* = I_r\) and \(\frac{1}{N} \Gamma^* \Sigma_{ee}^{-1} \Gamma^*\) being a diagonal matrix. Given the above analysis, we can impose in (2.2) the following restrictions, which we refer to as IB (Identification restrictions for Basic models).

IB1. \(M_{ff} = I_r\)

IB2. \(\frac{1}{N} \Gamma' \Sigma_{ee}^{-1} \Gamma = D,\) where \(D\) is a diagonal matrix with its diagonal elements distinct and arranged in descending order.

IB3. \(\bar{f} = \frac{1}{T} \sum_{t=1}^{T} f_t = 0.\)

Remark 2.4. The requirement that the diagonal elements of \(D\) are distinct in IB2 is not needed for the ML estimation of \(\beta\), but it is needed for the identification of factors and factor loadings. Under this requirement, the orthogonal matrix \(R\) in (2.3) can be uniquely determined up to a column sign change. This assumption does simplify the analysis for the MLE of \(\beta\).

2.3. Estimation. The objective function considered in this section is

\[\ln L = -\frac{1}{2N} \ln |\Sigma_z| - \frac{1}{2N} \text{tr} [(I_N \otimes B)M_z(I_N \otimes B')\Sigma_z^{-1}],\]

(2.4)
where $\Sigma_{zz} = \Gamma_{ff}\Gamma' + \Sigma_{ee}$ and $M_x = \frac{1}{T} \sum_{t=1}^T \dot{x}_t \dot{x}_t'$. Here $\Sigma_{zz}$ is the matrix consisting of the parameters other than $\beta$, the latter is contained in $B$; $M_x$ is the data matrix. The objective function (2.4) can be regarded as the likelihood function (omitting a constant). Note that the determinant of $I_N \otimes B$ is 1, so the Jacobian term does not depend on $B$. If $\varepsilon_t$ and $f_t$ are independent and normally distributed, the likelihood function for the observed data has the form of (2.4). Here recall that $f_t$ are fixed constants and $\varepsilon_t$ are not necessarily normal, (2.4) is a pseudo-likelihood function.

For further analysis, we partition the matrix $\Sigma_{zz}$ and $M_x$ as

$$\Sigma_{zz} = \begin{pmatrix}
\Sigma_{11}^{zz} & \Sigma_{12}^{zz} & \cdots & \Sigma_{1N}^{zz} \\
\Sigma_{21}^{zz} & \Sigma_{22}^{zz} & \cdots & \Sigma_{2N}^{zz} \\
\vdots & \vdots & & \vdots \\
\Sigma_{N1}^{zz} & \Sigma_{N2}^{zz} & \cdots & \Sigma_{NN}^{zz}
\end{pmatrix}, \quad M_x = \begin{pmatrix}
M_{11}^{zz} & M_{12}^{zz} & \cdots & M_{1N}^{zz} \\
M_{21}^{zz} & M_{22}^{zz} & \cdots & M_{2N}^{zz} \\
\vdots & \vdots & & \vdots \\
M_{N1}^{zz} & M_{N2}^{zz} & \cdots & M_{NN}^{zz}
\end{pmatrix}$$

where for any $(i, j), \Sigma_{ij}^{zz}$ and $M_{ij}^{zz}$ are both $(K+1) \times (K+1)$ matrices.

Let $\hat{\beta}, \hat{\Gamma}$ and $\hat{\Sigma}_{ee}$ denote the MLE. The first order condition for $\beta$ satisfies

$$\frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \hat{\Gamma}_{i} \hat{\Sigma}^{-1}_{ii} \{ (\dot{y}_{it} - \dot{x}_{it}\hat{\beta}) - \hat{\lambda}' \hat{G} \sum_{j=1}^N \hat{\Gamma}_{j} \hat{\Sigma}^{-1}_{jj} \{ \dot{y}_{jt} - \dot{x}_{jt}\hat{\beta} \} \} \dot{x}_{it} = 0$$

where $\hat{G} = (\hat{M}_{ff}^{-1} + \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma})^{-1}$. The first order condition for $\Gamma_j$ satisfies

$$\sum_{i=1}^N \hat{\Gamma}_{i} \hat{\Sigma}^{-1}_{ii} \left( \hat{B} M_{ij}^y \hat{B}' - \hat{\Sigma}_{ij}^y \right) = 0.$$

Post-multiplying $\hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j$ on both sides of (2.6) and then taking summation over $j$, we have

$$\sum_{i=1}^N \sum_{j=1}^N \hat{\Gamma}_{i} \hat{\Sigma}^{-1}_{ii} \left( \hat{B} M_{ii}^y \hat{B}' - \hat{\Sigma}_{ii}^y \right) \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j = 0.$$

The first order condition for $\Sigma_{ii}$ satisfies

$$\hat{B} M_{ii}^y \hat{B}' - \hat{\Sigma}_{ii}^y = \mathbb{W},$$

where $\mathbb{W}$ is a $(K+1) \times (K+1)$ matrix such that its upper-left $1 \times 1$ and lower-right $K \times K$ submatrices are both zero, but the remaining elements are undetermined. The undetermined elements correspond to the zero elements of $\Sigma_{ii}$. These first order conditions are needed for the asymptotic representation of the MLE.
2.4. Asymptotic properties of the MLE. As $N$ tends to infinity, the number of parameters goes to infinity, which makes consistency proof more difficult. Following [10], we establish the following average consistency results which serve as the basis for subsequent analysis.

**Proposition 2.1 (Consistency).** Let $\hat{\theta} = (\hat{\beta}, \hat{\Gamma}, \hat{\Sigma}_{\epsilon\epsilon})$ be the solution by maximizing (2.4). Under Assumptions A-D and the identification conditions IB, when $N,T \to \infty$, we have

$$\hat{\beta} - \beta \overset{p}{\to} 0$$

$$\frac{1}{N} \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}_i^{-1} (\hat{\Gamma}_i - \Gamma_i)' \overset{p}{\to} 0$$

$$\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \| \overset{p}{\to} 0$$

The derivation of Proposition 2.1 requires considerable work. The results of $\hat{\beta} - \beta \overset{p}{\to} 0$ and $\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \| \overset{p}{\to} 0$ can be directly derived by working with the objective function because they are free of rotational problems. To prove $\frac{1}{N} \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}_i^{-1} (\hat{\Gamma}_i - \Gamma_i)' \overset{p}{\to} 0$, we have to invoke the identification conditions. In addition, the identification condition used in this section has so-called sign problem. So the estimator $\hat{\Gamma}$ having the same signs as those of $\Gamma$ is assumed.

In order to derive the inferential theory, we need to strengthen Proposition 2.1. This result is stated in the following theorem.

**Theorem 2.1 (Convergence rate).** Under the assumptions of Proposition 2.1, we have

$$\hat{\beta} - \beta = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1})$$

$$\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii}^{-1} \| \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 = O_p(T^{-1})$$

$$\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 = O_p(T^{-1})$$

[8] considers an iterated principal components estimator for model (2.1). His derivation shows that, in the presence of heteroscedasticities over the cross section, the PC estimator for $\beta$ has a bias of order $O_p(N^{-1})$. As a
comparison, Theorem 2.1 shows that the MLE is robust to the heteroskedasticities over the cross section. So if \( N \) is fixed, the estimator in [8] is inconsistent unless there is no heteroskedasticity, but the estimator here is still consistent.

Although \( \Gamma \) and \( \Sigma_{ee} \) are not the parameters of interest and their asymptotic properties are not presented in this paper, Theorem 2.1 has implications for the limiting distributions of these parameters. Given that \( \hat{\beta} - \beta \) has a faster convergence rate, the limiting distributions of \( \text{vech}(\hat{\Gamma}_t - \Gamma_t) \) and \( \text{vech}(\Sigma_{ii} - \Sigma_{ii}) \) are not affected by the estimation of \( \beta \), and are the same as the case of without regressors. If we use \( \hat{f}_t = (\sum_{i=1}^N \hat{\Gamma}_i \Sigma_{ii}^{-1} \hat{\Gamma}_i')^{-1}(\sum_{i=1}^N \hat{\Gamma}_i \Sigma_{ii}^{-1} \hat{B}_{it}) \) to estimate \( f_t \), then the limiting distribution of \( \hat{f}_t - f_t \) is also the same as in pure factor models. The asymptotic representations on these estimators are implicitly contained in the appendix.

Now we present the most important result in this section. Throughout let \( \mathcal{M}(\mathcal{X}) \) denote the project matrix onto the space orthogonal to \( \mathcal{X} \), i.e. \( \mathcal{M}(\mathcal{X}) = I - \mathcal{X} \mathcal{X}' \mathcal{X}^{-1} \).

**Theorem 2.2 (Asymptotic representation).** Under the assumptions of Proposition 2.1, we have

\[
\hat{\beta} - \beta = \Omega^{-1} \frac{1}{NT} \sum_{t=1}^{N} \sum_{i=1}^{T} \Sigma_{ii}^{-1} e_{it} v_{itx} + O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1})
\]

where \( \Omega \) is a \( K \times K \) matrix, whose \((p, q)\) element \( \Omega_{pq} = \frac{1}{N} \sum_{i=1}^{N} \Sigma_{ii}^{-1} \Sigma_{ipq} \) with \( \Sigma_{ipq} \) being the \((p, q)\) element of matrix \( \Sigma_{ii} \).

**Remark 2.5.** In appendix A.3, we show that the asymptotic expression of \( \hat{\beta} - \beta \) can be alternatively expressed as

\[
(2.9) \quad \hat{\beta} - \beta = \left( \frac{\text{tr}[\hat{M} X_1 \mathcal{M}(\mathcal{F}) X_1'] \cdots \text{tr}[\hat{M} X_1 \mathcal{M}(\mathcal{F}) X_K']}{\text{tr}[\hat{M} X_K \mathcal{M}(\mathcal{F}) X_1'] \cdots \text{tr}[\hat{M} X_K \mathcal{M}(\mathcal{F}) X_K']} \right)^{-1} + O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1})
\]

where \( X_k = (x_{itk}) \) is \( N \times T \) (the data matrix for the \( k \)th regressor, \( k = 1, 2, \ldots, K \); \( e = (e_{it}) \) is \( N \times T \); \( \hat{M} = \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2}) \Sigma_{ee}^{-1/2} \) with \( \Sigma_{ee} = \text{diag}(\Sigma_{11e}, \Sigma_{22e}, \cdots, \Sigma_{NNe}) \) and \( \Lambda = (\lambda_1, \lambda_2, \ldots, \lambda_N) \); \( \mathcal{F} = (f_1, f_2, \cdots, f_T)' \); \( \mathcal{F} = (1_T, \mathcal{F}) \) where \( 1_T \) is a \( T \times 1 \) vector with all 1’s.
Remark 2.6. Theorem 2.2 shows that the asymptotic expression of $\hat{\beta} - \beta$ only involves variations in $e_{it}$ and $v_{itx}$. Intuitively, this is due to the fact that the error terms of the $y$ equation share the same factors with the explanatory variables. The variations from the common factor part of $x_{itk}$ (i.e., $\gamma_{ik}^{'}f_t$) do not provide information for $\beta$ since this part of information is offset by the common factor part of the error terms (i.e., $\lambda_{if}^{'}f_t$) in the $y$ equation.

Corollary 2.1 (Limiting distribution). Under the assumptions of Theorem 2.2, if $\sqrt{N/T} \to 0$, we have

$$\sqrt{N/T}(\hat{\beta} - \beta) \xrightarrow{d} N(0, \overline{\Omega}^{-1})$$

where $\overline{\Omega} = \lim_{N,T \to \infty} \Omega$, and $\overline{\Omega}$ is also the limit of

$$\overline{\Omega} = \lim_{N,T \to \infty} \frac{1}{NT} \begin{pmatrix}
\text{tr}[\hat{M}X_1M(\hat{F})X_1'] & \cdots & \text{tr}[\hat{M}X_1M(\hat{F})X_K'] \\
\vdots & \ddots & \vdots \\
\text{tr}[\hat{M}X_KM(\hat{F})X_1'] & \cdots & \text{tr}[\hat{M}X_KM(\hat{F})X_K']
\end{pmatrix}$$

Remark 2.7. The covariance matrix $\overline{\Omega}$ can be consistently estimated by

$$\frac{1}{NT} \begin{pmatrix}
\text{tr}[\hat{M}X_1M(\hat{F})X_1'] & \cdots & \text{tr}[\hat{M}X_1M(\hat{F})X_K'] \\
\vdots & \ddots & \vdots \\
\text{tr}[\hat{M}X_KM(\hat{F})X_1'] & \cdots & \text{tr}[\hat{M}X_KM(\hat{F})X_K']
\end{pmatrix},$$

where $X_k$ is the $N \times T$ data matrix for the $k$th regressor,

$$\hat{M} = \hat{\Sigma}_{ee}^{-1} - \hat{\Sigma}_{ee}^{-1}\hat{\Lambda}(\hat{\Lambda}'\hat{\Sigma}_{ee}^{-1}\hat{\Lambda})^{-1}\hat{\Lambda}'\hat{\Sigma}_{ee}^{-1};$$

$$\hat{F} = (1_T, \hat{F})$$ with $\hat{F} = (\hat{f}_1, \hat{f}_2, \ldots, \hat{f}_T)'$ and

$$\hat{f}_t = (N \sum_{i=1}^{N} \hat{\Gamma}_i\hat{\Sigma}_{ii}^{-1}\hat{\Gamma}_i')^{-1}(N \sum_{i=1}^{N} \hat{\Gamma}_i\hat{\Sigma}_{ii}^{-1}\hat{B}z_{it}).$$

Here $\hat{\Gamma}, \hat{\Lambda}, \hat{\Sigma}_{ii}, \hat{\Sigma}_{ee}$ and $\hat{B}$ are the maximum likelihood estimators.

Remark 2.8. We point out that the condition $\sqrt{N/T} \to 0$ is only needed for the limiting distribution to be of this simple form. The MLE for $\beta$ is still consistent under fixed $N$, but the limiting distribution will be different.
3. Common shock models with zero restrictions. The basic model in section 2 assumes that the explanatory variables \( x_{it} \) share the same factors with \( y_{it} \). This section relaxes this assumption. We assume that the regressors are impacted by additional factors that do not affect the \( y \) equation. An alternative view is that some factor loadings in the \( y \) equation are restricted to be zero. Consider the following model

\[
y_{it} = \alpha_i + x_{it1}\beta_1 + x_{it2}\beta_2 + \cdots + x_{itK}\beta_K + \psi'_it + e_{it} \\
x_{itk} = \mu_{ik} + \gamma'_{ik}g_t + \gamma'_{ik}h_t + \nu_{itk}
\]

for \( k = 1, 2, \cdots, K \), where \( g_t \) is an \( r_1 \times 1 \) vector representing the shocks affecting both \( y_{it} \) and \( x_{it} \), and \( h_t \) is an \( r_2 \times 1 \) vector representing the shocks affecting \( x_{it} \) only. Let \( \lambda_i = (\psi'_i, 0'_{r_2 \times 1})' \), \( \gamma_{ik} = (\gamma'_{ik}, \gamma'_{ik})' \) and \( f_t = (g'_t, h'_t)' \), the above model can be written as

\[
y_{it} = \alpha_i + x_{it1}\beta_1 + x_{it2}\beta_2 + \cdots + x_{itK}\beta_K + \lambda'_if_t + e_{it} \\
x_{itk} = \mu_{ik} + \gamma'_if_t + \nu_{itk}
\]

which is the same as model (2.1) except that \( \lambda_i \) now has \( r_1 \) free parameters and the remaining ones are restricted to be zeros. For further analysis, we introduce some notations. We define

\[
\Gamma^g_i = (\psi_i, \gamma'_{i1}, \cdots, \gamma'_{iK}), \quad \Gamma^h_i = (0_{r_2 \times 1}, \gamma'_{i1}, \cdots, \gamma'_{iK}), \\
\Gamma^g = (\Gamma^g_1, \Gamma^g_2, \cdots, \Gamma^g_N)' \quad \Gamma^h = (\Gamma^h_1, \Gamma^h_2, \cdots, \Gamma^h_N)'.
\]

We also define \( \mathbb{G} \) and \( \mathbb{H} \) similarly as \( \mathbb{F} \), i.e., \( \mathbb{G} = (g_1, g_2, \ldots, g_T)' \), \( \mathbb{H} = (h_1, h_2, \ldots, h_T)' \). This implies that \( \mathbb{F} = (\mathbb{G}, \mathbb{H}) \). The presence of zero restrictions in (3.1) requires different identification conditions from the previous model.

3.1. Identification conditions. Zero loading restrictions alleviate rotational indeterminacy. Instead of \( r^2 = (r_1 + r_2)^2 \) restrictions, we only need to impose \( r_1^2 + r_1r_2 + r_2^2 \) restrictions. These restrictions are referred to as IZ restrictions (Identification conditions with Zero restrictions). They are

IZ1 \( M_{ff} = I_r \)

IZ2 \( \frac{1}{N}\Gamma^g\Sigma^{-1}_{ee}\Gamma^g = D_1 \) and \( \frac{1}{N}\Gamma^h\Sigma^{-1}_{ee}\Gamma^h = D_2 \), where \( D_1 \) and \( D_2 \) are both diagonal matrices with distinct diagonal elements in descending order.

IZ3 \( 1'_f\mathbb{G} = 0 \) and \( 1'_f\mathbb{H} = 0 \).

In addition, we need an additional assumption for our analysis.

**Assumption E**: \( \Psi = (\psi'_1, \psi'_2, \ldots, \psi'_N)' \) is of full column rank.
Identification conditions IZ are less stringent than IB of the previous section. Assumption E says that the factors $g_t$ are pervasive for the $y$ equation. We next explain why $r_1^2 + r_1r_2 + r_2^2$ restrictions are sufficient. Let $R$ be an $r \times r$ invertible matrix, which we partition into

$$R = \begin{bmatrix} R_{11} & R_{12} \\ R_{21} & R_{22} \end{bmatrix}$$

where $R_{11}$ is $r_1 \times r_1$ and $R_{22}$ is $r_2 \times r_2$. The indeterminacy arises since equation (2.2) can be written as

$$(I_N \otimes B)z_t = \mu + \Gamma f_t + e_t = \mu + (\Gamma R)(R^{-1}f_t) + e_t$$

If we treat $\Gamma R$ as a new $\Gamma$ and $R^{-1}f_t$ as a new $f_t$, we have observationally equivalent models. However, in the present context there are many zero restrictions in $\Gamma$. If $\Gamma R$ is a qualified loading matrix, the same zero restrictions should be satisfied for $\Gamma R$. This leads to $\Psi R_{12} = 0$. If $\Psi$ is of full column rank, then left-multiplying $(\Psi \Psi)^{-1} \Psi'$ gives $R_{12} = 0$. This implies that we need $r_1^2 + r_1r_2 + r_2^2$ restrictions for full identification since $R_{11}$, $R_{21}$ and $R_{22}$ have $r_1^2 + r_1r_2 + r_2^2$ free parameters. As a comparison, if there are no restrictions in $\Gamma$, we need $r^2 = (r_1 + r_2)^2$ restrictions. Thus, zero loadings partially remove rotational indeterminacy. Notice IZ1 has $\frac{1}{2}r(r + 1)$ restrictions and IZ2 has $\frac{1}{2}r_1(r_1 - 1) + \frac{1}{2}r_2(r_2 - 1)$ restrictions. The total number of restrictions is thus $\frac{1}{2}r(r + 1) + \frac{1}{2}r_1(r_1 - 1) + \frac{1}{2}r_2(r_2 - 1) = r_1^2 + r_2^2 + r_1r_2$, the exact number we need.

3.2. Estimation. The likelihood function is now maximized under three sets of restrictions, i.e. $\frac{1}{N} \Gamma^y \Sigma_{ee}^{-1} \Gamma^y = D_1$, $\frac{1}{N} \Gamma^y \Sigma_{ee}^{-1} \Gamma^h = D_2$ and $\Phi = 0$ where $\Phi$ denotes the zero factor loading matrix in the $y$ equation. The likelihood function with the Lagrange multipliers is

$$\ln L = -\frac{1}{2N} \ln |\Sigma_{zz}| - \frac{1}{2N} \text{tr}
\left[ (I_N \otimes B)M_{zz}(I_N \otimes B')\Sigma_{zz}^{-1} \right]
+ \text{tr}
\left[ \Upsilon_1 \left( \frac{1}{N} \Gamma^y \Sigma_{ee}^{-1} \Gamma^y - D_1 \right) \right]
+ \text{tr}
\left[ \Upsilon_2 \left( \frac{1}{N} \Gamma^y \Sigma_{ee}^{-1} \Gamma^h - D_2 \right) \right] + \text{tr}(\Upsilon_3 \Phi),$$

where $\Sigma_{zz} = \Gamma' + \Sigma_{ee}$; $\Upsilon_1$ is $r_1 \times r_1$ and $\Upsilon_2$ is $r_2 \times r_2$, both are symmetric Lagrange multipliers matrices with zero diagonal elements; $\Upsilon_3$ is a Lagrange multiplier matrix of dimension $r_2 \times N$.

Let $U = \Sigma_{zz}^{-1}[(I_N \otimes B)M_{zz}(I_N \otimes B') - \hat{\Sigma}_{zz}]\hat{\Sigma}_{zz}^{-1}$. Notice $U$ is a symmetric matrix. The first order condition on $\Gamma^y$ gives

$$\frac{1}{N} \hat{\Gamma}^y U + \Upsilon_1 \frac{1}{N} \hat{\Gamma}^y \Sigma_{ee}^{-1} = 0.$$
Post-multiplying $\hat{\Gamma}^g$ yields

$$\frac{1}{N} \hat{\Gamma}^g U \hat{\Gamma}^g + \Upsilon_1 \frac{1}{N} \hat{\Gamma}^g \Sigma_x^{-1} \hat{\Gamma}^g = 0.$$  

Since $\frac{1}{N} \hat{\Gamma}^g U \hat{\Gamma}^g$ is a symmetric matrix, the above equation implies that $\Upsilon_1 \frac{1}{N} \hat{\Gamma}^g \Sigma_x^{-1} \hat{\Gamma}^g$ is also symmetric. But $\frac{1}{N} \hat{\Gamma}^g \Sigma_x^{-1} \hat{\Gamma}^g$ is a diagonal matrix. So the $(i, j)$th element of $\Upsilon_1 \frac{1}{N} \hat{\Gamma}^g \Sigma_x^{-1} \hat{\Gamma}^g$ is $\Upsilon_{1,ij}$, where $\Upsilon_{1,ij}$ is the $(i, j)$th element of $\Upsilon_1$ and $d_{ij}$ is the $j$th diagonal element of $D_1$. Given $\Upsilon_1 \frac{1}{N} \hat{\Gamma}^g \Sigma_x^{-1} \hat{\Gamma}^g$ is symmetric, we have $\Upsilon_{1,ij} d_{ij} = \Upsilon_{1,ji} d_{ij}$ for all $i \neq j$. However, $\Upsilon_1$ is also symmetric, so $\Upsilon_{1,ij} = \Upsilon_{1,ji}$. This gives $\Upsilon_{1,ij} (d_{ij} - d_{ii}) = 0$. Since $d_{ij} \neq d_{ii}$ by $IZ2$, we have $\Upsilon_{1,ij} = 0$ for all $i \neq j$. This implies $\Upsilon_1 = 0$ since the diagonal elements of $\Upsilon_1$ are all zeros.

Let $\Gamma^h_x = (\gamma^h_{1x}, \gamma^h_{2x}, \ldots, \gamma^h_{Nx})'$ with $\gamma^h_{1x} = (\gamma^h_{11}, \gamma^h_{12}, \ldots, \gamma^h_{1K})$, and $\Sigma_{xx} = \text{diag}(\Sigma_{11x}, \Sigma_{22x}, \ldots, \Sigma_{NNx})$, a block diagonal matrix of $NK \times NK$ dimension. We partition the matrix $U$ and define the matrix $\bar{U}$ as

$$U = \begin{pmatrix} U_{11} & U_{12} & \cdots & U_{1N} \\ U_{21} & U_{22} & \cdots & U_{2N} \\ \vdots & \vdots & \ddots & \vdots \\ U_{N1} & U_{N2} & \cdots & U_{NN} \end{pmatrix}, \quad \bar{U} = \begin{pmatrix} U_{11} & U_{12} & \cdots & U_{1N} \\ U_{21} & U_{22} & \cdots & U_{2N} \\ \vdots & \vdots & \ddots & \vdots \\ U_{N1} & U_{N2} & \cdots & U_{NN} \end{pmatrix}$$

where $U_{ij}$ is a $(K+1) \times (K+1)$ matrix and $U_{ij}$ is the lower-right $K \times K$ block of $U_{ij}$. Notice $\bar{U}$ is also a symmetric matrix. Then the first order condition on $\Gamma^h_x$ gives

$$\frac{1}{N} \hat{\Gamma}^h_x U + \Upsilon_2 \frac{1}{N} \hat{\Gamma}^h_x \Sigma_{xx}^{-1} \hat{\Gamma}^h_x = 0.$$  

Post-multiplying $\hat{\Gamma}^h_x$ yields

$$\frac{1}{N} \hat{\Gamma}^h_x U \hat{\Gamma}^h_x + \Upsilon_2 \frac{1}{N} \hat{\Gamma}^h_x \Sigma_{xx}^{-1} \hat{\Gamma}^h_x = 0.$$  

Notice $\frac{1}{N} \hat{\Gamma}^h_x U \hat{\Gamma}^h_x = \frac{1}{N} \hat{\Gamma}^h_x \Sigma_{xx}^{-1} \hat{\Gamma}^h_x = D_2$. By the similar arguments in deriving $\Upsilon_1 = 0$, we have $\Upsilon_2 = 0$. The interpretation for the zero Lagrange multipliers is that these constraints are non-binding for the likelihood. Whether or not these restrictions are imposed, the optimal value of the likelihood function is not affected, and neither is the efficiency of $\hat{\beta}$. In contrast, we cannot show $\Upsilon_3$ to be zero. Thus if $\Phi = 0$ is not imposed, the optimal value of the likelihood function and the efficiency of $\hat{\beta}$ will be affected. In Section 2, we did not use the Lagrange multiplier approach to impose the identification restrictions. Had it been used, we would have obtained zero valued
Lagrange multipliers. This is another view of why these restrictions do not affect the limiting distribution of $\hat{\beta}$. See Remark 2.4.

Now the likelihood function is simplified as

$$(3.2) \quad \ln L = -\frac{1}{2N} \ln |\Sigma_{zz}| - \frac{1}{2N} \text{tr}\left[(I_N \otimes B)M_{zz}(I_N \otimes B')\Sigma_{zz}^{-1}\right] + \text{tr} [\Upsilon_i' \Psi].$$

The first order condition on $\Gamma$ is

$$(3.3) \quad \hat{\Gamma}'\hat{\Sigma}_{zz}^{-1}[(I_N \otimes \hat{B})M_{zz}(I_N \otimes \hat{B}') - \hat{\Sigma}_{zz}]\hat{\Sigma}_{zz}^{-1} = W',$$

where $W'$ is a matrix having the same dimension as $\Gamma$, whose element is zero if the counterpart of $\Gamma$ is not specified to be zero, otherwise undetermined (containing the Lagrange multipliers). Post-multiplying $\hat{\Gamma}$ gives

$$\hat{\Gamma}'\hat{\Sigma}_{zz}^{-1}[(I_N \otimes \hat{B})M_{zz}(I_N \otimes \hat{B}') - \hat{\Sigma}_{zz}]\hat{\Sigma}_{zz}^{-1}\hat{\Gamma} = W'\hat{\Gamma}.$$

By the special structure of $W$ and $\hat{\Gamma}$, it is easy to verify that $W'\hat{\Gamma}$ has the form

$$\begin{bmatrix}
0_{r_1 \times r_1} & 0_{r_1 \times r_2} \\
0_{r_2 \times r_1} & 0_{r_2 \times r_2}
\end{bmatrix}.$$ 

However, the left hand side of the preceding equation is a symmetric matrix, so is the right side. It follows that the subblock "x" is zero, i.e. $W'\hat{\Gamma} = 0$. Thus, $\hat{\Gamma}'\hat{\Sigma}_{zz}^{-1}[(I_N \otimes \hat{B})M_{zz}(I_N \otimes \hat{B}') - \hat{\Sigma}_{zz}]\hat{\Sigma}_{zz}^{-1}\hat{\Gamma} = 0$. (This equation would be the first order condition for $M_{ff}$ if it were unknown.) This equality can be simplified as

$$(3.4) \quad \hat{\Gamma}'\hat{\Sigma}_{zz}^{-1}[(I_N \otimes \hat{B})M_{zz}(I_N \otimes \hat{B}') - \hat{\Sigma}_{zz}']\hat{\Sigma}_{zz}^{-1}\hat{\Gamma} = 0,$$

because $\hat{\Gamma}'\hat{\Sigma}_{zz}^{-1} = \hat{G}\hat{\Sigma}_{zz}^{-1}$ with $\hat{G} = (I + \hat{\Gamma}'\hat{\Sigma}_{zz}^{-1}\hat{\Gamma})^{-1}$. Next, we partition the matrix $\hat{G} = (I + \hat{\Gamma}'\hat{\Sigma}_{zz}^{-1}\hat{\Gamma})^{-1}$ and $\hat{H} = (\hat{\Gamma}'\hat{\Sigma}_{zz}^{-1}\hat{\Gamma})^{-1}$ as follows

$$\hat{G} = \begin{bmatrix} \hat{G}_{11} & \hat{G}_{12} \\ \hat{G}_{21} & \hat{G}_{22} \end{bmatrix}, \quad \hat{H} = \begin{bmatrix} \hat{H}_1 \\ \hat{H}_2 \end{bmatrix} = \begin{bmatrix} \hat{H}_{11} & \hat{H}_{12} \\ \hat{H}_{21} & \hat{H}_{22} \end{bmatrix},$$

where $\hat{G}_{11}, \hat{H}_{11}$ are $r_1 \times r_1$, while $\hat{G}_{22}, \hat{H}_{22}$ are $r_2 \times r_2$.

Notice $\hat{\Sigma}_{zz}^{-1} = \hat{\Sigma}_{zz}^{-1} - \hat{\Sigma}_{zz}^{-1}\hat{G}\hat{\Sigma}_{zz}^{-1}$ and $\hat{\Gamma}'\hat{\Sigma}_{zz}^{-1} = \hat{G}\hat{\Sigma}_{zz}^{-1}$. Substitute these results into (3.3) and use (3.4), the first order condition for $\psi_i$ can be simplified as

$$(3.5) \quad \hat{G}_{1i} \sum_{i=1}^{N} \hat{\Gamma}_i\hat{\Sigma}_{ii}^{-1}(\hat{B}M_{zz}^y\hat{B}' - \hat{\Sigma}_{ii}^y)\hat{\Sigma}_{jj}^{-1}I_{k+1} = 0,$$
where \( I_{K+1}^{i} \) is the first column of the identity matrix of dimension \( K + 1 \).

Similarly, the first order condition for \( \gamma_{jx} = (\gamma_{j1}, \gamma_{j2}, \ldots, \gamma_{jK}) \) is

\[
\sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{B}_M^i \hat{B}' - \hat{\Sigma}_{ii}^y) \hat{\Sigma}_{jj}^{-1} I_{K+1}^{i} = 0,
\]

where \( I_{K+1}^{i} \) is a \((K + 1) \times K\) matrix, obtained by deleting the first column of the identity matrix of dimension \( K + 1 \).

The first order condition for \( \Sigma_{jj} \) is

\[
\hat{B}_M^j \hat{B}' - \hat{\Sigma}_{jj}^y - \hat{\Gamma}_j \hat{G} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{B}_M^i \hat{B}' - \hat{\Sigma}_{ii}^y)
\]

\[
- \hat{\Sigma}_{jj}^y \hat{\Gamma}_j \hat{G} \hat{\Gamma}_j = W,
\]

where \( W \) is defined following (2.8).

The first order condition for \( \beta \) is

\[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{ii}^{-1} \left\{ (\hat{y}_{it} - \hat{x}_{it} \hat{\beta}) - \hat{\lambda}_i \hat{G} \sum_{j=1}^{N} \hat{\Gamma}_j \hat{\Sigma}_{jj}^{-1} \left[ \hat{y}_{jt} - \hat{x}_{jt} \hat{\beta} \right] \right\} \hat{x}_{it} = 0,
\]

which is the same as in Section 2.

We need an additional identity to study the properties of the MLE. Recall that, by the special structures of \( W \) and \( \hat{\Gamma} \), the three submatrices of \( W \hat{\Gamma} \) can be directly derived to be zeros. The remaining submatrix is also zero, as shown earlier. However, this submatrix being zero yields the following equation (the detailed derivation is delivered in Appendix B)

\[
\frac{1}{N} \hat{G} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{B}_M^i \hat{B}' - \hat{\Sigma}_{ii}^y) \hat{\Sigma}_{jj}^{-1} I_{K+1}^{j} \psi_{j}^{i} = 0.
\]

These identities for the MLE are used to derive the asymptotic representations.

3.3. Asymptotic properties of the MLE. The results on consistency and the rate of convergence are similar to those in the previous section, which are presented in Appendixes B.1 and B.2. For simplicity, we only state the asymptotic representation for the MLE here.
Under Assumptions A-E, we have

$$ P^0(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{itt}^{-1} e_{it} v_{itx} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{itt}^{-1} \gamma_{it} h_{it} e_{it} $$

$$ - \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{itt}^{-1} \psi_{i} \Pi_{\psi_{i}}^{-1} \left( \frac{1}{N} \sum_{j=1}^{N} \psi_{j} \Sigma_{j}^{-1} \gamma_{jt} h_{jt} \right) h_{it} e_{it} $$

$$ + O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) $$

where $P^0$ is a $K \times K$ symmetric matrix with its $(p,q)$ element equal to

$$ tr(\Gamma_p \hat{M} \Gamma_q^T) + \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i}^{-1} \Sigma_{itx}^{-1} \Sigma_{itt}^{-1} \gamma_{it} \gamma_{jt} \gamma_{kt} $$

$$ \Pi_{\psi_{i}} = \frac{1}{N} \sum_{i=1}^{N} \psi_{i} \Sigma_{itt}^{-1} \psi_{i} $$

and $\hat{M} = \Sigma_{e}^{-1/2} M(\Sigma_{e}^{-1/2} \Psi) \Sigma_{e}^{-1/2}$.  

Proposition 3.1 is derived under the identification conditions IZ. In Appendix B.3 of the supplement, we show that for any set of factors and factor loadings $(\psi_{i}, \gamma_{it}, g_{t}, h_{t})$, it can always be transformed into a new set $(\psi_{i}^{*}, \gamma_{it}^{*}, g_{t}^{*}, h_{t}^{*})$, which satisfies IZ, and at the same time, leaving $\Phi = 0$ intact. Given the asymptotic representation in Proposition 3.1, together with the relationship between the two sets, we have the following theorem, which doesn’t depend on IZ.

**Theorem 3.1.** Under Assumptions A-E, we have

$$ P(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{itt}^{-1} e_{it} v_{itx} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{itt}^{-1} \gamma_{it} h_{it}^{*} e_{it} $$

$$ - \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{itt}^{-1} \psi_{i} \Pi_{\psi_{i}}^{-1} \left( \frac{1}{N} \sum_{j=1}^{N} \psi_{j} \Sigma_{j}^{-1} \gamma_{jt} h_{jt}^{*} \right) h_{it}^{*} e_{it} $$

$$ + O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) $$

where

$$ h_{it}^{*} = h_{it} - \hat{G} \hat{G}' \hat{G} ' - \hat{g}_{it}^{*} $$

$P$ is a $K \times K$ symmetric matrix with its $(p,q)$ element equal to

$$ \frac{1}{NT} tr\left[ \hat{M} \hat{H}_{p} \hat{H}_{p}' \hat{M} \hat{H}_{p} \right] + \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i}^{-1} \Sigma_{itt}^{-1} \Sigma_{itx} $$

where $\overline{\Sigma} = (1, \Sigma); \Pi_{\psi_{i}} = \frac{1}{N} \sum_{i=1}^{N} \psi_{i} \Sigma_{itt}^{-1} \psi_{i}^{*}; \hat{M} = \Sigma_{e}^{-1/2} M(\Sigma_{e}^{-1/2} \Psi) \Sigma_{e}^{-1/2}, \hat{\Gamma}_{p} = \left( \gamma_{1p}^{h}, \gamma_{2p}^{h}, \ldots, \gamma_{Np}^{h} \right)$. 

**Proposition 3.1 (Asymptotic representation).**
Remark 3.1. In Appendix B.3, we also show that the asymptotic expression of $\hat{\beta} - \beta$ in Theorem 3.1 can be expressed alternatively as

$$\hat{\beta} - \beta = \left( \begin{array}{ccc} \text{tr}[\tilde{M}_1 M(\overline{G})_1 X_1'] & \cdots & \text{tr}[\tilde{M}_1 M(\overline{G})_K X_K'] \\ \vdots & \ddots & \vdots \\ \vdots & \ddots & \vdots \\ \text{tr}[\tilde{M}_K M(\overline{G})_1 X_1'] & \cdots & \text{tr}[\tilde{M}_K M(\overline{G})_K X_K'] \end{array} \right)^{-1}$$

$$\times \left( \begin{array}{c} \text{tr}[\tilde{M}_1 M(\overline{G}) e_1'] \\ \vdots \\ \text{tr}[\tilde{M}_K M(\overline{G}) e_K'] \end{array} \right) + O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}),$$

where $X_k$ and $e$ are defined below (2.9) and $\overline{G} = (1_T, G)$. Notice $\tilde{M}$ is defined as $\Sigma^{-1/2} e M(\Sigma^{-1/2} e \Psi) \Sigma^{-1/2} e$, which is equal to $\Sigma^{-1/2} M(\Sigma^{-1/2} e \Lambda) \Sigma^{-1/2} e$ since $\Lambda = (\Psi, 0_{N \times r_2})$ in the present context.

Remark 3.2. The alternative expression in the preceding remark can be explained in an intuitive way. Notice that the first equation of (3.1) can be written as

$$Y = X_1 \beta_1 + X_2 \beta_2 + \cdots + X_K \beta_K + \Psi G' + \alpha 1_T + e.$$

Post-multiplying $M(1_T)$, we have

$$Y' M(1_T) = X_1 M(1_T) \beta_1 + \cdots + X_K M(1_T) \beta_K + \Psi G' M(1_T) + e M(1_T).$$

Pre-multiplying $M(\Sigma^{-1/2} e) \Sigma^{-1/2} e$ and post-multiplying $M[M(1_T) G]$ and noticing $M(1_T) M[M(1_T) G] = M(\overline{G})$, we have

$$M(\Sigma^{-1/2} e) \Sigma^{-1/2} e Y' M(\overline{G}) = M(\Sigma^{-1/2} e) \Sigma^{-1/2} e M(\overline{G}) \beta_1 + \cdots + M(\Sigma^{-1/2} e) \Sigma^{-1/2} e X_K M(\overline{G}) \beta_K + M(\Sigma^{-1/2} e) \Sigma^{-1/2} e e M(\overline{G}).$$

The error term of the above equation, $M(\Sigma^{-1/2} e) \Sigma^{-1/2} e M(\overline{G})$, is asymptotically homoscedastic and uncorrelated over the cross section and over time. Applying the ordinary least squares (OLS) to the above equation, we will obtain the same limiting result. Of course, this method is infeasible because $\Lambda, \Sigma_{ee}$ and $G$ are unobservable. The maximum likelihood method amounts to making the unobservable factors and factor loadings observable.

Given Theorem 3.1 and Remark 3.1 we have the following corollary.
Corollary 3.1 (limiting distribution). Under Assumptions A-E, if $\sqrt{N/T} \to 0$, we have
$$\sqrt{NT}(\hat{\beta} - \beta) \overset{d}{\to} N(0, \mathcal{P}^{-1}),$$
where $\mathcal{P} = \lim_{N,T \to \infty} \mathcal{P}$, and $\mathcal{P}$ is also the probability limit of
$$\mathcal{P} = \text{plim}_{N,T \to \infty} \frac{1}{NT} \left( \begin{array}{ccc} \text{tr}[\hat{M}X_1\mathcal{M}(\mathcal{G})X_1'] & \cdots & \text{tr}[\hat{M}X_1\mathcal{M}(\mathcal{G})X_K'] \\ \vdots & & \vdots \\ \text{tr}[\hat{M}X_K\mathcal{M}(\mathcal{G})X_1'] & \cdots & \text{tr}[\hat{M}X_K\mathcal{M}(\mathcal{G})X_K'] \end{array} \right).$$

Remark 3.3. Compared with the model in Section 2, $\hat{\beta}$ is more efficient under the zero loading restrictions. The reason is intuitive. In the previous case, variations in $v_{itx}$ provide information for $\beta$. But in the present case, variations in $\gamma_{ik}h_t$ of $x_{it}$ also provide information for $\beta$. This can also be seen by comparing the limiting variances of Corollaries 2.1 and 3.1. Notice the projection matrix now only involves $\mathcal{G}$ instead of $\mathcal{F}$; and $\mathcal{G}$ is a submatrix of $\mathcal{F}$.

Remark 3.4. The covariance matrix $\mathcal{P}$ can be estimated by the same method as in estimating $\Omega$; see Remark 2.7.

Remark 3.5. Consider additional cases of zero restrictions. If $\gamma_{ik}^h = 0$ for all $i$, then model (3.1) reduces to the common shock model of Section 2 with $g_t$ as the common shocks only. If zero restrictions are imposed in the $x$ equation only instead of the $y$ equation, i.e.,
$$y_{it} = \alpha_i + x_{it1}\hat{\beta}_1 + x_{it2}\hat{\beta}_2 + \cdots + x_{itK}\hat{\beta}_K + \psi_i'g_t + \phi_i'h_t + e_{it},$$
$$x_{itk} = \mu_{ik} + \gamma_{ik}^hg_t + v_{itk}.$$ it can be shown that $\hat{\beta}$ obtained by imposing $\gamma_{ik}^h = 0$ has the same asymptotic representation as in Theorem 2.2; the zero restriction $\gamma_{ik}^h = 0$ does not improve the efficiency of $\hat{\beta}$. Finally, consider the following restricted model:
$$y_{it} = \alpha_i + x_{it1}\hat{\beta}_1 + x_{it2}\hat{\beta}_2 + \cdots + x_{itK}\hat{\beta}_K + \psi_i'g_t + e_{it},$$
$$x_{itk} = \mu_{ik} + \gamma_{ik}^{th}h_t + v_{itk}.$$ Here the $y$ and $x$ equations do not share the same factors, though $g_t$ and $h_t$ are correlated. In comparison with model (3.1), the above model imposes extra zero restrictions in the $x$ equation. Again, it can be shown that the zero restrictions in the $x$ equation do not improve the efficiency of $\hat{\beta}$. That is, the estimator with the additional restrictions has the same asymptotic representation as in Theorem 3.1.
4. Models with time-invariant regressors and common regressors. In this section, we extend the basic model in section 2 to include time-invariant regressors and common regressors. Examples of time-invariant regressors include gender, race and education; and examples for common regressors include price variables, unemployment rate, or macroeconomic policy variables. These types of regressors are important for empirical applications.

We first consider the model with only time-invariant regressors:

\[ y_{it} = \alpha_i + x_{it1}\beta_1 + x_{it2}\beta_2 + \cdots + x_{itK}\beta_K + \psi'_ig_t + \phi'_ih_t + \epsilon_{it} \]
\[ x_{itk} = \mu_{ik} + \gamma'_{itk}g_t + \gamma'_{itk}h_t + v_{itk} \]

for \( k = 1, 2, \cdots, K \), where \( g_t \) is an \( r_1 \)-dimensional vector and \( h_t \) is an \( r_2 \)-dimensional vector. Let \( f_t = (g'_t, h'_t)' \), an \( r \)-dimensional vector. The key point of model (4.1) is that the \( \phi_i \)'s are known (but not zeros). We treat \( \phi_i \) as new added time-invariant regressors, whose coefficient \( h_t \) is allowed to be time-varying. The parameters of interest, besides \( \beta \), include \( h_t \) for all \( t \). The model in the previous section can be viewed as \( \Phi = 0 \), where \( \Phi = (\phi_1, \phi_2, \cdots, \phi_N)' \). However, the earlier derivation is not applicable here because now \( \Phi \) is a general matrix with full column rank, which provides more information (restrictions) on the rotation matrix. Thus the number of restrictions required to eliminate rotational indeterminacy is even fewer than in Section 3. This point can be seen in the next subsection.

We define the following notation for further analysis:

\[ \Gamma^g_i = (\psi_i, \gamma'_{i1}, \cdots, \gamma'_{iK}), \quad \Gamma^h_i = (\phi_i, \gamma'_{i1}, \cdots, \gamma'_{iK}), \quad \Gamma_i = (\Gamma^g_i, \Gamma^h_i)', \]
\[ \Lambda = (\lambda_1, \lambda_2, \cdots, \lambda_N)', \quad \Psi = (\psi_1, \psi_2, \cdots, \psi_N)', \quad \lambda_i = (\psi_i', \phi_i)', \]
\[ \Phi = (\phi_1, \phi_2, \cdots, \phi_N)' \].

Then equation (4.1) has the same matrix expression as (2.2). Note that \( \Lambda = [\Psi, \Phi] \) is the factor loading matrix for the \( N \times 1 \) vector \((y_{1t}, y_{2t}, \cdots, y_{NT})'\).

4.1. Identification Conditions. We make the following identification conditions, which we refer to as IO (Identification conditions with partial Observable fixed effects), to emphasize the observed fixed effects.

IO1. We partition the matrix \( M_{ff} \) as

\[ M_{ff} = \begin{bmatrix} M_{gg} & M_{gh} \\ M_{hg} & M_{hh} \end{bmatrix} \]

and impose \( M_{gh} = 0 \) and \( M_{gg} = I_{r_1} \).
IO2. \( \frac{1}{N} \Gamma^g \Sigma_{ee}^{-1} \Gamma^g = D \), where \( D \) is a diagonal matrix with its diagonal elements distinct and arranged in descending order.

IO3. \( 1_T^r \mathcal{G} = 0 \) and \( 1_T^r \mathcal{H} = 0 \).

These restrictions can be imposed without loss of generality, as argued later and more formally in Appendix C.3. In addition, we make the following assumption.

**Assumption F:** The loading matrix \( \Lambda = [\Psi, \Phi] \) is of full column rank.

Now we use the method in Section 3.1 to show that IO is enough to achieve full identification. Let \( R \) be the rotation matrix partitioned in the same way as in Section 3.1. If there exists some matrix \( \Gamma^\dagger \), which shares the same structure with \( \Gamma \), satisfying \( \Gamma^\dagger = \Gamma R \), then we must have \( \Psi R_{12} + \Phi R_{22} = \Phi \) because \( \Phi \) is given. This is equivalent to

\[
\begin{bmatrix}
\Psi & \Phi
\end{bmatrix}
\begin{bmatrix}
R_{12} \\
I - R_{22}
\end{bmatrix} = \Lambda \begin{bmatrix}
R_{12} \\
I - R_{22}
\end{bmatrix} = 0.
\]

So if matrix \( \Lambda \) is of full column rank, then pre-multiplying \( (\Lambda')^{-1} \Lambda' \) gives \( R_{12} = 0 \) and \( R_{22} = I \). Thus only \( R_{11} \) and \( R_{12} \) are undetermined. This implies that we only need \( r_1^2 + r_1 r_2 \) restrictions. The number of restrictions implied by IO1 and IO2 is exactly \( r_2^2 + r_1 r_2 \). So there is further reduction in the number of restrictions to eliminate rotational indeterminacy.

### 4.2. Estimation.

For clarity, in this subsection, we use \( \Phi^* \) to denote the observed value for \( \Phi \). Recall that \( \Sigma_{zz} = \Gamma M_{ff} \Gamma' + \Sigma_{ee} \), where \( \Gamma \) contains the factor loading coefficients (including \( \Phi \)); \( M_{ff} \) contains the sub-blocks \( M_{gg}, M_{gh}, \) and \( M_{hh} \); \( \Sigma_{ee} \) contains the heteroskedasticity coefficients. The regression coefficient \( \beta \) is contained in matrix \( B \). The maximization of the likelihood function is now subject to four sets of restrictions, \( M_{gh} = 0, M_{gg} = I_{r_1}, \Phi = \Phi^*, \) and \( \frac{1}{N} \Gamma^g \Sigma_{ee}^{-1} \Gamma^g = D \). The likelihood function augmented with the Lagrange multipliers is

\[
\ln L = -\frac{1}{2N} \ln |\Sigma_{zz}| - \frac{1}{2N} \text{tr} \left[ (I_N \otimes B)(I_N \otimes B') \Sigma_{zz}^{-1} \right] + \text{tr} \left[ \Upsilon_1 M_{gh} \right] + \text{tr} \left[ \Upsilon_2 (M_{gg} - I_{r_1}) \right] + \text{tr} \left[ \Upsilon_3 \left( \frac{1}{N} \Gamma^g \Sigma_{ee}^{-1} \Gamma^g - D \right) \right] + \text{tr} \left[ \Upsilon_4 (\Phi - \Phi^*) \right],
\]

where \( \Upsilon_1, \Upsilon_2, \Upsilon_3, \) and \( \Upsilon_4 \) are all Lagrange multipliers matrices; \( \Upsilon_1 \) is an \( r_2 \times r_1 \) matrix; \( \Upsilon_2 \) is an \( r_1 \times r_1 \) symmetric matrix; \( \Upsilon_3 \) is an \( r_1 \times r_1 \) symmetric matrix with all diagonal elements zeros; \( \Upsilon_4 \) is an \( r_2 \times N \) matrix; and \( \Sigma_{zz} = \Gamma M_{ff} \Gamma' + \Sigma_{ee} \). Using the same arguments in deriving \( \Upsilon_1 = 0 \) in Section 3,
we have \( \Upsilon_3 = 0 \). Then the likelihood function is simplified as

\[
\ln L = -\frac{1}{2N} \ln |\Sigma_z| - \frac{1}{2N} \text{tr}\left[ (I_N \otimes B)M_z(I_N \otimes B')\Sigma_z^{-1} \right]
+ \text{tr}[\Upsilon_1 M_g] + \text{tr}[\Upsilon_2(M_{yy} - I_{r_1})] + \text{tr}[\Upsilon_4(\Phi - \Phi^*)].
\]  

(4.2)

The first order condition for \( \Gamma \) gives

\[
\hat{M}_{ff} \hat{\Gamma}' \hat{\Sigma}_z^{-1}[(I_N \otimes \hat{B})M_z(I_N \otimes \hat{B}') - \hat{\Sigma}_z] \hat{\Sigma}_z^{-1} = W',
\]

where \( W \) is defined in (3.3). Pre-multiplying \( \hat{M}_{ff}^{-1} \) and post-multiplying \( \hat{\Gamma} \), and by the special structures of \( W \) and \( \hat{\Gamma} \), we have

\[
\frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_z^{-1}[(I_N \otimes \hat{B})M_z(I_N \otimes \hat{B}') - \hat{\Sigma}_z] \hat{\Sigma}_z^{-1} \hat{\Gamma} = \begin{bmatrix} 0_{r_1 \times r_1} & 0_{r_1 \times r_2} \\ \frac{1}{N} M_{hh}^{-1} \Upsilon_1' \hat{\Psi} & \frac{1}{N} M_{hh}^{-1} \Upsilon_4' \Phi \end{bmatrix}.
\]

But the first order condition for \( \Sigma_{ii} \) is the same as (3.8), i.e.

\[
\hat{\Gamma}' \hat{\Sigma}_z^{-1}[(I_N \otimes \hat{B})M_z(I_N \otimes \hat{B}') - \hat{\Sigma}_z] \hat{\Sigma}_z^{-1} \hat{\Gamma} = 0.
\]

Comparing the proceeding two results and noting that the left hand side is a symmetric matrix, we have \( \hat{\Gamma}' \hat{\Sigma}_z^{-1}[(I_N \otimes \hat{B})M_z(I_N \otimes \hat{B}') - \hat{\Sigma}_z] \hat{\Sigma}_z^{-1} \hat{\Gamma} = 0 \).

But \( \hat{\Gamma}' \hat{\Sigma}_z^{-1} \) can be replaced by \( \hat{\Gamma}' \hat{\Sigma}_z^{-1} \) (see (S.2) in Appendix). Thus

\[
(4.3) \quad \hat{\Gamma}' \hat{\Sigma}_z^{-1}[(I_N \otimes \hat{B})M_z(I_N \otimes \hat{B}') - \hat{\Sigma}_z] \hat{\Sigma}_z^{-1} \hat{\Gamma} = 0.
\]

The above result implies that \( \Upsilon_1 = 0 \), \( \Upsilon_2 = 0 \), \( \Upsilon_4' \hat{\Psi} = 0 \) and \( \Upsilon_4' \Phi = 0 \).

The first order condition for \( \Sigma_{ii} \) is the same as (3.7), i.e.

\[
\hat{B}M_{zz} \hat{B}' - \hat{\Sigma}_{zz} - \hat{\Gamma}_j' \hat{G} \sum_{i=1}^N \hat{\Gamma}_{ii}^{-1} \left( \hat{B}M_{zz} \hat{B}' - \hat{\Sigma}_{zz} \right) - \sum_{i=1}^N \left( \hat{B}M_{zz} \hat{B}' - \hat{\Sigma}_{zz} \right) \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_j' \hat{G} \hat{\Gamma}_j = \mathcal{W},
\]

(4.4)

where \( \mathcal{W} \) is defined following (2.8).

The first order condition on \( \beta \) is the same as (3.8), i.e.

\[
(4.5) \quad \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \hat{\Sigma}_{ite}^{-1} \left\{ (\hat{y}_{it} - \hat{x}_{it} \hat{\beta}) - \hat{\lambda}_i \hat{G} \sum_{j=1}^N \hat{\Gamma}_{jj} \hat{\Sigma}_{jj}^{-1} \left[ \hat{y}_{jt} - \hat{x}_{jt} \hat{\beta} \right] \right\} \hat{x}_{it} = 0.
\]

We need an additional identify for the theoretical analysis in the appendix. The preceding analysis shows that \( \frac{1}{N} \Upsilon_4' \hat{\Psi} = 0 \) and \( \frac{1}{N} \Upsilon_4' \Phi = 0 \). They imply

\[
(4.6) \quad \frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N \hat{G}_j \hat{\Gamma}_{ii} \hat{\Sigma}_{ii}^{-1} \left( \hat{B}M_{zz} \hat{B}' - \hat{\Sigma}_{zz} \right) \hat{\Sigma}_{jj}^{-1} I_{K+i} \hat{\lambda}_j = 0,
\]

where \( \hat{\lambda}_j = (\hat{\psi}_j', \phi_j')' \).
4.3. Asymptotic properties. The following theorem states the asymptotic representation for $\hat{\beta} - \beta$.

**Proposition 4.1.** Under Assumptions A-D and F, and under the identification condition IO, we have

\[
Q_0(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iix} \varepsilon_{it} \varepsilon_{it} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \gamma_t \hat{\epsilon}_{it} \epsilon_{it} \\
- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iix} \lambda_i \Pi^{-1}_{\lambda \lambda} \left( \frac{1}{N} \sum_{j=1}^{N} \lambda_j \Sigma^{-1}_{ijx} \gamma_{jx} \right) \hat{\epsilon}_{it} \epsilon_{it} \\
+ O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}),
\]

where $Q_0$ is a $K \times K$ symmetric matrix with its $(p, q)$ element equal to $\text{tr}[M_p \Gamma_p \hat{\Gamma}_q] + \frac{1}{N} \sum_{i=1}^{N} \Sigma^{-1}_{iix} \gamma_t \hat{\epsilon}_{it} \epsilon_{it}$; $\hat{\Gamma}_q = \sum_{i=1}^{N} \lambda_i \Sigma^{-1}_{ijx} \gamma_{jx}$; and $\gamma^h_{iij}$, $\gamma^h_{ii}$, $\gamma^h_{i}$.

Proposition 4.1 is derived under the identification conditions IO. In Appendix C.3, we show that for any set of factors and factor loadings $(\psi^*, \gamma^*, g^*, h^*)$, we can always transform it to another set $(\psi^*_1, \gamma^*_1, g^*_1, h^*_1)$ which satisfies IO, and at the same time, still maintains the observability of $\Phi$ (i.e., $\Phi$ is untransformed). This is in agreement with the Lagrange multiplier analysis, in which $\Upsilon_j = 0$ ($j = 1, 2, 3$); the only binding restriction is $\Phi = \Phi^*$. Using the relationship between the two sets, we can generalize Proposition 4.1 into the following theorem, which does not depend on IO.

**Theorem 4.1.** Under Assumptions A-D and F, we have

\[
Q(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iix} \varepsilon_{it} \varepsilon_{it} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \gamma_t \hat{\epsilon}_{it} \epsilon_{it} \\
- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iix} \lambda_i \Pi^{-1}_{\lambda \lambda} \left( \frac{1}{N} \sum_{j=1}^{N} \lambda_j \Sigma^{-1}_{ijx} \gamma_{jx} \right) \hat{\epsilon}_{it} \epsilon_{it} \\
+ O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1})
\]

where

\[
h^*_t = h_t - \hat{\Gamma}_q \hat{G}_t \hat{G}_t \hat{\epsilon}_t,
\]

$Q$ is a $K \times K$ symmetric matrix with its $(p, q)$ element equal to

\[
\text{tr}[\hat{\Gamma}_p \hat{G}_t \hat{G}_t \hat{\epsilon}_t] + \frac{1}{N} \sum_{i=1}^{N} \Sigma^{-1}_{iix} \gamma_t \hat{\epsilon}_{it} \epsilon_{it}.
\]
and $\hat{M}_t$, $\Gamma_p$ and $\Pi_{\lambda\lambda}$ are defined in Proposition 4.1.

**Remark 4.1.** In Appendix C.3 we show that the asymptotic expression of $\hat{\beta} - \beta$ in Theorem 4.1 can be expressed alternatively as

$$
\hat{\beta} - \beta = \left( \begin{array}{ccc}
\text{tr}[\hat{M}X_1\mathcal{M}(\overline{e})X_1'] & \cdots & \text{tr}[\hat{M}X_1\mathcal{M}(\overline{e})X_k'] \\
\vdots & \ddots & \vdots \\
\text{tr}[\hat{M}X_K\mathcal{M}(\overline{e})X_1'] & \cdots & \text{tr}[\hat{M}X_K\mathcal{M}(\overline{e})X_K']
\end{array} \right)^{-1}
\times \left( \begin{array}{c}
\text{tr}[\hat{M}X_1\mathcal{M}(\overline{e})e'] \\
\vdots \\
\text{tr}[\hat{M}X_K\mathcal{M}(\overline{e})e']
\end{array} \right) + O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}),
$$

where $X_k$ and $e$ are defined below (2.9) and $\overline{e} = (1_T, G)$.

**Remark 4.2.** The expression in Remark 4.1 has a similar intuitive explanation as in Section 3. The first equation of (4.1) can be written as

$$
Y = X_1\beta_1 + X_2\beta_2 + \cdots + X_k\beta_K + \Psi G' + \Phi \overline{e}' + \alpha 1_T + e.
$$

Pre-multiply $\Sigma_{ee}^{-1/2}$ to eliminate the heteroscedasticity

$$
\Sigma_{ee}^{-1/2}Y = \Sigma_{ee}^{-1/2}X_1\beta_1 + \cdots + \Sigma_{ee}^{-1/2}X_k\beta_K + \Sigma_{ee}^{-1/2}\Psi G' + \Sigma_{ee}^{-1/2}\Phi \overline{e}' + \Sigma_{ee}^{-1/2}\alpha 1_T + \Sigma_{ee}^{-1/2}e.
$$

Both $\Phi$ and $1_T$ are observable. Pre-multiplying $\mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)$ and post-multiplying $\mathcal{M}(1_T)$, we eliminate $\Phi \overline{e}'$ and $\alpha 1_T$. Thus

$$
\mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}Y \mathcal{M}(1_T) = \mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}X_1\mathcal{M}(1_T)\beta_1 + \cdots + \mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}X_k\mathcal{M}(1_T)\beta_K + \mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}\Psi \mathcal{M}(1_T) + \mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}\alpha 1_T.
$$

Notice that both $\mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}\Psi$ and $G' \mathcal{M}(1_T)$ are unobservable. Pre-multiply $\mathcal{M}[\mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}\Psi]$ and $\mathcal{M}[\mathcal{M}(1_T)\mathcal{G}]$ to eliminate the unobservable common factors. Using the result that $\mathcal{M}(1_T)\mathcal{M}[\mathcal{M}(1_T)\mathcal{G}] = \mathcal{M}(\overline{e})$ and $\mathcal{M}[\mathcal{M}(\Sigma_{ee}^{-1/2}\Phi)\Sigma_{ee}^{-1/2}\Psi]\mathcal{M}(\Sigma_{ee}^{-1/2}\Phi) = \mathcal{M}(\Sigma_{ee}^{-1/2}\Lambda)$, we have

$$
\mathcal{M}(\Sigma_{ee}^{-1/2}\Lambda)\Sigma_{ee}^{-1/2}Y \mathcal{M}(\overline{e}) = \mathcal{M}(\Sigma_{ee}^{-1/2}\Lambda)\Sigma_{ee}^{-1/2}X_1\mathcal{M}(\overline{e})\beta_1 + \cdots + \mathcal{M}(\Sigma_{ee}^{-1/2}\Lambda)\Sigma_{ee}^{-1/2}X_K\mathcal{M}(\overline{e})\beta_K + \mathcal{M}(\Sigma_{ee}^{-1/2}\Lambda)\Sigma_{ee}^{-1/2}e \mathcal{M}(\overline{e}).
$$

The error term of the above equation, $\mathcal{M}(\Sigma_{ee}^{-1/2}\Lambda)\Sigma_{ee}^{-1/2}e \mathcal{M}(\overline{e})$, is asymptotically homoscedastic and uncorrelated over the cross section and over time. Applying OLS to the above equation, we have the same asymptotic expression. Again, this operation is infeasible in practice, but the MLE makes it asymptotically feasible.
From Theorem 4.1, we obtain the following corollary.

**Corollary 4.1.** Under the conditions of Theorem 4.1, if \( \sqrt{N/T} \to 0 \), we have
\[
\sqrt{NT}(\hat{\beta} - \beta) \overset{d}{\to} N(0, \mathbb{Q}^{-1}),
\]
where \( \mathbb{Q} = \lim_{N,T \to \infty} Q \), which has an alternative expression
\[
\mathbb{Q} = \text{plim}_{N,T \to \infty} \frac{1}{NT} \left( \begin{array}{c} \text{tr}[\tilde{M}X_1 \mathcal{M}(\bar{G})X'_1] \\ \vdots \\ \text{tr}[\tilde{M}X_K \mathcal{M}(\bar{G})X'_K] \end{array} \right).
\]

**Remark 4.3.** Compared with the model in Section 2, \( \hat{\beta} \) is more efficient with observable fixed effects (time-invariant regressors). The reason is provided in Remark 3.3.

We next consider estimating \( h_t \). It is worth emphasizing that unlike the asymptotic theory for \( \hat{\beta} \), where the identification conditions in IO are inessential but facilitate the theoretical analysis, the asymptotic theory of \( \hat{h}_t \) depends on the identification conditions, without which \( h_t \) is not identifiable. In what follows, we assume that the underlying parameters satisfy IO.

We estimate \( h_t \) by the following formula:
\[
\hat{h}_t = \left[ \Phi' \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi \right]^{-1} \left[ \Phi' \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi \right] \left[ \hat{Y}_t - \hat{X}_t \beta \right],
\]
where \( \hat{Y}_t = (y_{1t}, y_{2t}, \ldots, y_{Nt})' \) and \( \hat{Y}_t = \bar{Y}_t - T^{-1} \sum_{t=1}^{T} \bar{Y}_t \); \( \hat{X}_t \) is an \( N \times K \) matrix with its \( (i, k) \) element \( x_{itk} \) and \( \hat{X}_t = \bar{X}_t - T^{-1} \sum_{t=1}^{T} \bar{X}_t \). Now we state the limiting result for \( \hat{h}_t \).

**Theorem 4.2.** Under Assumptions A-D, F, and the identification conditions IO, if \( \sqrt{N/T} \to 0 \), then
\[
\sqrt{N}(\hat{h}_t - h_t) \overset{d}{\to} N(0, \left[ \lim_{N \to \infty} \frac{1}{N} \Phi' \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi \right]^{-1}).
\]

Theorem 4.2 also has an intuitive explanation. Consider the first equation of (4.1), which can be written as
\[
\hat{Y}_t = \hat{X}_t \beta + \Psi g_t + \Phi h_t + \alpha + e_t.
\]

First remove \( \alpha \) from the above equation, this gives (Note 1'G = 0, 1'X = 0)
\[
\hat{Y}_t = \hat{X}_t \beta + \Psi g_t + \Phi h_t + \hat{e}_t.
\]
Then pre-multiplying $\mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2}$, we have

$$\mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2}(\dot{\beta}_t - \bar{X}_t \beta) = \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi h_t + \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \dot{e}_t.$$  

The error term, $\mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \dot{e}_t$, is asymptotically homoscedastic and uncorrelated. Applying OLS to the above equation, we obtain the same limiting result as stated in Theorem 4.2.

Theorem 4.1 shows that $\hat{\beta} - \beta$ is $O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-3/2})$. But the convergence rate of $\hat{h}_t - h_t$ is only $\sqrt{N}$. So under the conditions $\sqrt{N}/T \to 0$, we can treat $\beta$ as known. The limiting distribution of $\hat{\beta}_t - h_t$ is the same as the case of a pure multiple battery factor model (without regressors) with known loading $\Phi$. The asymptotic representations of $\dot{\psi}_t - \psi_t$ and $\dot{g}_t - g_t$ can also be derived by the same arguments for $\hat{h}_t - h_t$. These results share the common feature that the limiting representations for one set of parameters are identical to the situation in which the remaining parameters can be treated as observable.

4.4. Models with time-invariant regressors and common regressors. In this subsection, we consider the joint presence of time-invariant regressors and common regressors. Consider the following model

$$(4.7) \quad y_{it} = x_{i1} \beta_1 + x_{i2} \beta_2 + \cdots + x_{iK} \beta_K + \psi'_i g_t + \phi'_i h_t + \kappa'_i d_t + e_{it},$$

$$x_{itk} = \gamma_{ik}' g_t + \gamma_{ik} h_t + \gamma_{ik}' d_t + v_{itk}$$

for $k = 1, 2, \ldots, K$, where $g_t$, $h_t$ and $d_t$ are $r_1 \times 1$, $r_2 \times 1$ and $r_3 \times 1$ vectors, respectively. A key feature of model (4.7) is that $d_t$ and $\phi_i$ are observable for all $i$ and $t$. We call $\phi_i$ the time-invariant regressors because they are invariant over time and $d_t$ the common regressors because they are the same for all the cross-sectional units. In this model, the time-invariant regressors have time-varying coefficients, and the common regressors have heterogeneous (individual-dependent) coefficients. The parameters of interest now, besides the coefficient $\beta$, include $\kappa_i$ and $h_t$. If $d_t \equiv 1$, $\kappa_i$ plays the role of $\alpha_i$ in (4.1). So the model here is more general.

Similarly as the previous subsection, we make the following assumption:

**Assumption G:** The matrices $(\Psi, \Phi, K)$ and $(G, H, D)$ are both of full column rank, where $K = (\kappa_1, \kappa_2, \ldots, \kappa_N)'$ and $D = (d_1, d_2, \ldots, d_T)'$.

Let $\lambda_i = (\psi_i', \phi_i')'$, $\gamma_{ik} = (\gamma_{ik}', \gamma_{ik} h_t)'$ and $\delta_i = (\kappa_i, \gamma_{ik}')$. The model can be written as

$$\begin{bmatrix} 1 & -\beta' \\ 0 & I_K \end{bmatrix} z_{it} = \Gamma'_i f_t + \delta'_i d_t + \varepsilon_{it},$$
where $z_{it}, \Gamma_i, \varepsilon_{it}$ are defined in Section 2; Let $\Delta = (\delta_1, \delta_2, \ldots, \delta_N)'$, then we have

\begin{equation}
(I_N \otimes B) z_{it} - \Delta d_t = \Gamma f_t + \varepsilon_t,
\end{equation}

where the symbols $\Gamma, z_{it}, B, \varepsilon_t$ are defined in Section 2.

The likelihood function can be written as

$$
\ln L = -\frac{1}{2N} \ln |\Sigma_{zz}| - \frac{1}{2NT} \sum_{t=1}^{T} [(I_N \otimes B) z_{it} - \Delta d_t]' \Sigma_{zz}^{-1} [(I_N \otimes B) z_{it} - \Delta d_t].
$$

Take $\Sigma_{zz}$ and $\beta$ as given, $\Delta$ maximizes the above function at

$$
\hat{\Delta} = (I_N \otimes B) (\sum_{s=1}^{T} z_s d_s') (\sum_{s=1}^{T} d_s d_s')^{-1}.
$$

Substituting $\hat{\Delta}$ into the above likelihood function, we obtain the concentrated likelihood function

$$
\ln L = -\frac{1}{2N} \ln |\Sigma_{zz}| - \frac{1}{2NT} \text{tr} [(I_N \otimes B) Z M(\mathcal{D}) Z'(I_N \otimes B') \Sigma_{zz}^{-1}],
$$

where $Z = (z_1, z_2, \ldots, z_T), \mathcal{D} = (d_1, d_2, \ldots, d_T)'$, and $M(\mathcal{D}) = I_T - \mathcal{D}' \mathcal{D}'^{-1} \mathcal{D}'$, a projection matrix. Consider (4.8), which is equivalent to

\begin{equation}
(I_N \otimes B) Z = \Gamma F' + \Delta D' + \varepsilon,
\end{equation}

where $\varepsilon = (\varepsilon_1, \varepsilon_2, \ldots, \varepsilon_T)$. Post-multiplying $M(\mathcal{D})$ on both sides, we have

\begin{equation}
(I_N \otimes B) Z M(\mathcal{D}) = \Gamma F' M(\mathcal{D}) + \varepsilon M(\mathcal{D}).
\end{equation}

If we treat $Z M(\mathcal{D})$ as the new observable data, $F' M(\mathcal{D})$ as the new unobservable factors, the preceding equation can be viewed as a special case of (4.1). Invoking Theorem 4.1, which does not need IO (the factors $F' M(\mathcal{D})$ may not satisfy IO), we have the following theorem:

**Theorem 4.3.** *Under Assumptions A-D and G, the asymptotic representation of $\hat{\beta}$ in the presence of time invariant and common regressors is*

\begin{equation}
\mathcal{R}(\hat{\beta} - \beta) = 
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ii}^{-1} e_{it} v_{itx} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ii}^{-1} \gamma_{ix} h_{it} e_{it}
\end{equation}

\begin{equation}
- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ii}^{-1} \chi_i \Pi_A^{-1} \frac{1}{N} \sum_{j=1}^{N} \chi_j \Sigma_{jj}^{-1} h_{jt} e_{it}
\end{equation}

\begin{equation}
+ O_p(T^{-3/2}) + O_p(N^{-1} T^{-1/2}) + O_p(N^{-1/2} T^{-1}),
\end{equation}
where

\[ h_t^* = h_t - \mathbb{H}' \mathbb{D} (\mathbb{D}' \mathbb{D})^{-1} d_t - \mathbb{H}' \mathcal{M}(\mathbb{D}) \mathbb{G} [G' \mathcal{M}(\mathbb{D}) \mathbb{G}]^{-1} (g_t - G' \mathbb{D} (\mathbb{D}' \mathbb{D})^{-1} d_t); \]

\( \mathcal{R} \) is a \( K \times K \) symmetric matrix with its \((p, q)\) element equal to

\[ \text{tr}[\mathcal{M}^h q \mathbb{H}' \mathcal{M} (\mathbb{B}) \mathbb{H} \Gamma^h_p] + \frac{1}{N} \sum_{i=1}^{N} \Sigma_{ii}^{-1} \lambda_i \sigma_{ii}; \]

where \( b_t = (g'_t, d'_t)' \) and \( \mathbb{B} = (b_1, b_2, \ldots, b_T)' = (\mathbb{G}, \mathbb{D}) \), a matrix of \( T \times (r_1 + r_3) \) dimension; \( \mathcal{M} = \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Lambda) \Sigma_{ee}^{-1/2}; \Gamma^h_p = (\gamma^h_1, \gamma^h_2, \ldots, \gamma^h_N)' \);

\[ \Pi_{\lambda \lambda} = \frac{1}{N} \sum_{i=1}^{N} \lambda_i \Sigma_{ii}^{-1} \lambda_i. \]

**Remark 4.4.** The asymptotic expression of \( \hat{\beta} - \beta \) can be alternatively expressed as

\[
\hat{\beta} - \beta = \left( \begin{array}{ccc}
\text{tr}[\mathcal{M} X_1 \mathcal{M}(\mathbb{B}) X_1] & \cdots & \text{tr}[\mathcal{M} X_K \mathcal{M}(\mathbb{B}) X_K]
\end{array} \right)^{-1}
\left( \begin{array}{c}
\text{tr}[\mathcal{M} X_1 \mathcal{M}(\mathbb{B}) e'] \\
\vdots \\
\text{tr}[\mathcal{M} X_K \mathcal{M}(\mathbb{B}) e']
\end{array} \right) + O_p(T^{-3/2}) + O_p(N^{-1} T^{-1/2}) + O_p(N^{-1/2} T^{-1}).
\]

If \( \mathbb{D} = 1_T \), the above asymptotic result reduces to the one in Theorem 4.1 since \( \mathbb{B} = (1_T, \mathbb{G}) = \mathbb{G} \).

Given Theorem 4.3 and Remark 4.4, we have the following corollary.

**Corollary 4.2.** Under Assumptions A-D and G, if \( \sqrt{N}/T \to 0 \), then

\[ \sqrt{NT} (\hat{\beta} - \beta) \overset{d}{\to} N(0, \mathcal{R}^{-1}), \]

where \( \mathcal{R} = \lim_{N,T \to \infty} \mathcal{R} \), and \( \mathcal{R} \) can also be expressed as

\[
\mathcal{R} = \underset{N,T \to \infty}{\text{plim}} \frac{1}{NT} \left( \begin{array}{ccc}
\text{tr}[\mathcal{M} X_1 \mathcal{M}(\mathbb{B}) X_1] & \cdots & \text{tr}[\mathcal{M} X_K \mathcal{M}(\mathbb{B}) X_K]
\end{array} \right).
\]
We next consider the estimation of $h_t$ and $\kappa_i$. Again, we emphasize that the estimation of $h_t$ and $\kappa_i$ is meaningful only when they are identifiable. To identify them, we impose the following IO$'$ conditions.

**The identification conditions IO$'$**: The identification conditions IO hold. In addition, $M_{fd} = \frac{1}{T} \sum_{t=1}^{T} f_t d'_t = 0$ with $f_t = (g'_t, h'_t)'$.

Under the identification conditions IO$'$, we estimate $f_t$ by

$$
\hat{f}_t = (\hat{\Lambda}' \hat{\Sigma}_{ee}^{-1} \hat{\Lambda})^{-1} \hat{\Lambda}' \hat{\Sigma}_{ee}^{-1} (\tilde{Y}_t - \tilde{X}_t \hat{\beta}),
$$

where $\tilde{Y}_t$ is the $t$th column of the matrix $Y \mathcal{M}(\mathcal{D})$, $\tilde{X}_t$ is an $N \times K$ matrix with its $k$th column equal to the $t$th column of the matrix $X_k \mathcal{M}(\mathcal{D})$ and $\hat{\Lambda} = (\hat{\Psi}, \Phi)$ (note $\Phi$ is observable). The estimator $\hat{h}_t$ is included in $\hat{f}_t$. After obtaining the estimators $\hat{\beta}, \hat{\psi}_i, \hat{g}_t, \hat{h}_t$, we estimate $\kappa_i$ by

$$
\hat{\kappa}_i = \left[ \sum_{t=1}^{T} d_t d'_t \right]^{-1} \left[ \sum_{t=1}^{T} d_t (y_{it} - x_{it} \hat{\beta} - \hat{\psi}_i' \hat{g}_t - \phi_i' \hat{h}_t) \right].
$$

The following Theorem states the limiting results on $\hat{h}_t$ and $\hat{\kappa}_i$.

**Theorem 4.4.** Under Assumptions A-D and G, together with the identification conditions IO$'$,

1. Under $\sqrt{N}/T \to 0$, we have

$$
\sqrt{N} (\hat{h}_t - h_t) \overset{d}{\to} N \left( 0, \left[ \lim_{N \to \infty} \frac{1}{N} \Phi \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi \right]^{-1} \right).
$$

2. Under $\sqrt{T}/N \to 0$, we have

$$
\sqrt{T} (\hat{\kappa}_i - \kappa_i) \overset{d}{\to} N \left( 0, \Sigma_{ii} \left[ \lim_{T \to \infty} \frac{1}{T} \sum_{t=1}^{T} d_t d'_t \right]^{-1} \right).
$$

**5. Computing algorithm.** To estimate the model by the maximum likelihood method, we adapt the ECM (Expectation and Conditional Maximization) procedures of [22]. The ECM procedure here can be viewed as the extension of the EM algorithm for the pure factor models considered by [29]. The E-step of the ECM algorithm is the same with the usual EM algorithm, but the M-step is broken into a sequence of maximizations instead of simultaneously maximization over the full parameter space. In the M-step, we split the parameter $\theta = (\beta, \Gamma, \Sigma_{ee}, M_{ff})$ into two blocks, $\theta_1 = (\Gamma, \Sigma_{ee}, M_{ff})$ and $\theta_2 = \beta$, and update $\theta_1^{(k)}$ to $\theta_1^{(k+1)}$ given $\theta_2^{(k)}$ and then update $\theta_2^{(k)}$ to $\theta_2^{(k+1)}$ given $\theta_1^{(k+1)}$, where $\theta^{(k)}$ is the estimated value at the $k$th iteration.
Because different models have different restrictions which lead to different identification conditions, we explain how to estimate each model under different identification conditions.

5.1. Basic Model. In this case, \( M_{ff} = I_r \). So the parameters to be estimated reduce to \( \theta = (\beta, \Gamma, \Sigma_{ee}) \). Let \( \hat{\theta}^{(k)} = (\hat{\beta}^{(k)}, \hat{\Gamma}^{(k)}, \hat{\Sigma}_{ee}^{(k)}) \) be the estimated value at the \( k \)th iteration. We update \( \Gamma^{(k)} \) according to

\[
\Gamma^{(k+1)} = \left[ \frac{1}{T} \sum_{t=1}^{T} E(z_t f_t' | Z, \theta^{(k)}) \right]^{-1} \left( \frac{1}{T} \sum_{t=1}^{T} E(f_t f_t' | Z, \theta^{(k)}) \right)^{-1} ,
\]

where

\[
\frac{1}{T} \sum_{t=1}^{T} E(f_t f_t' | Z, \theta^{(k)}) = I_r - \Gamma^{(k)} (\Sigma_{ee}^{(k)})^{-1} \Gamma^{(k)} ,
\]

\[
+ \Gamma^{(k)} (\Sigma_{ee}^{(k)})^{-1} (I_N \otimes B^{(k)}) M_N (I_N \otimes B^{(k)})' (\Sigma_{ee}^{(k)})^{-1} \Gamma^{(k)} ,
\]

with \( \Sigma_{ee}^{(k)} = \Gamma^{(k)} \Gamma^{(k)'} + \Sigma_{ee}^{(k)} \). We update \( \Sigma_{ee}^{(k)} \) and \( \beta^{(k)} \) according to

\[
\Sigma_{ee}^{(k+1)} = D_g \left\{ \left( I_N (K+1) - \Gamma^{(k+1)} (\Sigma_{ee}^{(k)})^{-1} \right) \times (I_N \otimes B^{(k)}) M_N (I_N \otimes B^{(k)'}), \right. \\
\left. \times \left( \frac{N}{I} \sum_{i=1}^{N} \sum_{t=1}^{T} \tilde{x}_{it} (\Sigma_{tie}^{(k+1)})^{-1} \tilde{x}_{it} \right)^{-1} \right\} ,
\]

\[
\beta^{(k+1)} = \left( \frac{N}{I} \sum_{i=1}^{N} \sum_{t=1}^{T} \tilde{x}_{it} (\Sigma_{tie}^{(k+1)})^{-1} \tilde{x}_{it} \right)^{-1} \left( \frac{N}{I} \sum_{i=1}^{N} \sum_{t=1}^{T} \tilde{x}_{it} (\Sigma_{tie}^{(k+1)})^{-1} (\tilde{y}_{it} - \lambda_t^{(k+1)} f_t^{(k)}) \right) ,
\]

where \( f_t^{(k)} \) is the transpose of the \( t \)-th row of

\[
E^{(k)} = E(E | Z, \theta^{(k)}) = \hat{Z}' (I_N \otimes B^{(k)'})(\Sigma_{ee}^{(k)})^{-1} \Gamma^{(k)}
\]

where \( \hat{Z} = (\hat{z}_1, \hat{z}_2, \ldots, \hat{z}_T) \) with \( \hat{z}_t = z_t - \frac{1}{T} \sum_{s=1}^{T} z_s \); \( D_g(\cdot) \) is the operator that sets the entries of its argument to zeros if the counterparts of \( E(\varepsilon_{it} \varepsilon_t' ) \) are zeros.

Putting together, we obtain \( \hat{\theta}^{(k+1)} = (\Gamma^{(k+1)}, \beta^{(k+1)}, \Sigma_{ee}^{(k+1)}) \).
5.2. Basic model with restrictions. Again we need not estimate $M_{ff}$.

We update $\Gamma^{(k)}$ by two steps. In the first step, we calculate $\Gamma^{(k+1)}$ according to (5.1). Let $T^{-1} \sum_{t=1}^{T} E(g_{1t}g_{1t} | Z, \theta^{(k)})$ be the left-upper $r_1 \times r_1$ submatrix of $T^{-1} \sum_{t=1}^{T} E(f_{1t}f_{1t} | Z, \theta^{(k)})$. Let $Z$ be the $N \times r_1$ matrix whose $i$-th row is the first $r_1$ elements of the $((i-1)(K+1)+1)$-th row of $T^{-1} \sum_{t=1}^{T} E(z_{1t}f_{1t} | Z, \theta^{(k)})$ for $i = 1, 2, \cdots, N$. Then calculate

$$
\Psi^{(k+1)} = Z \left( \frac{1}{T} \sum_{t=1}^{T} E(g_{1t}g_{1t} | Z, \theta^{(k)}) \right)^{-1}
$$

In the second step, $\Gamma^{(k+1)}$ is obtained by replacing the $((i-1)(K+1)+1)$-th row of the first step $\Gamma^{(k+1)}$ with $(\psi_{i}^{(k+1)}', 0_{1 \times r_2})$ for $i = 1, 2, \cdots, N$, where $\psi_{i}^{(k+1)}$ is the transpose of the $i$-th row of $\Psi^{(k+1)}$. We update $\Sigma_{\varepsilon \varepsilon}^{(k)}$ by (5.4) and $\beta^{(k)}$ by

$$
\beta^{(k+1)} = \left( \sum_{i=1}^{N} \sum_{t=1}^{T} \bar{x}_{it}(\Sigma_{i\varepsilon}^{(k)})^{-\frac{1}{2}} \bar{x}_{it} \right)^{-1} \left( \sum_{i=1}^{N} \sum_{t=1}^{T} \bar{x}_{it}(\Sigma_{i\varepsilon}^{(k)})^{-\frac{1}{2}} (\bar{y}_{it} - \psi_{i}^{(k+1)}') g_{1t} \right)
$$

where $\Sigma_{\varepsilon \varepsilon}^{(k)} = \Gamma_{i\varepsilon}^{(k)} \Gamma_{\varepsilon \varepsilon}^{(k)} + \Sigma_{\varepsilon \varepsilon}^{(k)}$, $g_{1t}$ is the transpose of the first $r_1$ elements of the $t$-th row of $F^{(k)}$ with

$$
F^{(k)} = E(F | Z, \theta^{(k)}) = Z' (I_{N} \otimes B^{(k)})^{-1} \Sigma_{\varepsilon \varepsilon}^{(k)}^{-1} \Gamma^{(k)}.
$$

This gives $\theta^{(k+1)} = (\Gamma^{(k+1)}, \beta^{(k+1)}, \Sigma_{\varepsilon \varepsilon}^{(k+1)})$.

5.3. Models with the time-invariant and common regressors. The identification restrictions are such that we only need to estimate $M_{ih}$. Let $\theta^{(k)} = (\beta^{(k)}, \Gamma^{(k)}, \Sigma_{\varepsilon \varepsilon}^{(k)}, M_{ih}^{(k)})$ be the estimator of the $k$th iteration. We update $\Gamma^{(k)}$ by two steps. First calculate $\Gamma^{(k+1)}$ as in (5.1). Notice that $f_{1t} = (g_{1t}', h_{1t}')'$. Once we have obtained $T^{-1} \sum_{t=1}^{T} E(f_{1t}f_{1t} | Z, \theta^{(k)})$, then

$$
\frac{1}{T} \sum_{t=1}^{T} E(g_{1t}g_{1t} | Z, \theta^{(k)}), \quad \frac{1}{T} \sum_{t=1}^{T} E(h_{1t}g_{1t} | Z, \theta^{(k)}), \quad \frac{1}{T} \sum_{t=1}^{T} E(h_{1t}h_{1t} | Z, \theta^{(k)})
$$

are all known. Let $Z$ be the $N \times r_1$ matrix, whose $i$-th row is the first $r_1$ elements of the $((i-1)(K+1)+1)$-th row of $T^{-1} \sum_{t=1}^{T} E(z_{1t}f_{1t} | Z, \theta^{(k)})$ for $i = 1, 2, \cdots, N$. Because the factor loadings $\Phi$ is known, we calculate

$$
\Psi^{(k+1)} = \left[ Z - \Phi \left( \frac{1}{T} \sum_{t=1}^{T} E(h_{1t}g_{1t} | Z, \theta^{(k)}) \right) \right]^{-1} \left[ \frac{1}{T} \sum_{t=1}^{T} E(g_{1t}g_{1t} | Z, \theta^{(k)}) \right]^{-1}
$$
Then we obtain $\Gamma^{(k+1)}$ by replacing the $((i-1)(K+1)+1)$-th row of the first step $\Gamma^{(k)}$ with $\lambda_i^{(k+1)} = (\psi_i^{(k+1)\prime}, \phi_i^{(k+1)\prime})$ for $i = 1, 2, \ldots, N$, where $\psi_i^{(k+1)}$ is the transpose of the $i$th row of $\Psi^{(k+1)}$. We update $\Sigma_{\varepsilon \varepsilon}^{(k)}$ by (5.4), $\beta^{(k)}$ by (5.5) and $M_{hh}^{(k)}$ according to

$$M_{hh}^{(k+1)} = \frac{1}{T} \sum_{t=1}^{T} E(h_t h_t' | Z, \theta^{(k)})$$

This gives $\theta^{(k+1)} = (\Gamma^{(k+1)}, \beta^{(k+1)}, \Sigma_{\varepsilon \varepsilon}^{(k+1)}, M_{hh}^{(k+1)})$.

For the model with both time-invariant and common regressors, we first post-multiply $\mathcal{M}(\mathbb{D}) = I_T - \mathbb{D}'(\mathbb{D}'\mathbb{D})^{-1}\mathbb{D}'$ on the data matrix $Z$ and apply the preceding procedures on $Z\mathcal{M}(\mathbb{D})$. Then all the estimators are obtained.

The above iteration continues until $\|\theta^{(k+1)} - \theta^{(k)}\|$ is smaller than a preset error tolerance. For the initial values, the iterated PC estimators of [8] are used.

6. Finite sample properties. In this section, we consider the finite sample properties of the MLE. Data are generated according to

$$y_{it} = \alpha_i + x_{it1}\beta_1 + x_{it2}\beta_2 + \psi_{i}^\prime g_{t} + \phi_{i}^\prime h_{t} + \kappa_{i}^\prime d_{t} + \epsilon_{it},$$

$$x_{itk} = \mu_{ik} + \gamma_{ik}^g g_{t} + \gamma_{ik}^h h_{t} + \gamma_{ik}^d d_{t} + v_{itk}, \quad k = 1, 2.$$  

The dimensions of $g_t, h_t, d_t$ are each fixed to 1. We set $\beta_1 = 1$ and $\beta_2 = 2$.

We consider four types of DGP (data generating process), which correspond to the four models considered in the paper.

**DGP1**: $\phi_i, \kappa_i, \gamma_{ik}^h$ and $\gamma_{ik}^d$ are fixed to zeros; $\alpha_i, \mu_{ik}, \psi_i$ and $g_t$ are generated from $N(0, 1)$ and $\gamma_{ik}^g = \psi_i + N(0, 1)$.

**DGP2**: $\phi_i, \kappa_i$ and $\gamma_{ik}^d$ are fixed to zeros; $\alpha_i, \mu_{ik}, \psi_i$, $\gamma_{ik}^h$, $g_t$ and $h_t$ are generated from $N(0, 1)$; $\gamma_{ik}^g = \psi_i + N(0, 1)$.

**DGP3**: $\kappa_i$ and $\gamma_{ik}^d$ are fixed to zeros; $\alpha_i, \mu_{ik}, \psi_i, \phi_i, g_t$ and $h_t$ are generated from $N(0, 1)$; $\gamma_{ik}^g = \psi_i + N(0, 1)$ and $\gamma_{ik}^h = \phi_i + N(0, 1)$. Here $\phi_i$ is observable.

**DGP4**: $\alpha_i, \mu_{ik}, \psi_i, \phi_i, \kappa_i, g_t$ and $h_t$ are generated from $N(0, 1)$; $d_t = 1 + N(0, 1)$, $\gamma_{ik}^g = \psi_i + N(0, 1)$, $\gamma_{ik}^h = \phi_i + N(0, 1)$ and $\gamma_{ik}^d = \kappa_i + N(0, 1)$. Here $\phi_i$ and $d_t$ are observable.

Using the method of writing (2.2), we can rewrite (6.1) as

$$(I_N \otimes B)z_{it} = \mu + L_{si} + \varepsilon_{it}$$
where \( \varsigma = g_t \) for DGP1; \( \varsigma = (g_t, h_t)' \) for DGP2 and DGP3; \( \varsigma = (g_t, h_t, d_t)' \) for DGP4 and \( L \) is the corresponding loadings matrix. Let \( \ell_i' \) be the \( i \)th row of \( L \). We generate the cross sectional heteroscedasticity \( \Xi \), an \( N(K + 1) \times 1 \) vector, according to
\[
\Xi_i = \frac{\eta_i}{1 - \eta_i} \ell_i't_i, \quad i = 1, 2, \ldots, N(K + 1)
\]
where \( \eta_i \) is drawn from \( U[u, 1 - u] \) with \( u = 0.1 \). A similar way of generating heteroscedasticity is also used in [13] and [16]. Let \( \Upsilon = \text{diag}(\Upsilon_1, \Upsilon_2, \ldots, \Upsilon_N) \) be an \( N(K + 1) \times N(K + 1) \) block diagonal matrix, in which \( \Upsilon_i = \text{diag}\{1, (M_i'M_i)^{-1/2}M_i\} \) with \( M_i \) being a \( K \times K \) standard normal random matrix for each \( i \). Once \( \Upsilon \) is generated, the error term \( \varepsilon_t \), which is defined as \( (\varepsilon_{1t}, \varepsilon_{2t}, \ldots, \varepsilon_{Nt})' \) with \( \varepsilon_{it} = (\epsilon_{it}, v_{1it}, v_{2it})' \), is calculated by \( \varepsilon_t = \sqrt{\text{diag}(\Xi)}\Upsilon\epsilon_t \), where \( \epsilon_t \) is an \( N(K + 1) \times 1 \) vector with all its elements being \( iid \ N(0, 1) \). Once \( \varepsilon_t \) is obtained, we use
\[
z_t = (I_N \otimes B)^{-1}(\mu + L\varsigma_t + \varepsilon_t)
\]
to yield the observable data.

Tables 1-4 report the simulation results based on 1000 repetitions. Bias and root mean square error (RMSE) are computed to measure the performance of the estimators. For the purpose of comparison, we also report the performance of the within-group (WG) estimators and Bai’s iterated principal components estimators (PC).

It is seen from the simulations that the WG estimators are inconsistent. The bias of the WG estimators shows no signs of decreasing as the sample size grows. The iterated PC estimators are consistent, but biased. As the sample size becomes large, the bias decreases noticeably. However, when the sample size is moderate, the bias of the iterated PC estimators is still pronounced. In comparison, the ML estimators are consistent and unbiased. For all the sample sizes, the biases of the ML estimators are very small and negligible. In addition, the RMSEs of the ML estimators are always the smallest among the three estimators, illustrating the efficiency of the ML method. The same pattern is observed for all of the four models considered.

7. Conclusion. This paper considers estimating panel data models with interactive effects, in which explanatory variables are correlated with the unobserved effects. Standard panel data methods (such as the within-group estimator) are not suitable for this type of models. We study the maximum likelihood method and provide a rigorous analysis for the asymptotic theory. While the analysis is difficult, the limiting distributions of the MLE are simple and have intuitive interpretations. The maximum likelihood method can
incorporate parameters restrictions to gain efficiency, a useful feature in view of the large number of parameters under large $N$ and large $T$. We analyze the restrictions via the Lagrange multiplier approach. This approach can reveal what kinds of restrictions are binding and lead to efficiency gain. We allow the model to include time invariant regressors and common regressors. The coefficients of the time invariant regressors are time dependent, and the coefficients of the common regressors are cross-section dependent. This is a sensible way for modelling the effects of such variables in panel data context and fits naturally into the framework of interactive effects. The likelihood method is easy to implement and perform very well, as demonstrated by the Monte Carlo simulations.
Table 1: The performance of WG, PC and ML estimators in the basic model

<table>
<thead>
<tr>
<th></th>
<th>WG</th>
<th></th>
<th>PC</th>
<th></th>
<th>MLE</th>
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<tr>
<td></td>
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<td>Bias</td>
<td>RMSE</td>
</tr>
<tr>
<td></td>
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</tr>
<tr>
<td>$N$</td>
<td>$T$</td>
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Table 2: The performance of WG, PC and ML estimators in the model with zero restrictions.

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<th>MLE</th>
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<tr>
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<td>Bias</td>
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</tr>
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Table 3: The performance of WG, PC and ML estimators in the model with time-invariant regressors

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<th>$\beta_1$ RMSE</th>
<th>$\beta_2$ Bias</th>
<th>$\beta_2$ RMSE</th>
<th>$\beta_1$ Bias</th>
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Table 4: The performance of WG, PC and ML estimators in the model with time-invariant regressors and common regressors

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<th>$\beta_2$ RMSE</th>
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References.


metrica, 46(1), 69–85.


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SUPPLEMENT TO “THEORY AND METHODS OF PANEL DATA MODELS WITH INTERACTIVE EFFECTS”

This supplement provides the detailed proofs for the propositions and theorems in the main text.

We first introduce the symbols to be used in this supplement.

Table 1: The symbols used in the supplement

<table>
<thead>
<tr>
<th>Symbol</th>
<th>Definition</th>
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<td>$H$</td>
<td>$(\Gamma'\Sigma_{ee^{-1}}\Gamma)^{-1}$</td>
</tr>
<tr>
<td>$\hat{H}$</td>
<td>$N \cdot \hat{H}$</td>
</tr>
<tr>
<td>$G$</td>
<td>$(M_{ff}^{-1} + \Gamma'\Sigma_{ee^{-1}}\Gamma)^{-1}$</td>
</tr>
<tr>
<td>$\hat{G}$</td>
<td>$(\hat{M}<em>{ff}^{-1} + \hat{\Gamma}'\hat{\Sigma}</em>{ee^{-1}}\hat{\Gamma})^{-1}$</td>
</tr>
<tr>
<td>$A$</td>
<td>$(\hat{\Gamma} - \Gamma)'\hat{\Sigma}_{ee^{-1}}\hat{\Gamma}$</td>
</tr>
<tr>
<td>$\hat{i}_t$</td>
<td>$(\hat{x}'<em>t, 0</em>{K \times K})'$</td>
</tr>
<tr>
<td>$\omega_{pq}$</td>
<td>$N^{-1} \sum_{i=1}^{N} \gamma_{ip} \gamma_{iq}^{-1}$</td>
</tr>
<tr>
<td>$v_p$</td>
<td>$N^{-1} \sum_{i=1}^{N} \lambda_i \Sigma_i^{-1} \gamma_{ip}$</td>
</tr>
<tr>
<td>$\xi_t$</td>
<td>$N^{-1} \sum_{i=1}^{N} \Gamma_i \Sigma_i^{-1} \epsilon_i$</td>
</tr>
<tr>
<td>$\chi_t$</td>
<td>$N^{-1} \sum_{i=1}^{N} \Gamma_i \Sigma_i^{-1} \epsilon_i$</td>
</tr>
</tbody>
</table>

From $(A + B)^{-1} = A^{-1} - A^{-1}B(A + B)^{-1}$, we have $\hat{H} = \hat{G}(I - \hat{M}_{ff}^{-1}\hat{G})^{-1}$.

From $\Sigma_{ee} = \Gamma M_{ff}\Gamma' + \Sigma_{ee}$, we have

(S.1) $\Sigma_{ee}^{-1} = \Sigma_{ee}^{-1} - \Sigma_{ee}^{-1}[\Gamma(M_{ff}^{-1} + \Gamma'\Sigma_{ee}^{-1}\Gamma)^{-1}\Gamma'\Sigma_{ee}^{-1}] = \Sigma_{ee}^{-1} - \Sigma_{ee}^{-1}\Gamma GT'\Sigma_{ee}^{-1}$,

and

(S.2) $\hat{\Gamma}'\hat{\Sigma}_{ee}^{-1} = \hat{\Gamma}'\hat{\Sigma}_{ee}^{-1} - \hat{\Gamma}'\hat{\Sigma}_{ee}^{-1}\hat{\Gamma}(\hat{M}_{ff}^{-1} + \hat{\Gamma}'\hat{\Sigma}_{ee}^{-1}\hat{\Gamma})^{-1}\hat{\Gamma}'\hat{\Sigma}_{ee}^{-1} = \hat{M}_{ff}^{-1}\hat{G}\hat{\Gamma}'\hat{\Sigma}_{ee}^{-1}$.

APPENDIX A: TECHNICAL MATERIALS FOR SECTION 2

APPENDIX A.1: PROOF OF PROPOSITION 2.1

The following six lemmas are useful for the proof of Proposition 2.1. The proofs of these six lemmas do not involve the special structure of the factor loadings and the identification conditions, they still hold in the context of Sections 3 and 4.

**Lemma A.1.** Let $Y$ and $Z$ be two real symmetric matrices, which have the orthogonal reduction

$$P'YP = D_Y \quad Q'ZQ = D_Z$$

where $D_Y$ ($D_Z$) is a diagonal matrix with its diagonal elements arranged in an increasing (descending) order. Then we have $\text{tr}(YZ) \geq \text{tr}(D_Y D_Z)$. 
Lemma A.2. Let $Q$ be an $r \times r$ matrix satisfying

$$QQ' = I_r$$
$$Q'VQ = D$$

where $V$ is an $r \times r$ diagonal matrix with strictly positive and distinct elements, arranged in decreasing order, and $D$ is also diagonal. Then $Q$ must be a diagonal matrix with elements either $-1$ or $1$ and $V = D$.

Lemma A.1 is given in Theobald (1975) and Lemma A.2 is given in [10].

To prove Proposition 2.1, we use a superscript "*" to denote the true parameters, for example $\Gamma^*$, $\Sigma^*_{ee}$, etc. The variables without the superscript "*" denote the function arguments (input variables) in the likelihood function.

Lemma A.3. Let $\theta = (\beta, \Gamma, \Sigma_{ee})$ and $\Theta$ be the parameter set such that Assumption D is satisfied and $\beta$ is in a compact set. Then we have

$$(a) \sup_{\theta \in \Theta} \frac{1}{N} \text{tr}\left[ B \Gamma^* \frac{1}{T} \sum_{t=1}^{T} f_t^* \varepsilon_t' B' \Sigma^{-1} \right] = o_p(1)$$

$$(b) \sup_{\theta \in \Theta} \frac{1}{N} \text{tr}\left[ B \frac{1}{T} \sum_{t=1}^{T} (\dot{\varepsilon}_t \dot{\varepsilon}_t' - \Sigma^*_{ee}) B' \Sigma^{-1} \right] = o_p(1)$$

where $B = (I_N \otimes B)(I_N \otimes B^*)^{-1}$.

Proof of Lemma A.3. Consider (a). Recall

$$B = \begin{pmatrix} 1 & -\beta' \\ 0 & I_K \end{pmatrix} \quad \text{so} \quad B^{-1} = \begin{pmatrix} 1 & \beta' \\ 0 & I_K \end{pmatrix} \quad \text{. Let} \quad L = \begin{pmatrix} 0 & (\beta - \beta^*)' \\ 0 & 0 \end{pmatrix} \quad \text{.}$$

Then we have $B = (I_N \otimes B)(I_N \otimes B^*)^{-1} = I_N \otimes (I_{K+1} - L)$. Now the left hand side of (a) is equal to

$$\text{tr}\left[ \frac{1}{NT} \sum_{t=1}^{T} \Gamma^* f_t^* \varepsilon_t' \Sigma^{-1} \right] - 2 \text{tr}\left[ \frac{1}{NT} \sum_{t=1}^{T} (I_N \otimes L) \Gamma^* f_t^* \varepsilon_t' \Sigma^{-1} \right]$$

$$+ \text{tr}\left[ \frac{1}{NT} \sum_{t=1}^{T} (I_N \otimes L) \Gamma^* f_t^* \varepsilon_t' (I_N \otimes L) \Sigma^{-1} \right]$$

Consider the first term of the above equation. Notice that

(A.1) $\sum_{i=1}^{N} \left\| H^{1/2} \Gamma_i \Sigma_i^{-1/2} \right\| ^2 = \sum_{i=1}^{N} \text{tr}\left( H^{1/2} \Gamma_i \Sigma_i^{-1} \Gamma_i' H^{1/2} \right)$
\[ \text{tr}[H^{1/2} \left( \sum_{i=1}^{N} \Gamma_i \Sigma_{ii}^{-1} \Gamma_i' \right) H^{1/2}] = \text{tr}(H^{1/2} H^{-1} H^{1/2}) = r \]

The last equation is due to the definition of \( H \). So Lemma A.2 of [10] holds in the present context. Using their arguments, we can prove sup\( \theta \in \Theta \text{tr} \left[ \frac{1}{NT} \sum_{t=1}^{NT} \Gamma^* f_t^* \Sigma_{ii}^{-1} \right] = o_p(1). \) The second term can be proved to be \( o_p(1) \) uniformly on \( \Theta \) similarly as the first since we can treat \((I_N \otimes L) \Gamma^*\) as a new \( \Gamma^* \) because of the boundedness of \( \beta \). Now consider the third term, which is equivalent to

\[
\sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*)(\beta_q - \beta_q^*) \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iie}^{-1} v_{itp} \gamma_{iq} f_t^* \right)
\]

\[
- \sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*)(\beta_q - \beta_q^*) \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} v_{itp} \Sigma_{iie}^{-1} \lambda_j^* G \Sigma_{jje}^{-1} \gamma_{jq} f_t^* \right) = a_1 + a_2,
\]

Consider \( a_1 \). By the Cauchy-Schwarz inequality, the term \( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iie}^{-1} v_{itp} \gamma_{iq} f_t^* \) is bounded by

\[
\left( \frac{1}{N} \sum_{i=1}^{N} ||\Sigma_{iie}^{-1}||^2 \cdot ||\gamma_{iq}||^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \frac{1}{T} \sum_{t=1}^{T} ||v_{itp}||^2 \right)^{1/2}
\]

The first factor is \( O_p(1) \) due to Assumptions C and D. The second factor is \( O_p(T^{-1/2}) \). So the above expression is \( O_p(T^{-1/2}) \). Noticing \( K \) is a finite number and \( ||\beta_p - \beta_p^*|| \) is bounded, we have \( a_1 = O_p(T^{-1/2}) \).

Consider \( a_2 \). We will show

\[
(A.2) \quad \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} v_{itp} \Sigma_{iie}^{-1} \lambda_j^* G \Sigma_{jje}^{-1} \gamma_{jq} f_t^* = O_p(T^{-1/2})
\]

By the Cauchy-Schwarz inequality, the left hand side above is bounded in norm by

\[
\left( \frac{1}{N} \sum_{i=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} ||v_{itp}||^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \Sigma_{iie}^{-1} \lambda_j^* H^{1/2} \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} ||\gamma_{jq}||^2 \right)^{1/2} \right) \leq \|\mathcal{U}\|
\]

with \( \mathcal{U} = (I + H^{1/2} M_f^{-1} H^{1/2})^{-1} \). The first factor is \( O_p(T^{-1/2}) \), the third is \( O(1) \) and \( \|\mathcal{U}\| \leq 1 \). Consider the term \( \sum_{i=1}^{N} \Sigma_{iie}^{-1} \lambda_j^* H^{1/2} \), which, due to the boundedness of \( \Sigma_{iie} \) by Assumption D, is bounded by \( C \sum_{i=1}^{N} \Sigma_{iie}^{-1} \lambda_j^* H^{1/2} \), which is further bounded by \( C \sum_{i=1}^{N} \Sigma_{iie}^{-1} \Gamma_i^* H^{1/2} \| = Cr \) by (A.1). Given
these results, we prove (A.2). Furthermore \( |\beta_p - \beta_p^*| \) is bounded, so we have \( a_2 = O_p(T^{-1/2}) \). Then (a) follows.

Consider (b). The left hand side of (b) can be written as

\[
\begin{align*}
\text{tr} \left[ \frac{1}{NT} \sum_{t=1}^T (\varepsilon_t \varepsilon'_t - \Sigma_{ee}) \Sigma_{zz}^{-1} \right] &- \text{tr} \left[ \frac{1}{N} \varepsilon \varepsilon' \Sigma_{zz}^{-1} \right] - 2 \text{tr} \left[ \frac{1}{NT} \sum_{t=1}^T (I_N \otimes L) (\varepsilon_t \varepsilon'_t - \Sigma_{ee}) \Sigma_{zz}^{-1} \right] \\
+ 2 \text{tr} \left[ \frac{1}{N} (I_N \otimes L) \varepsilon \varepsilon' \Sigma_{zz}^{-1} \right] &+ \text{tr} \left[ \frac{1}{NT} \sum_{t=1}^T (I_N \otimes L) (\varepsilon_t \varepsilon'_t - \Sigma_{ee}) (I_N \otimes L') \Sigma_{zz}^{-1} \right]
\end{align*}
\]

(A.3) \(- \text{tr} \left[ \frac{1}{N} (I_N \otimes L) \varepsilon \varepsilon' (I_N \otimes L') \Sigma_{zz}^{-1} \right] = b_1 - b_2 - 2b_3 + 2b_4 + b_5 - b_6\) where \( L \) is defined in result (a). The term \( b_1 \) is equal to

\[
\text{tr} \left[ \frac{1}{NT} \sum_{t=1}^T (\varepsilon_t \varepsilon'_t - \Sigma_{ee}^*) \Sigma_{zz}^{-1} \right] - \text{tr} \left[ \frac{1}{NT} \sum_{t=1}^T \Gamma' \Sigma_{ee}^{-1} (\varepsilon_t \varepsilon'_t - \Sigma_{ee}^*) \Sigma_{zz}^{-1} \Gamma G \right]
\]

The first term is \( \text{tr} \left[ \frac{1}{N} \sum_{i=1}^N \sum_{t=1}^T (\varepsilon_t \varepsilon'_t - \Sigma_{ii}^*) \Sigma_{zz}^{-1} \right] \), which is bounded in norm by (ignore the trace)

\[
\left( \frac{1}{N} \sum_{i=1}^N \| \Sigma_{ii}^{-1} \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^N \left\| \frac{1}{T} \sum_{t=1}^T (\varepsilon_t \varepsilon'_t - \Sigma_{ii}^*) \right\|^2 \right)^{1/2} = O_p(T^{-1/2})
\]

uniformly on \( \Theta \). The second term is equal to the trace of

\[
\frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N H^{1/2} \Gamma_i \Sigma_{ii}^{-1} \frac{1}{T} \sum_{t=1}^T \varepsilon_{it} \varepsilon_{jt} \Gamma_{jj}^{-1} H^{1/2} U
\]

with \( \varepsilon_{it} = \varepsilon_t \varepsilon_{jt} - E(\varepsilon_t \varepsilon_{jt}) \) and \( U = (I_r + H^{1/2} M^{-1} H^{1/2})^{-1} \). The above expression is bounded in norm by

\[
\left( \sum_{i=1}^N \| \Sigma_{ii}^{-1} \Gamma_{ii} H^{1/2} \|^2 \right) \left( \frac{1}{N^2} \sum_{i=1}^N \sum_{j=1}^N \left\| \frac{1}{T} \sum_{t=1}^T \varepsilon_{it} \varepsilon_{jt} \right\|^2 \right)^{1/2} \| U \|
\]

The term \( \sum_{i=1}^N \| \Sigma_{ii}^{-1} \Gamma_{ii} H^{1/2} \|^2 \) is bounded by \( \sum_{i=1}^N \| \Sigma_{ii}^{-1} \Gamma_{ii} H^{1/2} \|^2 = Cr \) by (A.1). So the above expression is \( O_p(T^{-1/2}) \) uniformly on \( \Theta \). This shows that \( b_1 \) is \( O_p(T^{-1/2}) \) uniformly on \( \Theta \).

Consider \( b_2 \). By \( \Sigma_{zz}^{-1} \leq \Sigma_{ee}^{-1} \), \( b_2 \) is bounded by \( \text{tr} \left[ \frac{1}{N} \varepsilon \varepsilon' \Sigma_{zz}^{-1} \right] \), which is equivalent to

\[
\frac{1}{T} \sum_{i=1}^N \left( \frac{1}{\sqrt{T}} \sum_{t=1}^T \varepsilon'_t \right) \Sigma_{ii}^{-1} \left( \frac{1}{\sqrt{T}} \sum_{s=1}^T \varepsilon_{is} \right)
\]
which is $O_p(T^{-1})$ uniformly on $\Theta$ due to the boundedness of $\Sigma_{ii}$. So $b_2$ is $O_p(T^{-1})$ uniformly on $\Theta$.

Consider $b_3$, which is equal to

$$
\sum_{p=1}^{K} \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ii}^{-1} e_{it} v_{itp}
$$

$$
-\text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i \Sigma_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [L_{\varepsilon_{it}^2} - E(L_{\varepsilon_{it}^2})] \Sigma_{ji}^{-1} \Gamma_j G \right]
$$

The first term of the above expression is bounded in norm by

$$
\sum_{p=1}^{K} |\beta_p - \beta_p^*| \left( \frac{1}{N} \sum_{i=1}^{N} \Sigma_{ii}^{-2} \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left| \frac{1}{T} \sum_{t=1}^{T} e_{it} v_{itp} \right|^2 \right)^{1/2}
$$

which is $O_p(T^{-1/2})$ due to the boundedness of $\Sigma_{ii}$ and $\beta$. The term inside the trace operator of the second term can be written alternatively as

$$
\sum_{p=1}^{K} (\beta_p - \beta_p^*) \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} H^{1/2} \Gamma_i \Sigma_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [L_{\varepsilon_{it}^2} - E(L_{\varepsilon_{it}^2})] \Sigma_{ji}^{-1} \Gamma_j H^{1/2} U
$$

with

$$
L_{\varepsilon_{it}^2} = \begin{pmatrix}
  v_{itp} e_{jt} & v_{itp} v_{jt1} & \cdots & v_{itp} v_{jTK} \\
  0 & 0 & \cdots & 0 \\
  \vdots & \vdots & \ddots & \vdots \\
  0 & 0 & \cdots & 0
\end{pmatrix}
$$

which is bounded in norm by

$$
\sum_{p=1}^{K} |\beta_p - \beta_p^*| \left( \frac{N}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} ||\Sigma_{ji}^{-1} \Gamma_j H^{1/2}||^2 \right) \left[ \frac{1}{N^4} \sum_{i=1}^{N} \sum_{j=1}^{N} ||\frac{1}{T} \sum_{t=1}^{T} [L_{\varepsilon_{it}^2} - E(L_{\varepsilon_{it}^2})] ||^2 \right]^{1/2} ||U||
$$

with $U = (I_T + H^{1/2} M_j^{-1} H^{1/2})^{-1}$. The term $\frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \frac{1}{T} \sum_{t=1}^{T} [L_{\varepsilon_{it}^2} - E(L_{\varepsilon_{it}^2})] ||^2$ is $O_p(T^{-1})$. From this, the above term is $O_p(T^{-1/2})$ uniformly on $\Theta$. This gives $b_3 = O_p(T^{-1/2})$ uniformly on $\Theta$.

Consider $b_4$, which is equal to

$$
\frac{1}{T} \sum_{p=1}^{K} (\beta_p - \beta_p^*) \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} e_{itp} \left( \frac{1}{\sqrt{T}} \sum_{s=1}^{T} v_{jtp} \right)
$$

$$
-\text{tr} \left[ \frac{1}{N} \sum_{i=1}^{N} \frac{1}{T} \sum_{t=1}^{T} [L_{\varepsilon_{it}^2} - E(L_{\varepsilon_{it}^2})] \Sigma_{ji}^{-1} \frac{1}{T} \sum_{t=1}^{T} e_{itp} \right] \left( \frac{1}{T} \sum_{t=1}^{T} e_{jt} \right)^{1/2} \Sigma_{ji}^{-1} \Gamma_j G
$$
The first term is \( O_p(T^{-1}) \) uniformly on \( \Theta \). The second term inside the trace operator is bounded in norm by

\[
\sum_{p=1}^{K} |\beta_p - \beta_p^*| \left( \sum_{i=1}^{N} \|H_1^{1/2} \gamma_i \Sigma_{ii}^{-1}\|^2 \right) \left( \frac{1}{N} \sum_{i=1}^{N} \left| \frac{1}{T} \sum_{t=1}^{T} v_{ip} \right|^2 \right)^{1/2} 
\times \left( \frac{1}{N} \sum_{j=1}^{N} \| \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{jt} \|^2 \right)^{1/2} \|U\|
\]

with \( U = (I_r + H_1^{1/2} M_f^{-1} H_1^{1/2})^{-1} \), which is \( O_p(T^{-1}) \) uniformly on \( \Theta \). So \( b_4 \) is \( O_p(T^{-1}) \) uniformly on \( \Theta \).

Consider \( b_5 \), which is equal to

\[
\sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*) (\beta_q - \beta_q^*) \frac{1}{NT} \sum_{i=1}^{N} \sum_{l=1}^{T} \Sigma_{ii}^{-1} [v_{ip}v_{ilq} - E(v_{ip}v_{ilq})] \]
\[- \sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*) (\beta_q - \beta_q^*) \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{T} \lambda_j \Sigma_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [v_{ip}v_{jqt} - E(v_{ip}v_{jqt})] \Sigma_{jj}^{-1} \lambda_j G \right] \]

The first term is bounded in norm by

\[
\sum_{p=1}^{K} \sum_{q=1}^{K} \left\{ |\beta_p - \beta_p^*| \cdot |\beta_q - \beta_q^*| \left( \frac{1}{N} \sum_{i=1}^{N} \Sigma_{ii}^{-2} \right) \right\}^{1/2} 
\times \left( \frac{1}{N} \sum_{i=1}^{T} \left( \frac{1}{T} \sum_{t=1}^{T} [v_{ip}v_{itq} - E(v_{ip}v_{itq})] \right)^2 \right)^{1/2}
\]

which is \( O_p(T^{-1/2}) \) by the boundedness of \( \beta \). For the second term, the trace is equal to the trace of

\[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{T} H_1^{1/2} \gamma_i \Sigma_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [v_{ip}v_{jqt} - E(v_{ip}v_{jqt})] \Sigma_{jj}^{-1} \lambda_j H_1^{1/2} \mathcal{U}
\]

which is bounded in norm by

\[
\left( \sum_{j=1}^{N} \| \Sigma_{jj}^{-1} \lambda_j H_1^{1/2} \|^2 \right) \left( \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} [v_{ip}v_{jqt} - E(v_{ip}v_{jqt})] \right)^2 \right)^{1/2} \| \mathcal{U} \|
\]

where \( \mathcal{U} = (I_r + H_1^{1/2} M_f^{-1} H_1^{1/2})^{-1} \). The above is \( O_p(T^{-1/2}) \) uniformly on \( \Theta \). So the second term is \( O_p(T^{-1/2}) \) because \( \beta \) is bounded. Thus \( b_5 = O_p(T^{-1/2}) \) uniformly on \( \Theta \).
Consider $b_6$. Since $\Sigma_{zz}^{-1} \leq \Sigma_{ee}^{-1}$, we have

$$b_6 \leq \text{tr} \left[ \frac{1}{N} (I_N \otimes L) \tilde{z}' (I_N \otimes L') \Sigma_{ee}^{-1} \right] = \frac{1}{N} \sum_{j=1}^N \Sigma_{je}^{-1} \left( \sum_{p=1}^K (\beta_p - \beta_p^*) \frac{1}{T} \sum_{t=1}^T v_{jt} \right)^2$$

which is $O_p(T^{-1})$ uniformly on $\Theta$.

Summarizing all the results, we obtain (b).

**Lemma A.4.** Under Assumptions A-D, if $\| \hat{\beta} - \beta^* \| = o_p(1)$, then

$$\frac{1}{N} \text{tr} \left[ \hat{B} \hat{\Gamma}^* M^*_f \hat{\Gamma}^* \hat{\Sigma}_{zz}^{-1} \right] = \frac{1}{N} \text{tr} \left[ \hat{\Gamma}^* M^*_f \hat{\Gamma}^* \hat{\Sigma}_{zz}^{-1} \right] + o_p(1)$$

where $\hat{B} = (I_N \otimes \hat{B})(I_N \otimes B^*)^{-1} = I_N \otimes (\hat{B}B^{-1})$.

**Proof of Lemma A.4.** Using the notations in result (a) of Lemma A.3, the left hand side is equal to

$$\frac{1}{N} \text{tr} \left[ \hat{\Gamma}^* M^*_f \hat{\Gamma}^* \hat{\Sigma}_{zz}^{-1} \right] - 2 \text{tr} \left[ (I_N \otimes \hat{L}) \hat{\Gamma}^* M^*_f \hat{\Gamma}^* \hat{\Sigma}_{zz}^{-1} \right]$$

$$+ \text{tr} \left[ (I_N \otimes \hat{L}) \hat{\Gamma}^* M^*_f \hat{\Gamma}^* (I_N \otimes \hat{L}') \hat{\Sigma}_{zz}^{-1} \right]$$

Given $\| \hat{\beta} - \beta^* \| = o_p(1)$, it suffices to prove

$$-2 \text{tr} \left[ (I_N \otimes \hat{L}) \hat{\Gamma}^* M^*_f \hat{\Gamma}^* \hat{\Sigma}_{zz}^{-1} \right] + \text{tr} \left[ (I_N \otimes \hat{L}) \hat{\Gamma}^* M^*_f \hat{\Gamma}^* (I_N \otimes \hat{L}') \hat{\Sigma}_{zz}^{-1} \right] = O_p(\| \hat{\beta} - \beta^* \|)$$

We use $c_1$ and $c_2$ to denote the two terms of the above expression. Consider $c_1$, which is equal to (omitting -2)

$$\text{tr} \left[ \sum_{p=1}^K (\hat{\beta}_p - \beta_p^*) \frac{1}{N} \sum_{i=1}^N \hat{\Sigma}_{ij}^{-1} \lambda_i^* \gamma_{ijp} \hat{\Sigma}_{ij}^{-1} \right]$$

$$- \text{tr} \left[ \sum_{p=1}^K (\hat{\beta}_p - \beta_p^*) \frac{1}{N} \sum_{i=1}^N \hat{\Sigma}_{ij}^{-1} \hat{\lambda}_i \gamma_{ijp} \hat{M}_f \right. \sum_{j=1}^N \Gamma_j^* \hat{\Sigma}_{ij}^{-1} \hat{G}$$

The first term of the above equation is $O_p(\| \hat{\beta} - \beta^* \|)$ by the boundedness of $\lambda_i^*, \gamma_{ijp}^*, \hat{\Sigma}_{ij}$. The second term is equivalent to

$$\text{tr} \left[ \sum_{p=1}^K (\hat{\beta}_p - \beta_p^*) \frac{1}{N} \sum_{i=1}^N \hat{H}^{1/2} \hat{\Sigma}_{ij}^{-1} \hat{\lambda}_i \gamma_{ijp} \hat{M}_f \right. \sum_{j=1}^N \Gamma_j^* \hat{\Sigma}_{ij}^{-1} \hat{G} \hat{H}^{1/2} \hat{U}$$
with \( U = (I + \hat{H}^{1/2} \hat{M}_{jj}^{-1} \hat{H}^{1/2})^{-1} \). Ignore the trace, the term is bounded in norm by
\[
\sum_{p=1}^{K} \left\{ |\hat{\beta}_p - \beta_p^*| \left( \sum_{i=1}^{N} \left\| \hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \right\|^2 \right) \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \gamma_{ip} \right\|^2 \right)^{1/2} \right\}^2
\]
where we use the fact that \( \| \hat{H}^{1/2} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1} \| \leq \| \hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \| \). The above expression is \( O_p(\| \hat{\beta} - \beta^* \|^2) \).

Consider \( c_2 \). Since \( \hat{\Sigma}_{ii}^{-1} \leq \hat{\Sigma}_{ee}^{-1} \), we only need to explore the term \( \text{tr}[(I_N \otimes \hat{L}) \hat{\Gamma}^* \hat{M}_{jj}^* \hat{\Gamma}^*(I_N \otimes \hat{L}^*) \hat{\Sigma}_{ee}^{-1}] \), which is equivalent to
\[
\text{tr} \left( \sum_{p=1}^{K} \sum_{q=1}^{Q} (\hat{\beta}_p - \beta_p^*)(\hat{\beta}_q - \beta_q^*) \frac{1}{N} \sum_{j=1}^{N} \hat{\Sigma}_{jj}^{-1} \gamma_{jp}^* \gamma_{jq}^* \hat{M}_{jj}^* \right).
\]
The above is \( O_p(\| \hat{\beta} - \beta^* \|^2) \) due to the boundedness of \( \gamma_{jp}^*, \hat{M}_{jj}^*, \hat{\Sigma}_{jj} \). So we have \( c_2 = O_p(\| \hat{\beta} - \beta^* \|^2) \).

Combining the results on \( c_1 \) and \( c_2 \), we have Lemma A.4.

**Lemma A.5.** **Under Assumptions A-D,**

(a) \( H_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_t^* \right] (I - A) = \| N^{1/2} \hat{H}^{1/2} \|^2 \cdot O_p(T^{-1/2}) \)

(b) \( \hat{H} \left[ \frac{1}{T} \sum_{t=1}^{T} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \eps_{ij,t} \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j \hat{H} = \| N^{1/2} \hat{H}^{1/2} \|^2 \cdot O_p(T^{-1/2}) \right) \)

(c) \( \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Sigma}_{ii} - \Sigma_{ii}^*) \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i \hat{H} \)
\[
= \| N^{1/2} \hat{H}^{1/2} \|^2 \cdot O_p \left( N^{-1/2} \left[ \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 \right]^{1/2} \right)
\]
where \( \eps_{ij,t} = \eps_{ij} e_{ij,t} - E(\eps_{ij} e_{ij,t}) \). The symbols of \( A, \hat{H}_N \) and \( \hat{\chi}_t \) are defined in Table 1.

**Proof of Lemma A.5.** Consider (a). By the definitions of \( A, \hat{H}_N \) and \( \hat{\chi}_t \), the left hand side of (a) is equivalent to
\[
\hat{H} \left[ \frac{1}{T} \sum_{t=1}^{T} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \eps_{ij,t} f_t^* \sum_{j=1}^{N} \gamma_{ij}^* \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j \hat{H} \right]
\]
By the Cauchy-Schwarz inequality, term $\sum_{j=1}^{N} \Gamma_j^* \hat{\Sigma}_j^{-1/2} \hat{\Sigma}_j' \hat{H}$ is bounded in norm by

$$\left(\frac{1}{N} \sum_{j=1}^{N} \|\Gamma_j^* \hat{\Sigma}_j^{-1/2}\|^2\right)^{1/2} \left(\sum_{j=1}^{N} \|\hat{\Sigma}_j^{-1/2} \hat{\Sigma}_j' \hat{H}^{1/2}\|^2\right)^{1/2} \cdot \|N^{1/2} \hat{H}^{1/2}\|$$

The first factor $\frac{1}{N} \sum_{j=1}^{N} \|\Gamma_j^* \hat{\Sigma}_j^{-1/2}\|^2 = O_p(1)$ by the boundedness of $\hat{\Sigma}_j$. The second factor is $\sqrt{T}$ by (A.1). So we have $\sum_{j=1}^{N} \Gamma_j^* \hat{\Sigma}_j^{-1/2} \hat{\Sigma}_j' \hat{H} = \|N^{1/2} \hat{H}^{1/2}\| \cdot O_p(1)$.

Also, the term $\hat{H}^{1/2} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} f_t \epsilon_t'$ is bounded in norm by

$$C \|N^{1/2} \hat{H}^{1/2}\| \cdot \left(\sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}\|^2\right)^{1/2} \left(\frac{1}{N} \sum_{i=1}^{N} \left\| \sum_{t=1}^{T} f_t \epsilon_t' \right\|^2\right)^{1/2}$$

which is $\|N^{1/2} \hat{H}^{1/2}\| \cdot O_p(T^{-1/2})$. Combining results, we have (a).

Consider (b). The left hand side of (b), by the Cauchy-Schwarz inequality, is bounded in norm by

$$C \|N^{1/2} \hat{H}^{1/2}\| \cdot \left(\sum_{j=1}^{N} \|\hat{H}^{1/2} \hat{\Gamma}_j \hat{\Sigma}_j^{-1/2}\|^2\right)^{1/2} \left(\frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \left\| \sum_{t=1}^{T} f_t \epsilon_t' \right\|^2\right)^{1/2}$$

which is $\|N^{1/2} \hat{H}^{1/2}\| \cdot O_p(T^{-1/2})$. So (b) follows.

Consider (c). The left hand side of (c) is bounded by

$$\|\hat{H}^{1/2}\|^2 \sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}\|^2 \cdot \|\hat{\Sigma}_{ii}^{-1/2}\|^2 \cdot \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2$$

$$\leq C \|\hat{H}^{1/2}\|^2 \sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}\|^2 \cdot \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2$$

By $\sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}\|^2 = r$, we have, for all $i$, $\|\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}\| \leq \sqrt{r}$. So the above equation is bounded by $C \sqrt{r} \|\hat{H}^{1/2}\|^2 \sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}\| \cdot \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|$. By the Cauchy-Schwarz inequality, this term is bounded by

$$C \sqrt{r} \frac{1}{\sqrt{N}} \|N^{1/2} \hat{H}^{1/2}\|^2 \left(\sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}\|^2\right)^{1/2} \left(\frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2\right)^{1/2}$$

which is $\|N^{1/2} \hat{H}^{1/2}\| \cdot O_p(N^{-1/2}[1/N \sum_{i=1}^{N} \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2]^{1/2})$. So (c) follows. \(\square\)
Lemma A.6. Let $\mathcal{E} = \|N^{1/2} \hat{H}^{1/2}\|^2 \cdot O_p(\|\hat{\beta} - \beta^*\|)$. Under Assumptions \(A-D\),

\(a\) \( (I - A)^t \left[ \frac{1}{T} \sum_{t=1}^{T} f_t^* (\hat{\beta} - \beta^*) \hat{\xi}_t^t \right] \hat{H}_N = \mathcal{E} \)

\(b\) \( \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta^*) \hat{\xi}_t^t \right] \hat{H}_N = \mathcal{E} \cdot o_p(1) \)

\(c\) \( \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^*) (\hat{\beta} - \beta^*)^t \hat{\xi}_t^t \right] \hat{H}_N = \mathcal{E} \cdot O_p(\|\hat{\beta} - \beta^*\|) \)

where $\hat{\chi}_t, \hat{\xi}_t, \hat{H}_N$ and $A$ are defined in Table 1.

Proof of Lemma A.6. Consider (a). By the definitions of $\hat{\xi}_t$ and $A$, the left hand side of (a) is equivalent to

$$
\hat{H} \left( \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \Gamma_i^{*t} \right) \left( \frac{1}{T} \sum_{t=1}^{T} \sum_{j=1}^{n} f_t^* (\hat{\beta} - \beta^*)^t \hat{\Gamma}_j \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j^t \hat{H} \right)
$$

The term $\hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \Gamma_i^{*t}$ has already been proved to be $\|N^{1/2} \hat{H}^{1/2}\| \cdot O_p(1)$ in Lemma A.5(a). By the definition of $\hat{l}_{jt}$ and the fact $\hat{\chi}_t \hat{\xi}_t = \gamma_{jt}^* f_t^* + \hat{v}_{jt}^p$,

$$
\frac{1}{T} \sum_{j=1}^{K} \sum_{t=1}^{T} f_t^* (\hat{\beta} - \beta^*)^t \hat{\Gamma}_j \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j^t \hat{H} = \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p^*) M_{jt}^* \sum_{j=1}^{N} \gamma_{jp}^* \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j \hat{H}
$$

$$
+ \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p^*) \frac{1}{T} \sum_{j=1}^{K} \sum_{t=1}^{T} f_t^* v_{jp} \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j \hat{H} = d_1 + d_2
$$

where $v_{jp}$ can be replaced by $v_{jp}$ since $\frac{1}{T} \sum_{t=1}^{T} f_t^* = 0$. Term $d_1$ is bounded in norm by

$$
\| M_{jt}^* \| \cdot \sum_{p=1}^{K} |\hat{\beta}_p - \beta_p^*| \cdot \left( \frac{1}{N} \sum_{j=1}^{N} \| \gamma_{jp}^* \hat{\Sigma}_{jj}^{-1/2} \| ^2 \right)^{1/2}
$$

$$
\times \left( \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1/2} \hat{\lambda}_j \hat{H}^{1/2} \|^2 \right)^{1/2} \| N^{1/2} \hat{H}^{1/2} \|
$$

Notice

$$
\sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1/2} \hat{\lambda}_j \hat{H}^{1/2} \|^2 = \sum_{j=1}^{N} \text{tr} \left[ \hat{H}^{1/2} \hat{\Sigma}_{jj}^{-1/2} \hat{\lambda}_j \hat{H}^{1/2} \right] = \text{tr} \left[ \left( \sum_{j=1}^{N} \hat{\lambda}_j \hat{\Sigma}_{jj}^{-1/2} \hat{\lambda}_j \right) \hat{H} \right]
$$
(A.4) \[ \text{tr} \left[ \left( \sum_{j=1}^{N} \lambda_j \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j' \right) \left( \sum_{j=1}^{N} \lambda_j \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j + \sum_{j=1}^{N} \hat{\gamma}_{jx} \hat{\Sigma}_{jx}^{-1} \hat{\gamma}_{jx}' \right)^{-1} \right] < r. \]

But \( \frac{1}{N} \sum_{j=1}^{N} \| \gamma_j^* \hat{\Sigma}_{jj}^{-1/2} \|^2 = O_p(1) \) by the boundedness of \( \hat{\Sigma}_{jj} \), we have \( d_1 = \| N^{1/2} \hat{H}^{1/2} \| \cdot O_p(\beta - \beta) \). Consider \( d_2 \). Ignore \( \sum_{p=1}^{K}(\hat{\beta}_p - \beta_p^*) \), the remaining expression is bounded in norm by

\[
C \left( \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1/2} \hat{\lambda}_j \hat{H}^{1/2} \| \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} \tilde{f}_t^* v_{jt} \right\|^2 \right)^{1/2} \| N^{1/2} \hat{H}^{1/2} \|
\]

which is \( \| N^{1/2} \hat{H}^{1/2} \| \cdot O_p(T^{-1/2}) \). So \( d_2 \) is bounded in norm by

\[
\sum_{p=1}^{K} |\hat{\beta}_p - \beta_p^*| \cdot \| N^{1/2} \hat{H}^{1/2} \| \cdot O_p(\| \beta - \beta^* \|)
\]

Given the results on \( d_1 \) and \( d_2 \), (a) follows.

Consider (b). The left hand side of (b) is equivalent to

\[
\hat{H} \frac{1}{T} \sum_{t=1}^{T} \left( \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} \right) \left( \hat{\beta} - \beta^* \right)' \left( \sum_{j=1}^{N} \hat{\beta}^*_j \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j \right) \hat{H}
\]

by the same arguments in (a), the above expression is equal to

\[
\sum_{p=1}^{K} \left( \hat{\beta}_p - \beta_p^* \right) \hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} \hat{f}_t^* \sum_{j=1}^{N} \gamma_j^* \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j \hat{H}
\]

\[+ \sum_{p=1}^{K} \left( \hat{\beta}_p - \beta_p^* \right) \hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} \hat{v}_{jp} \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j \hat{H} = d_3 + d_4, \text{ say}
\]

where we neglect term

\[
\sum_{p=1}^{K} \left( \hat{\beta}_p - \beta_p^* \right) \hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} \left( \sum_{j=1}^{N} \hat{v}_{jp} \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_j \hat{H} \right)
\]

since it is of smaller order than \( d_4 \). Consider \( d_3 \). Ignore \( \sum_{p=1}^{K}(\hat{\beta}_p - \beta_p^*) \), the remaining expression is bounded in norm by

\[
C \left( \sum_{i=1}^{N} \| \hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2} \|^2 \right)^{1/2} \left( \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1/2} \hat{\lambda}_j \hat{H} \|^2 \right)^{1/2}
\]
\[
\times \left( \frac{1}{N} \sum_{i=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} f^t_i \varepsilon_{it} \right)^2 \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \| \gamma_{jp} S_{\beta_j}^{-1/2} \|_p^2 \right)^{1/2} \| N^{1/2} \hat{H}^{1/2} \|_2^2
\]

which is \( \| N^{1/2} \hat{H}^{1/2} \|_2^2 \cdot O_p(1) \). Given this result, it follows \( d_3 = \| N^{1/2} \hat{H}^{1/2} \|_2^2 \cdot O_p(\| \beta - \beta^* \|) \). Consider \( d_4 \). Ignore \( \sum_{p=1}^{K} (\beta_p - \beta_p^*) \), the remaining expression is equal to

\[
\hat{H} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_{ii} \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} v_{jtp} - E(\varepsilon_{it} v_{jtp})] \hat{\Sigma}_{\beta_j}^{-1} \hat{\lambda}_j \hat{H} + \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_{ii} \hat{\Sigma}_{ii}^{-1} E(\varepsilon_{it} v_{jtp}) \hat{\Sigma}_{\beta_j}^{-1} \hat{\lambda}_j \hat{H}
\]

The first term of the above expression is bounded in norm by

\[
C \left( \sum_{i=1}^{N} \| \hat{H}^{1/2} \hat{\Gamma}_{ii} \hat{\Sigma}_{ii}^{-1/2} \|_2^2 \right)^{1/2} \left( \sum_{j=1}^{N} \| \hat{\Sigma}_{\beta_j}^{-1/2} \hat{\lambda}_j \hat{H} \|_2^2 \right)^{1/2}
\]

\[
\times \left( \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} v_{jtp} - E(\varepsilon_{it} v_{jtp})] \|_2^2 \right) \right)^{1/2} \| N^{1/2} \hat{H}^{1/2} \|_2^2
\]

which is \( \| N^{1/2} \hat{H}^{1/2} \|_2^2 \cdot O_p(T^{-1/2}) \). The second term is bounded in norm by

\[
C \frac{1}{N} \left( \sup_{i \leq N} \| E(\varepsilon_{it} v_{jtp}) \| \right) \left( \sum_{i=1}^{N} \| \hat{H}^{1/2} \hat{\Gamma}_{ii} \hat{\Sigma}_{ii}^{-1/2} \|_2^2 \right)^{1/2}
\]

\[
\times \left( \sum_{j=1}^{N} \| \hat{\Sigma}_{\beta_j}^{-1/2} \hat{\lambda}_j \hat{H} \|_2^2 \right)^{1/2} \| N^{1/2} \hat{H}^{1/2} \|_2^2
\]

Notice \( \sup_{i \leq N} \| E(\varepsilon_{it} v_{jtp}) \| \leq [\sup_{i \leq N} (\| \varepsilon_{it} \|_2^2)^{1/2}] [\sup_{i \leq N} (\| E v_{jtp} \|_2^2)^{1/2}] \), which is bounded by Assumption B. So the above expression is \( \| N^{1/2} \hat{H}^{1/2} \|_2^2 \cdot O_p(N^{-1}) \). Thus,

\[
\hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_{ii} \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} v_{jtp} \hat{\Sigma}_{\beta_j}^{-1} \hat{\lambda}_j \hat{H} = \| N^{1/2} \hat{H}^{1/2} \|_2^2 \cdot [O_p(N^{-1}) + O_p(T^{-1/2})]
\]

Given this result, it follows \( d_4 = \| N^{1/2} \hat{H}^{1/2} \|_2^2 \cdot O_p(\| \beta - \beta^* \|) \). Combining the results on \( d_3 \) and \( d_4 \), we have (b).

Consider (c). The left hand side of (c) is equivalent to

\[
\hat{H} \frac{1}{T} \sum_{t=1}^{T} \left( \sum_{i=1}^{N} \hat{\Gamma}_{ii} \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} \right) (\hat{\beta} - \beta^*)(\hat{\beta} - \beta^*)' \left( \sum_{j=1}^{N} \hat{\Sigma}_{\beta_j}^{-1} \hat{\lambda}_j \right) \hat{H}
\]
which is equal to
\[
\sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*) (\beta_q - \beta_q^*) \left\{ \hat{H} \sum_{i=1}^{N} \lambda_i \hat{\Sigma}_{ie}^{-1} \gamma_{ip}^* M_{ij}^* \sum_{j=1}^{N} \gamma_{jq}^* \hat{\Sigma}_{je}^{-1} \lambda_j \hat{H} \right. \\
+ \hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} \lambda_i \hat{\Sigma}_{ie}^{-1} v_{itp} f_{it}^* \sum_{j=1}^{N} \gamma_{jq}^* \hat{\Sigma}_{je}^{-1} \lambda_j \hat{H} \\
+ \hat{H} \sum_{i=1}^{N} \lambda_i \hat{\Sigma}_{ie}^{-1} \frac{1}{T} \sum_{j=1}^{N} \sum_{t=1}^{T} f_{it}^* v_{jtq} \hat{\Sigma}_{je}^{-1} \lambda_j \hat{H} \\
+ \hat{H} \sum_{i=1}^{N} \lambda_i \hat{\Sigma}_{ie}^{-1} E(v_{itp} v_{itq}) \hat{\Sigma}_{ie}^{-1} \lambda_i \hat{H} \right\}
\]
\[
= \sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*) (\beta_q - \beta_q^*) (d_5 + d_6 + \cdots + d_9), \quad \text{say}
\]
where we neglect term
\[
\sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*) (\beta_q - \beta_q^*) \left[ \hat{H} \sum_{i=1}^{N} \sum_{j=1}^{N} \lambda_i \hat{\Sigma}_{ie}^{-1} v_{itp} v_{jtq} \hat{\Sigma}_{je}^{-1} \lambda_j \hat{H} \right]
\]
since it is of smaller order than \( \sum_{p=1}^{K} \sum_{q=1}^{K} (\beta_p - \beta_p^*) (\beta_q - \beta_q^*) d_8 \).

Consider \( d_5 \), which is bounded in norm by
\[
\| M_{ff}^* \| \cdot \| N^{1/2} \hat{H}^{1/2} \| ^2 \left( \frac{1}{N} \sum_{i=1}^{N} \| \gamma_{ip}^* \hat{\Sigma}_{ie}^{-1/2} \| ^2 \right) ^{1/2} \\
\times \left( \frac{1}{N} \sum_{j=1}^{N} \| \gamma_{jq}^* \hat{\Sigma}_{je}^{-1/2} \| ^2 \right) ^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{H}^{1/2} \lambda_i \hat{\Sigma}_{ie}^{-1/2} \| ^2 \right) = \| N^{1/2} \hat{H}^{1/2} \| ^2 \cdot O_p(1).
\]

Consider \( d_6 \), which is bounded in norm by
\[
C \| N^{1/2} \hat{H}^{1/2} \| ^2 \left( \frac{1}{N} \sum_{i=1}^{N} \| \gamma_{ip}^* \hat{\Sigma}_{ie}^{-1/2} \| ^2 \right) ^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \| \hat{H}^{1/2} \lambda_i \hat{\Sigma}_{ie}^{-1/2} \| ^2 \right) ^{1/2} \\
\times \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{H}^{1/2} \lambda_i \hat{\Sigma}_{ie}^{-1/2} \| ^2 \right) = \| N^{1/2} \hat{H}^{1/2} \| ^2 \cdot O_p(T^{-1/2}).
\]
The term $d_7$ is also $\|N^{1/2}\hat{H}^{1/2}\|^2 \cdot O_p(T^{-1/2})$, similar to $d_6$. Term $d_8$ is bounded in norm by

$$C\|N^{1/2}\hat{H}^{1/2}\|^2\left(\sum_{i=1}^{N} \|\hat{H} \hat{\lambda} \hat{\Sigma}_{ite}^{-1/2}\|^2\right)$$

$$\times \left(\frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left[\frac{1}{T} \sum_{t=1}^{T} v_{ipt} v_{jtq} - E(v_{ipt} v_{jtq})\right]^2\right)^{1/2},$$

which is $\|N^{1/2}\hat{H}^{1/2}\|^2 \cdot O_p(T^{-1/2})$.

Consider $d_9$, which is bounded in norm by

$$C \frac{1}{N} \|N^{1/2}\hat{H}^{1/2}\|^2 \left(\sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\lambda} \hat{\Sigma}_{ite}^{-1/2}\|^2\right) \cdot \sup_{i \leq N} |E(v_{ipt} v_{jtq})|$$

which is $\|N^{1/2}\hat{H}^{1/2}\|^2 \cdot O_p(N^{-1})$.

Summarizing all the results, we have (c).

\[ \Box \]

**Proof of Proposition 2.1.** We consider the objective function,

$$\ln L = -\frac{1}{N} \ln |\Sigma_\alpha| - \frac{1}{N} \text{tr}\left[(I_N \otimes B) M_\alpha (I_N \otimes B') \Sigma_\alpha^{-1}\right] + K + \frac{1}{N} \ln |\Sigma_\alpha^*|$$

Here we add a centering constant to (2.4). Throughout the paper, we assume $\hat{\beta}$ is in a compact set. Since the likelihood function is a quadratic form of $\hat{\beta}$, it cannot achieve the maximum value at too large $\hat{\beta}$. So this assumption doesn’t loss generality.

Let $\Gamma^\dagger = (I_N \otimes B)^{-1}\Gamma$ and $\Sigma^\dagger_{\epsilon \epsilon} = (I_N \otimes B)^{-1}\Sigma_{\epsilon \epsilon}(I_N \otimes B')^{-1}$, then $\Sigma^\dagger_\alpha = \Gamma^\dagger M_{\alpha} \Gamma^\dagger + \Sigma^\dagger_{\epsilon \epsilon} = (I_N \otimes B)^{-1}\Sigma_{\epsilon \epsilon}(I_N \otimes B')^{-1}$. Notice $\ln |I_N \otimes B| = 0$ by the definition of $B$. So the likelihood function can be written as

(A.5) \[ \ln L = -\frac{1}{N} \ln |\Sigma_\alpha^*| - \frac{1}{N} \text{tr}[M_\alpha \Sigma_\alpha^{-1}] + K + \frac{1}{N} \ln |\Sigma_\alpha^*| \]

By $(I_N \otimes B^*) \dot{z}_t = \Gamma^* f_t^* + \dot{\epsilon}_t$, we have

$$(I_N \otimes B^*) M_\alpha (I_N \otimes B^*) = \Gamma^* M_{\alpha} \Gamma^* + \Sigma_{\epsilon \epsilon} + \Gamma^* \frac{1}{T} \sum_{t=1}^{T} f_t^* \dot{\epsilon}_t$$

$$+ \frac{1}{T} \sum_{t=1}^{T} \dot{z}_t f_t^* \Gamma^* + \frac{1}{T} \sum_{t=1}^{T} [\dot{z}_t \dot{\epsilon}_t - \Sigma_{\epsilon \epsilon}]$$
Let $\Gamma^d = (I_N \otimes B^*)^{-1} \Gamma^*$, $\Sigma^d_{ee} = (I_N \otimes B^*)^{-1} \Sigma_{ee}^* (I_N \otimes B^*)^{-1}$ and $\Sigma^d_{z\hat{z}} = \Gamma^d M^d_{zf} \Gamma^{d\prime} + \Sigma^d_{z\hat{z}}$. Then the proceeding equation can be rewritten as

$$M_z = \Sigma^d_{z\hat{z}} + \Gamma^d \frac{1}{T} \sum_{t=1}^T \hat{f}_t^* \hat{f}'_t (I_N \otimes B^{s_t-1}) + (I_N \otimes B^{s_t-1}) \frac{1}{T} \sum_{t=1}^T \hat{f}_t^* \hat{f}'_t \Gamma^{s_t\prime}$$

$$+(I_N \otimes B^{s_t-1}) \frac{1}{T} \sum_{t=1}^T |\hat{e}_t \hat{e}'_t - \Sigma_{ee}^*|(I_N \otimes B^{s_t-1})$$

Equation (A.5) can be rewritten as

(A.6) \n \ln L = \mathcal{L}(\theta) + R(\theta)

where

$$\mathcal{L}(\theta) = -\frac{1}{N} \ln |\Sigma^d_{z\hat{z}}| - \frac{1}{N} \text{tr} \left[ \Sigma^d_{z\hat{z}} (\Sigma^d_{z\hat{z}})^{-1} \right] + K + \frac{1}{N} \ln |\Sigma^*_{z\hat{z}}|$$

and

$$R(\theta) = -\frac{1}{N} \text{tr} \left[ (M_z - \Sigma^d_{z\hat{z}}) (\Sigma^d_{z\hat{z}})^{-1} \right]$$

Lemma A.3 implies that $\sup_{\theta \in \Theta} |R(\theta)| = o_p(1)$. So we have $|R(\theta^*) - R(\theta)| \leq 2 \sup_{\theta \in \Theta} |R(\theta)| = o_p(1)$. Since $\theta$ maximizes $\ln L$, it follows that $\mathcal{L}(\theta) + R(\theta) = \mathcal{T}(\theta^*) + R(\theta)$. This yields $\mathcal{T}(\theta) = \mathcal{T}(\theta^*) + R(\theta) - R(\theta^*) \geq \mathcal{L}(\theta^*) - |o_p(1)| = -|o_p(1)|$, the last equation uses the fact that $\mathcal{L}(\theta^*) = 0$ by the centering. However, $\mathcal{L}(\theta)$ is maximized at $\theta = \theta^*$ and $\mathcal{L}(\theta^*) = 0$. This yields $\mathcal{T}(\theta) \leq 0$. Combining $\mathcal{T}(\theta) \geq -|o_p(1)|$ and $\mathcal{T}(\theta) \leq 0$, we have $\mathcal{T}(\theta) = o_p(1)$.

Consider $\mathcal{L}(\hat{\theta})$, which is equivalent to

$$\mathcal{L}(\hat{\theta}) = -\frac{1}{N} \ln \left| \hat{\Gamma}^\dagger M^d_{zf} \hat{\Gamma}^{d\prime} + \hat{\Sigma}^d_{z\hat{z}} \right| - \frac{1}{N} \text{tr} \left[ \Gamma^d M^d_{zf} \Gamma^{d\prime \dagger} \hat{\Sigma}^d_{z\hat{z}} \right]$$

Notice that $\hat{\Sigma}^d_{z\hat{z}} = \Sigma^d_{z\hat{z}} - \Sigma^d_{z\hat{z}} \hat{\Gamma}^\dagger \hat{\Gamma}^{d\prime} \hat{\Sigma}^d_{z\hat{z}}$ with $\hat{\Gamma}^\dagger = (M^d_{zf} + \hat{\Gamma}^{d\prime} \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger})^{-1}$. So $\frac{1}{N} \text{tr}[\Sigma^d_{z\hat{z}} (\Sigma^d_{z\hat{z}})^{-1}]$ is equal to $\frac{1}{N} \text{tr}[\Sigma^d_{z\hat{z}} (\Sigma^d_{z\hat{z}})^{-1}] - \frac{1}{N} \text{tr}[\hat{\Gamma}^\dagger \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger} \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger} \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger} \hat{\Gamma}^{d\prime \dagger}]$. The term $\frac{1}{N} \text{tr} \left[ \Sigma^d_{z\hat{z}} (\Sigma^d_{z\hat{z}})^{-1} \right]$ is equal to $\frac{1}{N} \sum_{i=1}^N (\hat{\Sigma}^d_{z\hat{z}})^{-1} (\hat{\Sigma}^d_{z\hat{z}})^{-1} (\hat{\Sigma}^d_{z\hat{z}})^{-1} (\hat{\Sigma}^d_{z\hat{z}})$. The term $\frac{1}{N} \text{tr}[\hat{\Gamma}^\dagger \hat{\Gamma}^{d\prime \dagger} \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger} \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger} \hat{\Gamma}^{d\prime \dagger}]$ is equal to $\left[ \frac{1}{N} \sum_{i=1}^N (\hat{\Sigma}^d_{z\hat{z}})^{-1} (\hat{\Sigma}^d_{z\hat{z}})^{-1} (\hat{\Sigma}^d_{z\hat{z}})^{-1} (\hat{\Gamma}^{d\prime \dagger} \hat{\Gamma}^{d\prime \dagger}) \right]$, where $\hat{\Sigma}^d_{z\hat{z}} = B^* \Sigma^d_{z\hat{z}} B^* - 1$, $\Sigma^d_{z\hat{z}} = B^* \Sigma^d_{z\hat{z}} B^* - 1$ and $\hat{\Gamma} = B^* \hat{\Gamma}$. The latter term is bounded in norm by (ignore the trace)

$$\frac{1}{N} \sup_{i \leq N} \left( \Sigma^d_{z\hat{z}} \right) \cdot \left( \frac{1}{N} \sum_{i=1}^N \left( \hat{\Gamma}^\dagger \hat{\Gamma}^{d\prime \dagger} \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger} \hat{\Sigma}^d_{z\hat{z}} \hat{\Gamma}^{d\prime \dagger} \hat{\Gamma}^{d\prime \dagger} \right) \right) \cdot \left( 1 + \hat{\Gamma}^\dagger \hat{\Gamma}^{d\prime \dagger} \hat{\Gamma}^{d\prime \dagger} \hat{\Gamma}^{d\prime \dagger} \right)$$
Using the arguments in proving their Proposition 1 of [10], we can show
\[ (A.9) \]
which are arranged in descending order. Similarly, \( (A.10) \)
the second factor is \( r \). So we have \( \frac{1}{N} \text{tr}[\hat{\Sigma}_e^{-1} \Sigma_{ee} \hat{\Sigma}_e^{-1} \Sigma_{ee} - I_N] = O_p(N^{-1}). \)
Using the arguments in proving their Proposition 1 of [10], we can show
\[ (A.8) \]
\[ (A.7) \]
The main three terms of the above equation are all non-positive. By \( \mathcal{L}(\hat{\theta}) = o_p(1) \), each of the three terms must be \( o_p(1) \). That is,
\[ (A.7) \]
\[ (A.8) \]
\[ (A.9) \]
Consider (A.7). The matrix \( \Sigma_{iix} \) is definite positive matrix for all \( i \) and \( \hat{\Sigma}_{iie} \) is a scaler which is bounded by \([C^{-1}, C]\). So we have
\[ (A.10) \]
Consider (A.8). Let \( \hat{\omega}_i, \hat{\omega}_2, \cdots, \hat{\omega}_K \) be the eigenvalues of the matrix \( \hat{\Sigma}_{ii} \) which are arranged in descending order. Similarly, \( \omega_{i1}, \omega_{i2}, \cdots, \omega_{iK} \) are the eigenvalues of the matrix \( \Sigma_{ii} \) in descending order. Consider the following function
\[ f(\hat{\Sigma}_{ii}) = \ln |\hat{\Sigma}_{ii}| + \text{tr}[\hat{\Sigma}_{ii}^{-1} \Sigma_{ii}^*] - K - \ln |\Sigma_{ii}^*| - b\| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 \]
all of the above equation can be written as

$$2f(\hat{\Sigma}_{ii}) = \sum_{p=1}^{K} \left[ (\ln \hat{\omega}_{ip} - 1 - \ln \omega_{ip}^*) - b(\hat{\omega}_{ip}^2 + \omega_{ip}^*) + tr[(\hat{\Sigma}_{ii}^{-1} + 2b\hat{\Sigma}_{ii})\Sigma_{ii}^*] \right]$$

The matrix $\hat{\Sigma}_{ii}^{-1} + 2b\hat{\Sigma}_{ii}$ has the eigenvalues $\hat{\omega}_{ip}^{-1} + 2b\hat{\omega}_{ip}$ for $p = 1, 2, \cdots, K$. We can choose $b$ small enough to guarantee the order of $\hat{\omega}_{i1}^{-1} + 2b\hat{\omega}_{i1}, \hat{\omega}_{i2}^{-1} + 2b\hat{\omega}_{i2}, \cdots, \hat{\omega}_{iK}^{-1} + 2b\hat{\omega}_{iK}$ is the same as the order of $\hat{\omega}_{i1}^{-1}, \hat{\omega}_{i2}^{-1}, \cdots, \hat{\omega}_{iK}^{-1}$ because all $\hat{\omega}_{ip}$ are bounded by Assumption D. So we have

$$f(\hat{\Sigma}_{ii}) \geq \sum_{p=1}^{K} \left\{ (\ln \hat{\omega}_{ip} - 1 - \ln \omega_{ip}^*) - b(\hat{\omega}_{ip}^2 + \omega_{ip}^*) + \hat{\omega}_{ip}^{-1} \omega_{ip}^* + 2b\hat{\omega}_{ip}\omega_{ip}^* \right\}$$

by Lemma A.1. Using the arguments of [10], there exists a constant $b$ small enough such that $f(\hat{\Sigma}_{ii}) \geq 0$. So we have

$$o_p(1) = \frac{1}{N} \sum_{i=1}^{N} (\ln |\hat{\Sigma}_{ii}| + tr[\hat{\Sigma}_{ii}^{-1}\Sigma_{ii}^*] - K - \ln |\Sigma_{ii}^*|) \geq \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii} - \Sigma_{ii}^*||^2$$

This implies that

(A.11) \[ \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii} - \Sigma_{ii}^*||^2 \xrightarrow{P} 0 \]

Consider (A.9), which is equal to $\frac{1}{N} tr\left( [I_N \otimes (\hat{B}B^{-1})] \Gamma^s M_{jj}^s \Gamma^s [I_N \otimes (\hat{B}B^{-1})]^{\dagger} \hat{\Sigma}_{zz}^{-1} \right) = o_p(1)$. By (A.10) and Lemma A.4 , we have

$$\frac{1}{N} tr\left[ \Gamma^s M_{jj}^s \Gamma^s \hat{\Sigma}_{zz}^{-1} \right] = o_p(1)$$

Since $\hat{\Sigma}_{zz}^{-1} = \hat{\Sigma}_{ee}^{-1} - \hat{\Sigma}_{ee}^{-1} \hat{G} \hat{G}^\dagger \hat{\Sigma}_{ee}^{-1}$ and $\hat{G} = \hat{H} - \hat{H} \hat{M}_{jj}^{-1} \hat{G}$, the left hand side of the above equation can be written as

$$tr\left[ \frac{1}{N} \Gamma^s \hat{\Sigma}_{ee}^{-1} \Gamma^s M_{jj}^s \right] - \frac{1}{N} tr\left( \Gamma^s \hat{\Sigma}_{ee}^{-1} \Gamma^s \hat{G} \hat{G}^\dagger \hat{\Sigma}_{ee}^{-1} \Gamma^s M_{jj}^s \right)$$

$$+ tr\left[ \frac{1}{N} \Gamma^s \hat{\Sigma}_{ee}^{-1} \Gamma^s \hat{G} \hat{G}^\dagger \hat{\Sigma}_{ee}^{-1} \Gamma^s M_{jj}^s \right]$$

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The matrix inside each of the trace operator is semi-definitely positive. Given $\frac{1}{N} \text{tr} \left[ \Gamma^* M_{ff} \Gamma^* \Sigma^{-1}_{ee} \right] = o_p(1)$ together with the fact that $M = o_p(1)$ if $\text{tr}(M) = o_p(1)$ for a semi-positive definite matrix $M$, we have

\begin{equation}
\frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma^* - \frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \hat{\Gamma} (\hat{\Gamma}^* \Sigma^{-1}_{ee} \hat{\Gamma})^{-1} \hat{\Gamma}^* \Sigma^{-1}_{ee} \Gamma^* = o_p(1)
\end{equation}

and

\begin{equation}
\frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma (\hat{\Gamma}^* \Sigma^{-1}_{ee} \hat{\Gamma})^{-1} \hat{M}_{ff} \hat{\Gamma}^* \Sigma^{-1}_{ee} \Gamma^* = o_p(1)
\end{equation}

Notice that $\frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma^* = \frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma^* + \frac{1}{N} \Gamma^* (\hat{\Sigma}^{-1}_{ee} - \Sigma^{-1}_{ee}) \Gamma^*$. Since the term $\frac{1}{N} \Gamma^* (\hat{\Sigma}^{-1}_{ee} - \Sigma^{-1}_{ee}) \Gamma^*$ is equal to $\frac{1}{N} \sum_{i=1}^{N} \Gamma_i^* (\hat{\Sigma}^{-1}_{ii} - \Sigma^{-1}_{ii}) \Sigma^{-1}_{ii} \Gamma_i^*$, which is bounded in norm by $\frac{1}{N} \sum_{i=1}^{N} ||\Gamma_i^* \Sigma^{-1}_{ii}|| \cdot ||\Sigma^{-1}_{ii} - \hat{\Sigma}^{-1}_{ii}|| \cdot ||\hat{\Sigma}^{-1}_{ii} - \Sigma^{-1}_{ii}||$ and is further bounded by $C \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}^{-1}_{ii} - \Sigma^{-1}_{ii}||$ for a large $C$ by the boundedness of $||\Gamma_i^*||$, $||\hat{\Sigma}^{-1}_{ii}||$, $||\Sigma^{-1}_{ii}||$ for all $i$ in view of Assumptions C and D. By the Cauchy-Schwarz inequality, $C \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}^{-1}_{ii} - \Sigma^{-1}_{ii}||$ is bounded by $C (\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}^{-1}_{ii} - \Sigma^{-1}_{ii}||^2)^{1/2}$. So we have

\begin{equation}
\left\| \frac{1}{N} \Gamma^* (\hat{\Sigma}^{-1}_{ee} - \Sigma^{-1}_{ee}) \Gamma^* \right\| = \left\| \frac{1}{N} \sum_{i=1}^{N} \Gamma_i^* (\hat{\Sigma}^{-1}_{ii} - \Sigma^{-1}_{ii}) \Sigma^{-1}_{ii} \Gamma_i^* \right\|
\leq C \left( \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}^{-1}_{ii} - \Sigma^{-1}_{ii}||^2 \right)^{1/2}
\end{equation}

Then we have $\frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma^* = o_p(1)$ by (A.11). So $\frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma^* = \frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma^* + o_p(1)$. Given this result, together with (A.12), it follows

\begin{equation}
\frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma (\hat{\Gamma}^* \Sigma^{-1}_{ee} \hat{\Gamma})^{-1} \hat{\Gamma}^* \Sigma^{-1}_{ee} \Gamma^* \overset{p}{\rightarrow} C^*
\end{equation}

where $C^*$ is the limit of $\frac{1}{N} \Gamma^* \Sigma^{-1}_{ee} \Gamma^*$. Comparing the above result with (A.13), we have $\hat{M}_{ff} \hat{G} = o_p(1)$. So $\hat{G} = o_p(1)$ due to the boundedness of $\hat{M}_{ff}$ by Assumption D. Since $\hat{G} = \hat{H} (I_r - \hat{M}_{ff}^{-1} \hat{G})$, we also have $\hat{H} = o_p(1)$. We summarize these results as

\begin{equation}
\hat{G} = o_p(1); \quad \hat{H} = o_p(1)
\end{equation}

Let $A = (\hat{\Gamma} - \Gamma^*) \Sigma_{ee}^{-1} \hat{\Gamma} \hat{H}$, then $\Gamma^* \Sigma_{ee}^{-1} \Gamma \hat{H} = I_r - A$. By (A.14) and (A.11), equation (A.12) can be written, in terms of $A$, as

\begin{equation}
\frac{1}{N} \Gamma^* \Sigma_{ee}^{-1} \Gamma^* - (I_r - A) \left( \frac{1}{N} \hat{\Gamma}^* \Sigma_{ee}^{-1} \hat{\Gamma} \right) (I_r - A)' = o_p(1)
\end{equation}
Using (A.20), equation (2.7) is equal to

$$\frac{1}{N} \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i') \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i')' - A \left( \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{\epsilon \epsilon}^{-1} \hat{\Gamma} \right) A' = o_p(1)$$

Now we turn to the first order condition for the proof of consistency. Notice that

$$\hat{\Gamma} = \Gamma^* + l' \hat{\beta} - (\hat{\beta}^*)'$$

where \( l^* = (\hat{x}'_j, 0_{K \times K})' \) is a \((K + 1) \times K\) matrix. Also notice that

$$\hat{\Sigma}_{zz} = \Gamma^* M_{ff}^* \Gamma_j^* + \frac{1}{T} \sum_{t=1}^{T} \epsilon_{it}' M_{ff}^* \epsilon_j + \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_{it}' M_{ff}^* \epsilon_j$$

$$- \frac{1}{T} \sum_{t=1}^{T} l' \hat{\beta}^* - \frac{1}{T} \sum_{t=1}^{T} \epsilon_{it}' (\hat{\beta}^*)' l_j + \frac{1}{T} \sum_{t=1}^{T} l' \hat{\beta}^* - \frac{1}{T} \sum_{t=1}^{T} \epsilon_{it}' (\hat{\beta}^*)' l_j$$

where \( 1(i = j) = 1 \) if \( i = j \), 0 otherwise, and \( \epsilon_{ij} = \epsilon_{it} \epsilon_j - E(\epsilon_{it} \epsilon_j) \). For simplicity, we neglect the smaller order term \( \bar{\epsilon}_i \epsilon_j = T^{-2} (\sum_{t=1}^{T} \epsilon_{it}) (\sum_{j=1}^{T} \epsilon_{jt}) \). Using (A.20), equation (2.7) is equal to

$$M_{ff} - M_{ff}' = -A'M_{ff}' - M_{ff} A + A'M_{ff}' A + (I - A)' \left[ \frac{1}{T} \sum_{t=1}^{T} f_t' \hat{\epsilon}_t \right] H_N$$

$$+ \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\epsilon}_t f_t' \right] (I - A) + \hat{H} \left[ \frac{1}{T} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\epsilon}_{it} \hat{\Sigma}_{ii}^{-1} \hat{\epsilon}_{jt}^* \hat{\Gamma}_j^* \right] \hat{H}$$

$$- (I - A)' \left[ \frac{1}{T} \sum_{t=1}^{T} f_t' (\hat{\beta}^*)' \hat{\epsilon}_t^* \right] \hat{H}_N - \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta}^*)' \hat{\epsilon}_t^* \right] \hat{H}_N$$
where \( \hat{\chi}_t, \hat{\xi}_t, \hat{N}, \hat{H} \) and \( A \) are defined in Table 1 and \( \epsilon_{i,j,t} = \epsilon_{i,j,t} - E(\epsilon_{i,j,t}) \).

Consider (A.21). The 4th-6th and the last terms of the right hand side of (A.21) are summarized in Lemma A.5. The 7th-11th terms are summarized in Lemma A.6. Since we have already proved that \( \hat{\beta} - \beta^* = o_p(1) \) and

\[
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 = o_p(1),
\]

(A.21) can be written as

\[
\hat{M}_{ff} = (I_r - A)' \hat{M}_{ff}^* (I_r - A) + \| N^{1/2} \hat{H}^{1/2} \|^2 \cdot o_p(1)
\]

However, (A.16) indicates that \( N \hat{H} = (I_r - A)' \left( \frac{1}{N} \Gamma^* \hat{\Sigma}_{ee}^{-1} \Gamma^* \right)^{-1} (I_r - A) + o_p(\| I_r - A \|^2) \). So \( \| N^{1/2} \hat{H}^{1/2} \|^2 = \text{tr}(N \hat{H}) = \text{tr}((I_r - A)' \left( \frac{1}{N} \Gamma^* \hat{\Sigma}_{ee}^{-1} \Gamma^* \right)^{-1} (I_r - A)) + o_p(\| I_r - A \|^2) \). Given this result, we can show \( A = O_p(1) \). Otherwise, if \( A \) is stochastically unbounded, the first term of the right hand side of (A.22), which dominates the second term, will diverge to infinity. But the left hand side of (A.22) is stochastically bounded by Assumption D (it is \( I_r \) under our identification condition). A contradiction is obtained. Given \( A = O_p(1) \), we immediately obtain

\[
N \hat{H} = O_p(1)
\]

Then (A.22) can be simplified as \( \hat{M}_{ff} = (I_r - A)' \hat{M}_{ff}^* (I_r - A) + o_p(1) \). Notice that the identification condition requires \( M_{ff} = \hat{M}_{ff} = I_r \). This yields

\[
(I_r - A)' (I_r - A) = I_r + o_p(1)
\]

By (A.16) and (A.24), applying Lemma A.2 with \( Q = (I_r - A)' \), \( V = \frac{1}{N} \Gamma^* \hat{\Sigma}_{ee}^{-1} \Gamma^* \) and \( D = \frac{1}{N} \Gamma^* \hat{\Sigma}_{ee}^{-1} \Gamma^* \), it follows that \( I_r - A \) converges in probability to a diagonal matrix with diagonal elements either 1 or \(-1\). By assuming \( \hat{\Gamma} \) and \( \Gamma^* \) have the same column signs, we rule out \(-1\) as the diagonal element. So \( A = o_p(1) \). Then (A.17) implies the second result of Proposition 2.1.

This completes the proof of Proposition 2.1.

**Corollary A.1.** Under the assumptions of Proposition 2.1, we have

\[
\begin{align*}
(a) \quad & \frac{1}{N} \Gamma^* \hat{\Sigma}_{ee}^{-1} \Gamma^* = o_p(1) \\
(b) \quad & \frac{1}{N} \Gamma^* \hat{\Sigma}_{ee}^{-1} \Gamma^* = o_p(1) \\
(c) \quad & A = (\hat{\Gamma} - \Gamma^*) \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \hat{H} = o_p(1)
\end{align*}
\]
APPENDIX A2: PROOFS OF THEOREMS 2.1 AND 2.2

Given consistency, we now drop the superscript “*” from the true parameters for notational simplicity. The following lemmas are useful to prove Theorem 2.1.

**Lemma A.7.** Under Assumptions A-D,

(a) \( \frac{1}{N} \sum_{j=1}^{N} \left\| (I - A) \frac{1}{T} \sum_{t=1}^{T} f_t \varepsilon_{jt} \right\|^2 = O_p(T^{-1}) \)

(b) \( \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_t \right) \Gamma_j \right\|^2 = O_p(T^{-1}) \)

(c) \( \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{jt} f_t \right\|^2 = O_p(T^{-1}) \)

(d) \( \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H} \hat{\Gamma}_j \hat{\Sigma}_{jj}^{-1} (\hat{\Sigma}_{jj} - \Sigma_{jj}) \right\|^2 = O_p(\frac{1}{N} \sum_{j=1}^{N} \left\| \hat{\Sigma}_{jj} - \Sigma_{jj} \right\|^2) \)

where \( \varepsilon_{jt} = \varepsilon_{it} \varepsilon_{jt}' - E(\varepsilon_{it} \varepsilon_{jt}') \).

**Proof of Lemma A.7.** Consider (a). The left hand side of (a) is bounded by

\[ \left\| I - A \right\|^2 \frac{1}{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} f_t \varepsilon_{jt} \right\|^2 \]

which is \( O_p(T^{-1}) \) because \( \frac{1}{T} \sum_{t=1}^{T} f_t \varepsilon_{jt}' = O_p(T^{-1/2}) \) and \( A = o_p(1) \) by Corollary A.1(c).

Consider (b). The left hand side of (b) is equal to

\[ \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{jt} f_t \Gamma_j \right\|^2 \]

which is bounded by

\[ C \cdot ||\hat{H}||^2 \cdot \left( \frac{1}{N} \sum_{j=1}^{N} \left\| \Gamma_j \right\|^2 \right) \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \right\|^2 \right) \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} f_t \varepsilon_{it}' \right\|^2 \right) \]

which is \( O_p(T^{-1}) \) by Corollary A.1(b) and \( \hat{H} = O_p(1) \).

Consider (c). The left hand side of (c) is bounded by

\[ C \cdot ||\hat{H}||^2 \cdot \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \right\|^2 \right) \left( \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{jt} \right\|^2 \right) \]
which is \(O_p(T^{-1})\) by the same reason as (b) and Assumption B.

Consider (d). The left hand side of (d) is bounded by

\[
C \frac{1}{N} \|N \hat{H}\|^2 \cdot \left( \frac{1}{N} \sum_{j=1}^{N} \|\hat{\Gamma}_j \hat{\Sigma}_j^{-1/2}\|^2 \right) \left( \frac{1}{N} \sum_{j=1}^{N} \|\hat{\Sigma}_j - \Sigma_j\|^2 \right)
\]

which is \(O_p(N^{-2} \sum_{j=1}^{N} \|\hat{\Sigma}_j - \Sigma_j\|^2)\) by the same reason as (b). \(\square\)

**Lemma A.8.** Under Assumptions A-D,

\[\begin{align*}
(a) & \quad \frac{1}{N} \sum_{j=1}^{N} \left\| (I - A) \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{l}_{jt} \right\|^2 = O_p(\|\hat{\beta} - \beta\|^2) \\
(b) & \quad \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{l}_{jt} \right) \right\|^2 = o_p(\|\hat{\beta} - \beta\|^2) \\
(c) & \quad \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) f_t' \right) \hat{\epsilon}_{jt} \right\|^2 = O_p(\|\hat{\beta} - \beta\|^2) \\
(d) & \quad \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) (\hat{\beta} - \beta)' \right) \hat{l}_{jt} \right\|^2 = o_p(\|\hat{\beta} - \beta\|^2) \\
(e) & \quad \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) (\hat{\beta} - \beta)' \right) \hat{l}_{jt} \right\|^2 = O_p(\|\hat{\beta} - \beta\|^4)
\]

**Proof of Lemma A.8.** Consider (a). The left hand side of (a) is bounded in norm by

\[
\|I - A\|^2 \cdot \left( \frac{1}{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{l}_{jt} \right\|^2 \right)
\]

The first factor is \(O_p(1)\) by Corollary A.1(c). Consider the second factor, which is equal to

\[
\frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) M_{jp} f_j + \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \frac{1}{T} \sum_{t=1}^{T} f_{vjp} \right\|^2
\]

The above expression is bounded by

\[
2 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) M_{jp} f_j \right\|^2 + 2 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \frac{1}{T} \sum_{t=1}^{T} f_{vjp} \right\|^2
\]
Notice

\[(A.25) \quad |\hat{\beta}_p - \beta_p| \leq \|\hat{\beta} - \beta\| \quad \text{for } \forall p\]

The first term is \(O_p(\|\hat{\beta} - \beta\|^2)\), the second term is \(O_p(T^{-1}\|\hat{\beta} - \beta\|^2)\). Thus (a) follows.

Consider (b). The left hand side of (b) is equivalent to

\[
\frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} (\hat{\beta} - \beta)^j \frac{\partial f_j}{\partial \beta} \right\|^2
\]

which is equal to

\[
\frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} \hat{x}_{jtp} \right\|^2
\]

By \(\hat{x}_{jtp} = \gamma_{jtp} f_t + \hat{v}_{jtp}\), the above expression is bounded by

\[
4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} f'_j \gamma_{jp} \right\|^2
\]

\[+ 4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} v_{jtp} - E(\varepsilon_{it} v_{jtp}) \right\|^2
\]

\[+ 4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} E(\varepsilon_{it} v_{jtp}) \right\|^2
\]

\[+ 4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} \hat{v}_{jtp} \right\|^2 = e_1 + e_2 + e_3 + e_4, \quad \text{say}
\]

Consider \(e_1\). Ignore the factor 4, it is bounded, due to (A.25), by

\[
CK\|\hat{\beta} - \beta\|^2 \|\hat{\Sigma}^{1/2} \hat{H}^{1/2}\|^2 \left( \sum_{i=1}^{N} \|\hat{\Sigma}^{1/2} \hat{\Gamma}_i \hat{\Sigma}^{-1/2}_{ii}\|^2 \right)
\]

\[
\times \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} f'_i \right\|^2 \right) \sum_{p=1}^{K} \frac{1}{N} \sum_{j=1}^{N} \|\gamma_{jp}\|^2
\]

So \(e_1 = O_p(T^{-1}\|\hat{\beta} - \beta\|^2)\) by \(\hat{\Sigma} = O_p(1)\) and \(\sum_{i=1}^{N} \|\hat{\Sigma}^{1/2} \hat{\Gamma}_i \hat{\Sigma}^{-1/2}_{ii}\|^2 = r\).

Consider \(e_2\). Ignore the factor 4, it is bounded by

\[
CK\|\hat{\beta} - \beta\|^2 \cdot \|\hat{\Sigma}^{1/2} \hat{H}^{1/2}\|^2 \left( \sum_{i=1}^{N} \|\hat{\Sigma}^{1/2} \hat{\Gamma}_i \hat{\Sigma}^{-1/2}_{ii}\|^2 \right)
\]
\[
\times \left( \sum_{p=1}^{K} \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it}v_{jtp} - E(\varepsilon_{it}v_{jtp})] \right\|^2 \right)
\]

So \( e_2 = O_p(T^{-1} \| \hat{\beta} - \beta \|^2) \) by the similar arguments as \( e_1 \).

Consider \( e_3 \), which is bounded in norm by

\[
CK \frac{1}{N} \| \hat{\beta} - \beta \|^2 \cdot \| N^{1/2} \hat{H}^{1/2} \|_2^2 \left( \sum_{i=1}^{N} \| \hat{H}^{1/2} \hat{\Sigma}_{ii}^{-1/2} \|_2^2 \right)
\]

\[
\times \sum_{p=1}^{K} \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \| E(\varepsilon_{it}v_{jtp}) \|^2
\]

Notice \( E(\varepsilon_{it}v_{jtp}) = 0 \) if \( i \neq j \). So \( e_3 = O_p(N^{-1} \| \hat{\beta} - \beta \|^2) \). \( e_4 \) is of smaller order than \( e_2 \). Given these results, (b) follows.

Consider (c). The left hand side of (c) is equivalent to

\[
\frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \hat{\lambda}_t (\hat{\beta} - \beta) f_j \right\|^2 (\frac{1}{N} \sum_{j=1}^{N} \| \Gamma_j \|^2)
\]

which is bounded in norm by

\[
\left\| \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \hat{\lambda}_t (\hat{\beta} - \beta) f_j \right\|^2 (\frac{1}{N} \sum_{j=1}^{N} \| \Gamma_j \|^2)
\]

Since \( \frac{1}{N} \sum_{j=1}^{N} \| \Gamma_j \|^2 = O(1) \), we only need to consider \( \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \hat{\lambda}_t (\hat{\beta} - \beta) f_j \), which is equal to

\[
\sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1} \gamma_{ip} M_{ff}
\]

\[
+ \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \hat{H} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1} \gamma_{itp} f_t = e_5 + e_6 \quad \text{say}
\]

The term \( \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1} \gamma_{ip} M_{ff} \) is bounded in norm by

\[
\| N^{1/2} \hat{H}^{1/2} \| \left( \sum_{i=1}^{N} \| \hat{H} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1/2} \|_2^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii}^{-1/2} \gamma_{ip} \|_2^2 \right)^{1/2} \| M_{ff} \|
\]

which is \( O_p(1) \) by (A.4). So \( e_5 = O_p(\| \hat{\beta} - \beta \|) \). Consider \( e_6 \). The term \( \hat{H} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1} \gamma_{itp} f_t \) is bounded in norm by

\[
C \| N^{1/2} \hat{H}^{1/2} \| \left( \sum_{i=1}^{N} \| \hat{H} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1/2} \|_2^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \| \frac{1}{T} \sum_{t=1}^{T} f_t \|_2^2 \right)^{1/2}
\]
which is \( O_p(T^{-1/2}) \). So \( e_6 = o_p(\|\hat{\beta} - \beta\|) \). Given these results, (c) follows.

Consider (d). The left hand side of (d) is equivalent to

\[
\frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1} \frac{1}{T} \sum_{t=1}^{T} \hat{t}_{it}(\beta - \beta) \hat{\varepsilon}'_j \right\|^2
\]

which is equal to

\[
\frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\beta_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1} \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\varepsilon}_j \right\|^2
\]

By \( \hat{x}_{ip} = \gamma_{ip} f_t + \varepsilon_{it} \), the above expression is bounded by

\[
4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\beta_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1} \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\varepsilon}_j \right\|^2
\]

\[
+ 4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\beta_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1} \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\varepsilon}_j \right\|^2
\]

\[
+ 4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\beta_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1} \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\varepsilon}_j \right\|^2
\]

\[
+ 4 \frac{1}{N} \sum_{j=1}^{N} \left\| \sum_{p=1}^{K} (\beta_p - \beta_p) \hat{H} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1} \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\varepsilon}_j \right\|^2 = e_7 + e_8 + e_9 + e_{10}, \text{ say}
\]

Using (A.25), term \( e_7 \) is bounded by (ignoring the factor 4)

\[
K \|\hat{\beta} - \beta\|^2 \cdot \|N^{1/2} \hat{H}^{1/2}\| \left( \sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1/2}\| ^2 \right)
\]

\[
\times \left( \sum_{p=1}^{K} \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{\text{tie}}^{-1/2} \gamma_{ip}\| ^2 \right) \sum_{j=1}^{N} \frac{1}{N} \sum_{i=1}^{N} \| \hat{H}^{1/2} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1/2} \| ^2
\]

which is \( O_p(T^{-1}\|\hat{\beta} - \beta\|^2) \) by (A.4).

Consider \( e_8 \), which is bounded by (ignoring the factor 4)

\[
C K \|\hat{\beta} - \beta\|^2 \cdot \|N^{1/2} \hat{H}^{1/2}\| \left( \sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\lambda}_i \hat{\Sigma}_{\text{tie}}^{-1/2}\| ^2 \right)
\]

\[
\times \sum_{p=1}^{K} \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} [v_{itp} \hat{\varepsilon}'_j - E(v_{itp} \hat{\varepsilon}'_j)] \right\|^2
\]
which is $O_p(T^{-1/2}||\hat{\beta} - \beta||^2)$ by the similar arguments as $e_7$.

Consider $e_9$, which is bounded by (ignoring the factor 3)

$$CK \frac{1}{N} \||\hat{\beta} - \beta\||^2 \cdot N^{1/2} \hat{H}^{1/2} \left( \sum_{i=1}^{N} \||\hat{H}_{\lambda i}^1 \hat{\Sigma}_{iie}^{1/2}\||^2 \right) \sum_{p=1}^{K} \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \|E(v_{itp} \epsilon_{jt})\|^2$$

which is $O_p(N^{-1}||\hat{\beta} - \beta||^2)$ by $E(v_{itp} \epsilon_{jt}) = 0$ for $i \neq j$. Term $e_{10}$ is of smaller order than $e_8$. Given these results, we have (d).

Consider (e). The left hand side of (e) is equivalent to

$$\frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H} \sum_{i=1}^{N} \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \hat{\lambda}_{ii} \hat{\Sigma}_{iie}^{1/2} \hat{x}_{itp} \hat{x}_{jtp} \right\|^2$$

which is equal to

$$\frac{1}{N} \sum_{j=1}^{N} \left| \sum_{p=1}^{K} \sum_{q=1}^{K} (\hat{\beta}_p - \beta_p)(\hat{\beta}_q - \beta_q) \hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\lambda}_{ii} \hat{\Sigma}_{iie}^{1/2} \hat{x}_{itp} \hat{x}_{jtp} \right|^2$$

By (A.25), the above expression is bounded by

$$K^2 ||\hat{\beta} - \beta||^4 \left( \sum_{p=1}^{K} \sum_{q=1}^{K} \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\lambda}_{ii} \hat{\Sigma}_{iie}^{1/2} \hat{x}_{itp} \hat{x}_{jtp} \right\|^2 \right)$$

which is further bounded by

$$CK^2 ||\hat{\beta} - \beta||^4 \cdot N^{1/2} \hat{H}^{1/2} \left( \sum_{i=1}^{N} \||\hat{H}_{\lambda i}^1 \hat{\Sigma}_{iie}^{1/2}||^2 \right) \times \left( \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} \hat{x}_{itp} \hat{x}_{jtp} \right)^2 \right)$$

Notice, for $\forall i, t, p$, $E(\hat{x}_{itp}^4) \leq C$ for some sufficiently large constant $C$ by Assumptions B and C. So the above expression is $O_p(||\hat{\beta} - \beta||^4)$, obtaining (e).

**PROPOSITION A.1.** Under Assumptions A-D and the identification conditions IB, we have

$$A \equiv \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i^T \hat{H} = O_p(T^{-1/2}) + O_p\left( \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii} - \Sigma_{ii}||^2 \right)^{1/2}$$

$$+ O_p\left( \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii}^{-1}|| \cdot ||\hat{\Gamma}_i - \Gamma_i||^2 \right) + O_p(||\hat{\beta} - \beta||)$$
Proof of Proposition A.1. Notice the equation

\[
\frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} - \frac{1}{N} \Gamma' \Sigma_{ee}^{-1} \Gamma = \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma} - \Gamma)
\]

\[
+ \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} - \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma} - \Gamma) + \frac{1}{N} \Gamma' (\hat{\Sigma}_{ee}^{-1} - \Sigma_{ee}^{-1}) \Gamma
\]

Since \( \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \) and \( \frac{1}{N} \Gamma' \Sigma_{ee}^{-1} \Gamma \) are both diagonal matrices, we have

\[
\text{ndiag} \left\{ \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma} - \Gamma) \right\} + \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} = \text{ndiag} \left\{ \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma} - \Gamma) - \frac{1}{N} \Gamma' (\hat{\Sigma}_{ee}^{-1} - \Sigma_{ee}^{-1}) \Gamma \right\}
\]

where ndiag denotes the non-diagonal elements. However, by Lemmas A.5 and A.6, and noting \( M_{ff} = M_{ff} = I_r \), equation (A.21) can be simplified as

\[
\hat{H} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma} - \Gamma) + (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \hat{H} = O_p(T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|)
\]

which can be written as

\[
\hat{H} N \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma} - \Gamma) + \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \hat{H} N = O_p(T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|)
\]

where \( \hat{H} N = N \hat{H} \) is a diagonal matrix. We use \( \hat{q}_1, \hat{q}_2, \ldots, \hat{q}_r \) to denote its diagonal elements.

Equation (A.26) puts \( \frac{1}{2} r (r - 1) \) restrictions and equation (A.27) puts \( \frac{1}{2} r (r + 1) \) restrictions on the matrix \( \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \). So it can be uniquely determined by the equation system (A.26) and (A.27). Solving this equation system, we have

\[
\left[ \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \right]_{ii} = \frac{1}{2 \hat{q}_i} \left[ O_p(T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|) \right]
\]

for \( i = 1, 2, \ldots, r \), and

\[
\left[ \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \right]_{ij} = \frac{1}{\hat{q}_i - \hat{q}_j} \left[ O_p(T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|) \right]
\]

\[
- \hat{q}_i \left[ \frac{1}{N} (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma} - \Gamma) - \frac{1}{N} \Gamma' (\hat{\Sigma}_{ee}^{-1} - \Sigma_{ee}^{-1}) \Gamma \right]
\]

for \( i \neq j \). However, Corollary A.1(b) shows \( \hat{q}_j - q_j \overset{p}{\to} 0 \), where \( q_j \) is the \( j \)th diagonal element of \( Q \) defined in Assumption C.3. So we have

\[
(\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \hat{H} = O_p(T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|)
\]
which implies Proposition A.1 because $O_p\left(\frac{1}{N} \Gamma'(\hat{\Sigma}^{-1} - \Sigma^{-1})\right) = O_p\left(\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii} - \Sigma_{ii}||^{2} \right)$ by (A.14).

This completes the proof of Proposition A.1 \hfill \Box

**Proposition A.2.** Under the assumptions of Proposition 2.1, we have

\[ \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}^{-1}_{ii}\| \cdot \|\hat{\Gamma}_i - \Gamma_i\|^2 = O_p(T^{-1}) + O_p(||\hat{\beta} - \beta||^2) \]

\[ \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2 = O_p(T^{-1}) + O_p(||\hat{\beta} - \beta||^2) \]

**Proof of Proposition A.2.** Consider the first order condition (2.6).

Some algebra computation shows that (2.6) is equal to

\[ \hat{\Gamma}_j - \Gamma_j = -\hat{M}_{jj}^{-1} (\hat{M}_{jj} - M_{jj}) \Gamma_j - \hat{M}_{jj}^{-1} A' M_{jj} \Gamma_i \]

\[ + \hat{M}_{jj}^{-1} (I-A)' \frac{1}{T} \sum_{t=1}^{T} f_t \epsilon_j' \chi_t' \hat{f}_t' \Gamma_j + \hat{M}_{jj}^{-1} \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \epsilon_{j,t} \]

\[ - \hat{M}_{jj}^{-1} (I-A)' \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{f}_t' \Gamma_j + \hat{M}_{jj}^{-1} \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \epsilon_{j,t} \]

\[ - \hat{M}_{jj}^{-1} \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Sigma}_{ii} - \Sigma_{ii}) \]

where $\epsilon_{j,t} = \epsilon_{it} \epsilon_{jt} - E(\epsilon_{it} \epsilon_{jt})$. Notice that the identification condition requires $\hat{M}_{jj} = M_{jj} = I_r$. So the above result can be simplified as

\[ \hat{\Gamma}_j - \Gamma_j = -A' \Gamma_j + (I-A)' \frac{1}{T} \sum_{t=1}^{T} f_t \epsilon_j' \chi_t' \hat{f}_t' \Gamma_j \]

\[ + \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \epsilon_{j,t} - (I-A)' \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{f}_t' \Gamma_j \]

\[ - \hat{H} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{f}_t' \]\n
\[ - \hat{H} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{f}_t' \]
\[
\hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) f'_t \right] \Gamma_j - \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) \hat{e}_{jt} \right] 
\]

\[
+ \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) \hat{\beta}' \hat{e}_{jt} \right] - \hat{H} \hat{\Gamma}_j \hat{\Sigma}^{-1}_{ji} (\hat{\Sigma}_{ji} - \Sigma_{ji}) 
\]

The right hand side of (A.29) has 10 terms. We use \(a_{i1}, a_{i2}, \ldots, a_{i10} \) to denote them. So we have \(\| \hat{\Gamma}_i - \Gamma_i \| \leq \| a_{i1} \| + \| a_{i2} \| + \cdots + \| a_{i10} \| \). Using the fact \(\sum_{p=1}^{10} \| a_p \|^2 \leq 10 \sum_{p=1}^{10} \| a_p \|^2 \), we have

\[
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}^{-1}_{ii} \| \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 \leq C \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Gamma}_i - \Gamma_i \|^2
\]

\[
= 10C \frac{1}{N} \sum_{i=1}^{N} (\| a_{i1} \|^2 + \| a_{i2} \|^2 + \cdots + \| a_{i10} \|^2)
\]

The first inequality uses the fact that \(\| \hat{\Sigma}^{-1}_{ii} \| \) is bounded by some \(C \). Consider the first term, which is equal to \(\frac{1}{N} \sum_{j=1}^{N} \| A \Gamma_j \|^2 \). This term is bounded by \(\| A \|^2 \cdot \frac{1}{N} \sum_{j=1}^{N} \| \Gamma_j \|^2 \). Since \(A = O_p(T^{-1/2}) + O_p((\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2)^{1/2}) + O_p(\| \hat{\beta} - \beta \|) \) by Propositions A.1, together with \(\frac{1}{N} \sum_{j=1}^{N} \| \Gamma_j \|^2 = O(1) \), we have \(\frac{1}{N} \sum_{i=1}^{N} \| a_{i1} \|^2 = O_p(T^{-1}) + O_p((\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2) + O_p(\| \hat{\beta} - \beta \|^2) \).

The 2nd-4th and 10th terms are summarized in Lemma A.7. The 5th-9th terms are summarized in Lemma A.8. From these results, we have

\[
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} \| \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 = O_p(T^{-1}) + O_p(\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2)
\]

\[
+ O_p(\| \hat{\beta} - \beta \|^2)
\]

Now we turn to the first order condition (2.8). By (A.18) and (A.19), equation (2.8) is equivalent to

\[
\hat{\Sigma}_{ji} - \Sigma_{ji} = - (\hat{\lambda}_j - \lambda_j)' \hat{M}_{ff} (\hat{\lambda}_j - \lambda_j) - 2(\hat{\lambda}_j - \lambda_j)' \hat{M}_{ff} \lambda_j - \lambda_j' (\hat{M}_{ff} - M_{ff}) \lambda_j 
\]

\[
+ 2 \frac{1}{N} \sum_{t=1}^{T} \lambda_j f_t e_{jt} + \frac{1}{T} \sum_{t=1}^{T} (e_{jt}^2 - \Sigma_{ji}) - 2 \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \frac{1}{T} \sum_{t=1}^{T} \lambda_j f_t \hat{x}_{jtp} 
\]

\[-2 \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \left[ \frac{1}{T} \sum_{t=1}^{T} e_{jt} \hat{x}_{jtp} \right] + \sum_{p=1}^{K} \sum_{q=1}^{K} \hat{x}_{jtp} (\hat{\beta}_p - \beta_p) (\hat{\beta}_q - \beta_q) \hat{x}_{jtp} 
\]

\[
+ \sum_{p=1}^{K} \sum_{q=1}^{K} \hat{x}_{jtp} (\hat{\beta}_p - \beta_p) (\hat{\beta}_q - \beta_q) \hat{x}_{jtp}
\]
Consider (A.32). The second term of right hand side of (A.32) involves \( \hat{\lambda}_j - \lambda_j \). The third term involves \( \hat{M}_{ff} - M_{ff} \). But the expression of \( \hat{\lambda}_j - \lambda_j \) is given in (A.29) (the first column) and \( \hat{M}_{ff} - M_{ff} \) is given in (A.21)\(^1\). Using (A.29) and (A.21) to replace \( \hat{\lambda}_j - \lambda_j \) and \( \hat{M}_{ff} - M_{ff} \) from (A.32), we have

\begin{equation}
\hat{\Sigma}_{jj} = \Sigma_{jj} = \frac{1}{T} \sum_{t=1}^{T} [v_{jtx} v_{jtx}' - \Sigma_{jj}] - (\hat{\gamma}_{jx} - \gamma_{jx})^T \hat{M}_{ff} (\hat{\gamma}_{jx} - \gamma_{jx})
\end{equation}

(A.33)  

\(- (\hat{\gamma}_{jx} - \gamma_{jx})^T \hat{M}_{ff} \gamma_{jx} - \gamma_{jx}^T \hat{M}_{ff} (\hat{\gamma}_{jx} - \gamma_{jx}) - \gamma_{jx} (\hat{M}_{ff} - M_{ff}) \gamma_{jx} + \frac{1}{T} \sum_{t=1}^{T} \gamma_{jx}' f_t v_{jtx}' + \frac{1}{T} \sum_{t=1}^{T} v_{jtx} f_t \gamma_{jx}
\)

Consider (A.32). The second term of right hand side of (A.32) involves \( \hat{\lambda}_j - \lambda_j \). The third term involves \( \hat{M}_{ff} - M_{ff} \). But the expression of \( \hat{\lambda}_j - \lambda_j \) is given in (A.29) (the first column) and \( \hat{M}_{ff} - M_{ff} \) is given in (A.21)\(^1\). Using (A.29) and (A.21) to replace \( \hat{\lambda}_j - \lambda_j \) and \( \hat{M}_{ff} - M_{ff} \) from (A.32), we have

\begin{equation}
\hat{\Sigma}_{jj} - \Sigma_{jj} = \frac{1}{T} \sum_{t=1}^{T} (\hat{\varepsilon}_{it}^2 - \Sigma_{jj}) - (\hat{\lambda}_j - \lambda_j)' (\hat{\lambda}_j - \lambda_j) + \lambda_j^T A^T \lambda_j
\end{equation}

(A.34)  

\(+ 2 \lambda_j^T A^T \frac{1}{T} \sum_{t=1}^{T} f_t e_{jt} - 2 \lambda_j^T A^T \frac{1}{T} \sum_{t=1}^{N} \sum_{i=1}^{T} f_t \varepsilon_{it}' \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i \hat{H} \lambda_j
\)

\(+ \lambda_j^T \hat{H} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j \hat{H} \lambda_j - \lambda_j^T \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Sigma}_{ii} - \Sigma_{ii}) \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j \hat{H} \lambda_j
\)

\(- 2 \lambda_j^T \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} e_{jt} - E (\varepsilon_{it} e_{jt})] + 2 \lambda_j^T \hat{H} \lambda_j \hat{\Sigma}_{jj}^{-1} (\hat{\Sigma}_{jj} - \Sigma_{jj}) + O_{j1}
\)

where \( A = (\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{cc}^{-1} \hat{\Gamma} \hat{H}, e_{jt} = \varepsilon_{jt} - E (\varepsilon_{jt}) \) and \( O_{j1} \) is defined as

\( O_{j1} = -2 \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \lambda_j A^T \frac{1}{T} \sum_{t=1}^{T} f_t \hat{x}_{ip} [\frac{1}{T} \sum_{t=1}^{T} \chi_t (\hat{\beta} - \beta)' \hat{\xi}_t] \hat{H} N \lambda_j
\)

\(- 2 \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{jt} \hat{x}_{tp} + \sum_{p=1}^{K} \sum_{q=1}^{K} \frac{1}{T} \sum_{t=1}^{T} \hat{x}_{tp} (\hat{\beta}_p - \beta_p) (\hat{\beta}_q - \beta_q) \hat{x}_{tq}
\)

\(+ 2 \lambda_j A^T \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{\xi}_t \hat{H} N \lambda_j + 2 \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \lambda_j \hat{H} N \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t \hat{x}_{tp}
\)

\(^1\)Under the identification condition \( \hat{M}_{ff} = M_{ff} = I_r \), we immediately obtain that \( \hat{M}_{ff} - M_{ff} = 0 \). However, this result, while simple, is not as useful as it looks. If we use this result, we would face the term \( \hat{H} \Gamma^{\hat{\Sigma}^{-2}} (\hat{\Gamma} - \Gamma) + (\hat{\Gamma} - \Gamma)^{\hat{\Sigma}^{-2}} \hat{H} \) in the subsequent expression, which still requires invoke (A.21).
Applying similar arguments to (A.33), we have

\[
\Sigma_{jix} - \Sigma_{ijx} = \frac{1}{T} \sum_{t=1}^{T} (v_{jtx}v_{jtx} - \Sigma_{jix}) - \gamma'_{jx} \hat{H} \sum_{i=1}^{N} \Gamma_i \hat{\Sigma}_{ii}^{-1}(\Sigma_{ii} - \Sigma_{ij}) \hat{\Gamma}_i \hat{H} \gamma_{jx} \\
- \gamma'_{jx} A' \left[ \frac{1}{T} \sum_{t=1}^{T} f_i \hat{\chi}_i \right] \hat{H} \gamma_{jx} - \gamma'_{jx} \hat{H} \sum_{t=1}^{T} \hat{\chi}_t f'_t A \gamma_{jx} + \gamma'_{jx} A' \frac{1}{T} \sum_{t=1}^{T} f_t v'_{jtx} \\
+ \frac{1}{T} \sum_{t=1}^{T} v_{jtx} f_t A \gamma_{jx} - \gamma'_{jx} \hat{H} \sum_{i=1}^{N} \Gamma_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \eta_{j,t} - \sum_{i=1}^{N} \frac{1}{T} \sum_{t=1}^{T} \eta_{j,t} \hat{\Sigma}_{ii}^{-1} \Gamma_i \hat{H} \gamma_{jx} \\
+ \gamma'_{jx} \hat{H} \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}' \hat{\Sigma}_{ee}^{-1} (\Sigma_{ee} - \Sigma_{jix}) \hat{\Gamma} \hat{H} \gamma_{jx} + \gamma'_{jx} A' A \gamma_{jx} - (\hat{\gamma}_{jx} - \gamma_{jx})'(\hat{\gamma}_{jx} - \gamma_{jx}) + O_{j2}
\]

(A.36)

where \( \eta_{j,t} = \epsilon_{it} v'_{jtx} - E(\epsilon_{it} v'_{jtx}) \) and \( O_{j2} \) is defined as

\[
O_{j2} = \gamma'_{jx} A' \left[ \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta) \hat{\xi}'_t \right] \hat{H} \gamma_{jx} + \gamma'_{jx} \hat{H} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) f'_t A \gamma_{jx} \\
- \gamma'_{jx} \hat{H} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) \hat{\xi}'_t \hat{H} \gamma_{jx} - \gamma'_{jx} \hat{H} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) \hat{\chi}'_t \hat{H} \gamma_{jx} \\
+ \gamma'_{jx} \hat{H} \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) v'_{jtx} + \frac{1}{T} \sum_{t=1}^{T} v_{jtx} (\hat{\beta} - \beta) \hat{\xi}'_t \hat{H} \gamma_{jx}
\]

(A.37)

Terms \( O_{j1} \) and \( O_{j2} \) depend on \( \hat{\beta} - \beta \). By the special structure of \( \Sigma_{ii} \), we have

\[
\frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jj} - \Sigma_{jj} \|^2 = \frac{1}{N} \sum_{j=1}^{N} (\hat{\Sigma}_{jje} - \Sigma_{jje})^2 + \frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jix} - \Sigma_{jix} \|^2
\]
Consider \( \frac{1}{N} \sum_{j=1}^{N} (\Sigma_{jie} - \Sigma_{jie})^2 \). The right hand side of (A.34) has 10 terms and we use \( b_{j1}, b_{j2}, \cdots, b_{j9} \) to denote the first 9 terms. By the Cauchy-Schwarz inequality, we have

\[
\frac{1}{N} \sum_{j=1}^{N} (\Sigma_{jie} - \Sigma_{jie})^2 \leq 10 \frac{1}{N} \sum_{j=1}^{N} (\|b_{j1}\|^2 + \cdots + \|b_{j9}\|^2 + \|O_{j1}\|^2)
\]

The first term is \( O_p(T^{-1}) \) can be easily verified. For the second term, we need to use (A.29) (the first column) to substitute \( \hat{\lambda}_j - \lambda_j \). By a little tedious computation, we have

\[
\frac{1}{N} \sum_{j=1}^{N} (\Sigma_{jie} - \Sigma_{jie})^2 = O_p(T^{-1}) + o_p\left(\frac{1}{N} \sum_{j=1}^{N} \|\hat{\Sigma}_{ii}^{-1}\| \cdot \|\hat{\Gamma}_i - \Gamma_i\|^2\right)
\]

The above result, together with Lemma A.9 below, implies

\[
\frac{1}{N} \sum_{j=1}^{N} (\Sigma_{jie} - \Sigma_{jie})^2 = O_p(T^{-1}) + o_p\left(\frac{1}{N} \sum_{j=1}^{N} \|\hat{\Sigma}_{ii}^{-1}\| \cdot \|\hat{\Gamma}_i - \Gamma_i\|^2\right) + O_p(\|\hat{\beta} - \beta\|^2)
\]

Similar arguments are applicable for \( \frac{1}{N} \sum_{j=1}^{N} \|\hat{\Sigma}_{ix} - \Sigma_{ix}\|^2 \). So we have

\[
\frac{1}{N} \sum_{j=1}^{N} \|\hat{\Sigma}_{ij} - \Sigma_{ij}\|^2 = O_p(T^{-1}) + o_p\left(\frac{1}{N} \sum_{j=1}^{N} \|\hat{\Sigma}_{ii}^{-1}\| \cdot \|\hat{\Gamma}_i - \Gamma_i\|^2\right) + O_p(\|\hat{\beta} - \beta\|^2)
\]

Substituting (A.31) into (A.38), we have \( \frac{1}{N} \sum_{j=1}^{N} \|\hat{\Sigma}_{ij} - \Sigma_{ij}\|^2 = O_p(T^{-1}) + O_p(\|\hat{\beta} - \beta\|^2) \). Substituting this result into (A.31), we get the remaining result of Proposition A.2. This completes the proof of Proposition A.2.

**Corollary A.2.** Under Assumptions A-D and the identification conditions IB, we have

\[
\sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i)\hat{\Sigma}_{ii}^{-1}\hat{\Gamma}_i'\hat{H} = O_p(T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|)
\]

**Lemma A.9.** Under Assumptions A-D,

\[
(a) \quad \frac{1}{N} \sum_{j=1}^{N} \|O_{j1}\|^2 = o_p(\|\hat{\beta} - \beta\|^2)
\]

\[
(b) \quad \frac{1}{N} \sum_{j=1}^{N} \|O_{j2}\|^2 = o_p(\|\hat{\beta} - \beta\|^2)
\]
where $O_{j1}$ is defined in (A.35) and $O_{j2}$ defined in (A.37) above.

**Proof of Lemma A.9.** Consider (a). There are 9 terms on the left hand side of (A.34). We use $c_{j1}, c_{j2}, \cdots c_{j9}$ to denote them. By the Cauchy-Schwarz inequality, we have

$$
\frac{1}{N} \sum_{j=1}^{N} \|O_{j1}\|^2 \leq \frac{9}{N} \sum_{j=1}^{N} (\|c_{j1}\|^2 + \|c_{j2}\|^2 + \cdots + \|c_{j9}\|^2)
$$

Now we check the terms one by one. Consider (a). There are 9 terms on the left hand side of (A.34). We use $c_{j1}, c_{j2}, \cdots c_{j9}$ to denote them. By the Cauchy-Schwarz inequality, we have

$$
\frac{1}{N} \sum_{j=1}^{N} \|c_{j1}\|^2 \leq 4K \|\hat{\beta} - \beta\|^2 \cdot \|A\|^2 \frac{1}{N} \sum_{j=1}^{N} \|T \sum_{t=1}^{T} f_{it} \hat{x}_{jtp}\|^{2}
$$

But $\frac{1}{N} \sum_{j=1}^{N} \|c_{j1}\|^2$ is $O_p(1)$ and $\|A\| = o_p(1)$. Thus $\frac{1}{N} \sum_{j=1}^{N} \|c_{j1}\|^2 = O_p(\|\hat{\beta} - \beta\|^2)$.

The second term $\frac{1}{N} \sum_{j=1}^{N} \|c_{j2}\|^2$ and the third term $\frac{1}{N} \sum_{j=1}^{N} \|c_{j3}\|^2$ are both $o_p(\|\hat{\beta} - \beta\|^2)$ given that $\frac{1}{N} \sum_{j=1}^{N} \|T \sum_{t=1}^{T} x_{jtp} \hat{x}_{jtp}\|^{2} = O_p(1)$ and $\frac{1}{N} \sum_{j=1}^{N} \|T \sum_{t=1}^{T} c_{jt} \hat{x}_{jtp}\|^{2} = O_p(1)$.

The 4th, 5th and 7th terms are all $o_p(\|\hat{\beta} - \beta\|^2)$ which are implicitly implied by the results in Lemma A.6 and Proposition 2.1.

Consider $c_{j6}$. Term $\lambda_{j}^{i} H \sum_{i=1}^{N} \hat{\Gamma}_{i} \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} \hat{x}_{jtp}$ is equal to $\text{tr}[H \sum_{i=1}^{N} \hat{\Gamma}_{i} \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} \hat{x}_{jtp}]$. So we have

$$
\frac{1}{N} \sum_{j=1}^{N} \|c_{j6}\|^2 \leq 4K \sum_{p=1}^{K} \|\hat{\beta} - \beta\| \|H \sum_{i=1}^{N} \hat{\Gamma}_{i} \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} f_{it}\| (\frac{1}{N} \sum_{j=1}^{N} \|\gamma_{jp}\|^2 \|\lambda_{j}\|^2)
$$

By $\hat{x}_{jtp} = \gamma_{jp} f_{t} + \hat{v}_{jtp}$, the above term is further bounded by

$$
16K \sum_{p=1}^{K} \|\hat{\beta} - \beta\|^2 \|H \sum_{i=1}^{N} \hat{\Gamma}_{i} \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} f_{it}\| (\frac{1}{N} \sum_{j=1}^{N} \|\gamma_{jp}\|^2 \|\lambda_{j}\|^2)
$$

$$
+ 16K \sum_{p=1}^{K} \|\hat{\beta} - \beta\|^2 \frac{1}{N} \sum_{j=1}^{N} \|H \sum_{i=1}^{N} \hat{\Gamma}_{i} \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} v_{jtp} - E(\varepsilon_{it} v_{jtp})\| \lambda_{j}\|^2
$$

$$
+ 16K \sum_{p=1}^{K} \|\hat{\beta} - \beta\|^2 \frac{1}{N} \sum_{j=1}^{N} \|H \sum_{i=1}^{N} \hat{\Gamma}_{i} \hat{\Sigma}_{ii}^{-1} E(\varepsilon_{it} v_{jtp})\| \lambda_{j}\|^2
$$
The first term of the above is $o_p(||\hat{\beta} - \beta||^2)$ by $\hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} \varepsilon_{it} \hat{v}_{jp} \lambda_j'$ which is implied by Lemma A.6(a). The second term is also $o_p(||\hat{\beta} - \beta||^2)$ because term $\frac{1}{N} \sum_{i=1}^{N} |\hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} \hat{v}_{jp} - E(\varepsilon_{it} \hat{v}_{jp})] \lambda_j'|^2$ is bounded by

$$C ||\hat{H}||^2 \left( \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}||^2 \right) \left( \frac{1}{N^2} \sum_{j=1}^{N} \sum_{i=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} \hat{v}_{jp} - E(\varepsilon_{it} \hat{v}_{jp})] \lambda_j' \right)^2 \right)$$

which is $O_p(T^{-1})$. The third term is also $o_p(||\hat{\beta} - \beta||^2)$ since $\sum_{i=1}^{N} ||\hat{H}^{1/2} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1/2}||^2 = r$. The last term is of smaller order than the second one. Given these results, we have $\frac{1}{N} \sum_{j=1}^{N} ||c_{jt}||^2 = o_p(||\hat{\beta} - \beta||^2)$.

The 8th term is $o_p(||\hat{\beta} - \beta||^2)$ which can be proved similarly as the 6th term. The last term is $o_p(||\hat{\beta} - \beta||^2)$, which can be proved similarly as the 7th term. We thus obtain (a).

Result (b) can be proved similarly as (a).
for all $m, n = 1, 2, \cdots, K$. By $\dot{x}_{jtp} = \gamma'_{j/ip} f_t + \dot{\nu}_{jtp}$, equation (A.39) is equivalent to

\begin{equation}
(A.41) \quad \vartheta_{p1}(\dot{\beta}_1 - \beta_1) + \vartheta_{p2}(\dot{\beta}_2 - \beta_2) + \cdots + \vartheta_{pK}(\dot{\beta}_K - \beta_K)
= -\text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \gamma_{ip} (\hat{\lambda}_i - \lambda_i)' M_{ff} \right] + \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{T} \hat{\Sigma}_{iie}^{-1} \lambda_{ip}' t v_{itp} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{T} \hat{\Sigma}_{iie}^{-1} \lambda_{ip}' t v_{itp} - \text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \lambda_{ip}' (\hat{\lambda}_i - \lambda_i)' M_{ff} \right]
\end{equation}

\[+ \text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \gamma_{ip} (\hat{\lambda}_i - \lambda_i)' G \right] - \text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \lambda_{ip}' (\hat{\lambda}_i - \lambda_i)' G \right]\]

where $I_{K+1}^{p+1}$ is the $(p + 1)$th column of the identity matrix $I_{K+1}$. We neglect term $\frac{1}{NT} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \varepsilon_{itp}$ and $\text{tr}\left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{T} \sum_{t=1}^{T} \hat{\Gamma}_j \hat{\Sigma}_{jii}^{-1} (\Sigma_{ii} - \Sigma_{ji}) \frac{1}{I_{K+1}^{p+1}} \hat{\Sigma}_{iie}^{-1} \lambda_j' \hat{\lambda}_j \right]$ since the former is of smaller order than the $7$th term and the latter is of smaller order than the last term. The first term on the right involves $\hat{\lambda}_j - \lambda_j$. However, the expression of $\hat{\lambda}_j - \lambda_j$ has been given in (A.29) (the first column).

Using (A.29) to replace $\lambda_j - \lambda_j$ from (A.41), we have

\begin{equation}
(A.42) \quad \vartheta_{p1}(\dot{\beta}_1 - \beta_1) + \vartheta_{p2}(\dot{\beta}_2 - \beta_2) + \cdots + \vartheta_{pK}(\dot{\beta}_K - \beta_K)
= \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{T} \hat{\Sigma}_{iie}^{-1} \varepsilon_{itp} v_{itp} - \text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{T} \hat{\Gamma}_j \hat{\Sigma}_{jii}^{-1} [\varepsilon_{itp} v_{itp} - E(\varepsilon_{itp} v_{itp})] \right] \hat{\Sigma}_{iie}^{-1} \lambda_j' \hat{\lambda}_j \right]
\end{equation}

\[+ \text{tr}\left[ \frac{1}{T} \sum_{t=1}^{T} \hat{\Gamma}_t \hat{\Sigma}_{te}^{-1} (\varepsilon_{te} - \Sigma_{te}) \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}_H \hat{v}_p \right] - \text{tr}\left[ (I - A)' \frac{1}{T} \sum_{t=1}^{T} f_i (\hat{\beta} - \beta)' \hat{\xi}_t \hat{H}_N \hat{v}_p \right]
\end{equation}

\[+ \text{tr}\left[ \frac{1}{T} \sum_{t=1}^{T} f_i (\hat{\beta} - \beta)' \hat{\xi}_t \hat{H}_N \hat{v}_p \right] \]
where $\hat{v}_p$ is defined in Table 1 and $J_{p1}$ and $J_{p2}$ are defined by

$$J_{p1} = \text{tr} \left[ G^2 (\hat{f} - \Gamma)^T \hat{\Sigma}_{\hat{\epsilon} e l} \hat{f} \hat{\Gamma} \frac{1}{N} \sum_{j=1}^{N} \hat{\Sigma}_{\hat{j} e l}^{-1} \hat{\lambda}_j \hat{\gamma}_{jp} \right] + \text{tr} \left[ A \frac{1}{N} \sum_{j=1}^{N} \hat{\Sigma}_{\hat{j} e l}^{-1} (\hat{\lambda}_j - \lambda_j) \hat{\gamma}_{jp} \right]
$$

$$+ \text{tr} \left[ H_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t \hat{f}_t \right) A \hat{v}_p \right] + \text{tr} \left[ \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{\hat{j} e l}^{-1} f_t \hat{\lambda}_j v_{tp} \hat{G} \hat{\Gamma}^T \hat{\Sigma}_{\hat{\epsilon} e l} (\hat{f} - \Gamma) \right]
$$

$$+ \text{tr} \left[ A' \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{\hat{j} e l}^{-1} f_t \hat{\gamma}_{jp} e_{jt} \right] - \text{tr} \left[ \sum_{i=1}^{N} \hat{\Sigma}_{ii}^{-1} (\hat{\lambda}_i - \lambda_i) \hat{\gamma}_{ip} \hat{G} \left( \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\chi}_t \right) \right]
$$

$$+ \text{tr} \left[ \hat{v}_p \hat{G} \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\chi}_t \right) \right] + \text{tr} \left[ A' A \hat{v}_p \hat{M} f_f \right] - \text{tr} \left[ A' \left( \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\chi}_t \right) \hat{H}_N \hat{v}_p \right]
$$

$$- \text{tr} \left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\gamma}_{ip} \hat{\gamma}_{ip} \right] + \text{tr} \left[ \frac{1}{N} \hat{H} \sum_{j=1}^{N} \hat{\lambda}_j \hat{\Sigma}_{\hat{j} e l}^{-1} (\hat{\Sigma}_{\hat{j} e l} \hat{\Sigma}_{\hat{j} e l}^{-1} - \hat{\Sigma}_{\hat{j} e l}) \hat{\gamma}_{jp} \right]
$$

$$+ \text{tr} \left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\gamma}_{ip} \hat{\gamma}_{ip} \right] - \text{tr} \left[ H \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Sigma}_{ii} - \Sigma_{ii}) \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i \hat{H} \hat{v}_p \right]
$$

and

$$J_{p2} = - \text{tr} \left[ H_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)' \hat{\xi}_t \right) \hat{H}_N \hat{v}_p \right] - \text{tr} \left[ \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)' \hat{\xi}_t \right) \hat{H}_N \hat{v}_p \right]
$$

$$+ \text{tr} \left[ \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)' \hat{\xi}_t \right) \hat{H}_N \hat{v}_p \right] + \text{tr} \left[ \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)' \hat{\xi}_t \right) \hat{H}_N \hat{v}_p \right]
$$

$$+ \text{tr} \left[ \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)' \hat{\xi}_t \right) \hat{H}_N \hat{v}_p \right]
$$

Equation (A.42) is dealt with in detail by the three lemmas below. The terms involving $\hat{\beta} - \beta$ are analyzed in Lemma A.10. The terms without $\hat{\beta} - \beta$ are
summarized in Lemma A.12. The residual terms $J_{p1}$ and $J_{p2}$ are treated in Lemma A.11.

Let $\Omega_{pq} = \frac{1}{N} \sum_{i=1}^{N} \Sigma_{iix}^{-1}(p,q)$, by Lemmas A.10, A.11 and A.12, we have

$$\Omega_{p1}(\hat{\beta}_1 - \beta_1) + \Omega_{p2}(\hat{\beta}_2 - \beta_2) + \cdots + \Omega_{pK}(\hat{\beta}_K - \beta_K) = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + O_p(N^{-1}T^{-1/2})$$

The above equation hold for all $p = 1, 2, \cdots, K$. So we have

$$\Omega(\hat{\beta} - \beta) = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + O_p(N^{-1}T^{-1/2})$$

Given the above result, in combination with Proposition A.2, we obtain the remaining three results of the theorem. This completes the proof of Theorem 2.1  

**Lemma A.10.** Under Assumptions A-D, we have

(a) $\vartheta_{pq} = \text{tr}(\omega_{qp}) - \text{tr}(v_q'Q^{-1}v_q) + \Omega_{pq} + o_p(1)$

(b) $\text{tr}[(I - A)\left(\frac{1}{T} \sum_{t=1}^{T} f_t(\hat{\beta} - \beta)' \hat{\Sigma}_{iix}^{-1}(\hat{\beta}' + \omega_{iix})\right)]$

$$= \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}(v_q'Q^{-1}v_q) + o_p(\|\hat{\beta} - \beta\|)$$

(c) $\text{tr}[(I - A)\left(\frac{1}{T} \sum_{t=1}^{T} f_t(\hat{\beta} - \beta)' \left(\frac{1}{N} \sum_{j=1}^{N} x_{itj}' \hat{\Sigma}_{iix}^{-1}(\hat{\beta} + \omega_{iix})\right)\right)]$

$$= \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}(\omega_{qp}) + o_p(\|\hat{\beta} - \beta\|)$$

where $\vartheta_{pq}$ is defined in (A.40), $\omega_{pq} = \frac{1}{N} \sum_{i=1}^{N} \gamma_{ip}' \hat{\Sigma}_{iix}^{-1}(\hat{\beta}' + \omega_{iix})$, $v_p = \frac{1}{N} \sum_{i=1}^{N} \lambda_i \hat{\Sigma}_{iix}^{-1}(\hat{\beta}' + \omega_{iix})$ and $\Omega_{pq} = \frac{1}{N} \sum_{i=1}^{N} \Sigma_{iix}^{-1}(p,q)$ with $\Sigma_{iix}$ being the $(p,q)$ entry of the matrix $\Sigma_{iix}$.

**Proof of Lemma A.10.** Consider (a). By the definition of $\vartheta_{pq}$,

$$\vartheta_{pq} = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iix}^{-1}\hat{x}_{itp}\hat{x}_{itq} - \frac{1}{NT} \sum_{i=1}^{N} \left( \sum_{t=1}^{T} \hat{\Sigma}_{iix}^{-1}\hat{x}_{itp}\hat{\lambda}_i \right) \hat{G} \left( \sum_{j=1}^{N} \hat{\Sigma}_{iix}^{-1}\hat{x}_{ij}\hat{x}_{ijq} \right)$$

Consider the first term of the above expression. It can be split into two terms (note that both $\Sigma_{iix}$ and $\Sigma_{iix}$ are scalars)

$$\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iix}^{-1}\hat{x}_{itp}\hat{x}_{itq} - \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \frac{\Sigma_{iix} - \Sigma_{iix}}{\Sigma_{iix}} \hat{x}_{itp}\hat{x}_{itq} = a_1 - a_2,$$
The term $a_1$ is equal to
\[
\frac{1}{N} \sum_{i=1}^{N} \Sigma_{i\ell e}^{-1} \gamma_{ip} \gamma_{iq} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{i\ell e}^{-1} \gamma_{ip} f_{v_{itq}} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{i\ell e}^{-1} \gamma_{ip} f_{v_{itq}}
\]
\[
+ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{i\ell e} (v_{itp} v_{itq} - \Sigma_{iix} (p, q)) + \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i\ell e}^{-1} \gamma_{ip} \omega_{iix} - \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i\ell e}^{-1} \bar{v}_{ip} \bar{v}_{iq}
\]

The first term of the above is $\text{tr}(\omega_{qp})$ by the definition of $\omega_{qp}$. The 2th-4th terms are $O_p(N^{-1/2}T^{-1/2})$. The 5th term is $\Omega_{pq}$. The last term is $O_p(T^{-1})$. Thus, $a_1 = \text{tr}(\omega_{qp}) + \Omega_{pq} + o_p(1)$.

The term $a_2$ is equal to
\[
\frac{1}{N} \sum_{i=1}^{N} \Sigma_{i\ell e} - \Sigma_{i\ell e} (p, q) + \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i\ell e} - \Sigma_{i\ell e} \gamma_{ip} \bar{v}_{itq} \sum_{t=1}^{T} f_{v_{itq}}
\]
\[
+ \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i\ell e} - \Sigma_{i\ell e} \gamma_{iq} \bar{v}_{itq} \sum_{t=1}^{T} f_{v_{itq}} - \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i\ell e} - \Sigma_{i\ell e} \bar{v}_{ip} \bar{v}_{iq}
\]

By the boundedness of $\Sigma_{i\ell e}, \Sigma_{i\ell s}, \Sigma_{i\ell s} (p, q), \gamma_{ip}$ and $M_{ff}$, the first and 3rd terms of the above expression are bounded in norm by
\[
C \left( \frac{1}{N} \sum_{i=1}^{N} (\Sigma_{i\ell e} - \Sigma_{i\ell e})^2 \right)^{1/2} \leq C \left( \frac{1}{N} \sum_{i=1}^{N} || \Sigma_{ii} - \Sigma_{ii} ||^2 \right)^{1/2}
\]
which is $o_p(1)$ by Proposition 2.1. The second term, by similar arguments, is bounded in norm by
\[
C \left( \frac{1}{N} \sum_{i=1}^{N} || \Sigma_{ii} - \Sigma_{ii} ||^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} f_{v_{itq}} \right)^2 \right)^{1/2}
\]
which is $o_p(T^{-1/2})$. The 4th term is bounded in norm by
\[
C \left( \frac{1}{N} \sum_{i=1}^{N} || \Sigma_{ii} - \Sigma_{ii} ||^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} v_{itp} v_{itq} - \Sigma_{iix} (p, q) \right)^2 \right)^{1/2}
\]
which is $O_p(T^{-1})$. The 5th term can be proved to be $o_p(1)$ similarly as the second one. The last term is of smaller order than the 4th term. Given these results, we have $a_2 = o_p(1)$. 
Consider the term $\frac{1}{NT} \sum_{t=1}^{T} (\sum_{i=1}^{N} \hat{\Sigma}_{tie}^{-1} \hat{\Sigma}_{ip} \hat{\lambda}_i \hat{G} (\sum_{j=1}^{N} \hat{\Sigma}_{jie}^{-1} \hat{\lambda}_j \hat{x}_{jtq})), which is equal to

\[
\text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{tie}^{-1} \hat{\lambda}_i \hat{G} \sum_{j=1}^{N} \hat{\lambda}_j \hat{\Sigma}_{jie}^{-1} \hat{\gamma}_{jq} \right] + \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} f_{itp} \hat{\Sigma}_{tie}^{-1} \hat{\lambda}_i \hat{G} \sum_{j=1}^{N} \hat{\lambda}_j \hat{\Sigma}_{jie}^{-1} \hat{\gamma}_{jq} \right]
\]

+ \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \lambda_i \hat{\Sigma}_{tie}^{-1} \hat{\gamma}_{jq} \hat{v}_{jtp} \hat{f}_t \right] + \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{tie}^{-1} \hat{\gamma}_{jq} \hat{G} \right]

- \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{tie}^{-1} \hat{\gamma}_{jq} \hat{\Sigma}_{jie}^{-1} \hat{\lambda}_j \hat{G} \right] = a_3 + \cdots + a_8, \text{ say}

Consider $a_3.$ The term $\frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{tie}^{-1} \hat{\lambda}_i$ is equal to

\[
\frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{tie}^{-1} \hat{\lambda}_i - \frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} (\hat{\Sigma}_{tie}^{-1} - \hat{\Sigma}_{tie}^{-1}) \hat{\lambda}_i + \frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{tie}^{-1} (\hat{\lambda}_i - \lambda_i)'
\]

The first term is $v'_p.$ The second term can be proved to be bounded in norm by $C(\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2)^{1/2}$ similarly as (A.14). The third term is bounded in norm by

\[
\left( \frac{1}{N} \sum_{i=1}^{N} \| \gamma_{ip} \hat{\Sigma}_{tie}^{-1/2} \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{tie}^{-1} \| \| \hat{\lambda}_i - \lambda_i \|^2 \right)^{1/2}
\]

\[
\leq \left( \frac{1}{N} \sum_{i=1}^{N} \| \gamma_{ip} \hat{\Sigma}_{tie}^{-1/2} \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{tie}^{-1} \| \| \hat{\Gamma}_i - \Gamma_i \|^2 \right)^{1/2}
\]

which is $O_p(T^{-1/2}) + O_p(\| \hat{\beta} - \beta \|)$ by Proposition A.2. Thus,

(A.45) \[
\frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{tie}^{-1} \hat{\lambda}_i = v_p + o_p(1), \quad \hat{v}_p = v_p + o_p(1)
\]

Note $\hat{G}_N - Q^{-1} = o_p(1),$ so we have $a_3 = \text{tr}[v'_p Q^{-1} v_q] + o_p(1)$.

Consider $a_4.$ The term $\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} f_{itp} \hat{\Sigma}_{tie}^{-1} \hat{\lambda}_i$ is bounded in norm by

\[
C \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\lambda}_i \hat{\Sigma}_{tie}^{-1} \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \| f_{itp} \|^2 \right)^{1/2}
\]
which is $O_p(T^{-1/2})$. So $a_4$ is $O_p(T^{-1/2})$ by (A.45) and $\hat{G}_N - Q^{-1} = o_p(1)$.

The term $a_5$ can be proved similarly as $a_4$.

Consider $a_6$, which is bounded in norm by

$$C \sup_{i \leq N} |\Sigma_{iix}^{(p,q)}| \frac{1}{N} \text{tr} \left( \left( \sum_{i=1}^{N} \|\hat{H}^{1/2} \hat{\lambda}_i \hat{\Sigma}_{iix}^{-1/2} \|_2^2 \right) \cdot \|\left(\hat{H}^{1/2} M_f^{-1} \hat{H}^{1/2} + I\right)^{-1}\| \right)$$

which is $O_p(N^{-1})$.

Consider $a_7$, the term inside the trace operator is bounded in norm by

$$C \left( \frac{1}{N} \sum_{i=1}^{N} \|\hat{\lambda}_i \hat{\Sigma}_{iix}^{-1/2} \|_2^2 \right) \left( \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} [v_{itp} v_{jqt} - E(v_{itp} v_{jqt})] \right)^2 \right)^{1/2} \|\hat{G}_N\|$$

The above term is $O_p(T^{-1/2})$ by $\hat{G}_N - Q^{-1} = o_p(1)$. So we have $a_7 = O_p(T^{-1/2})$. Term $a_8$ is of smaller order than $a_7$.

Summarizing all the results, we obtain (a).

Consider (b). The term $\frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)_{q} (\hat{\xi}_t)$ is equal to

$$\sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t \hat{x}_{jtq} \hat{\Sigma}_{jjc}^{-1} \hat{\lambda}_j \right)$$

which can be split into

$$\sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \left( \frac{1}{N} \sum_{j=1}^{N} \gamma_{jq} \hat{\Sigma}_{jjc}^{-1} \hat{\lambda}_j \right) + \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t \hat{x}_{jtq} \hat{\Sigma}_{jjc}^{-1} \hat{\lambda}_j \right)$$

The term $\frac{1}{N} \sum_{j=1}^{N} \gamma_{jq} \hat{\Sigma}_{jjc}^{-1} \hat{\lambda}_j$ is $v_q' + o_p(1)$ by (A.45). The second term is bounded in norm by

$$C \sum_{q=1}^{K} ||\hat{\beta}_q - \beta_q|| \left( \frac{1}{N} \sum_{j=1}^{N} ||\hat{\lambda}_j \hat{\Sigma}_{jjc}^{-1/2} \|_2^2 \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} f_t v_{jtq} \right\|^2 \right)^{1/2}$$

which is $o_p(||\hat{\beta} - \beta||)$. So we have

$$\frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{\xi}_t = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) v_q' + o_p(||\hat{\beta} - \beta||)$$

Notice $(I - A)' E \rightarrow I_r$ by Corollary A.1(a). Given this result, together with $H_N \rightarrow Q^{-1}$ by Corollary A.1(b), result (b) follows.
Consider (c). The term \( \sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t \bar{x}_{jtp} \hat{\gamma}_{jp}^{-1} \gamma_{jp}^{-1} \) can be split into

\[
\sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \frac{1}{N} \sum_{j=1}^{N} \gamma_{jq} \hat{\Sigma}_{joe}^{-1} \gamma_{jp}^{-1} + \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t v_{jtp} \hat{\gamma}_{jp}^{-1} \gamma_{jp}^{-1}
\]

The term \( \frac{1}{N} \sum_{j=1}^{N} \gamma_{jq} \hat{\Sigma}_{joe}^{-1} \gamma_{jp}^{-1} \) has already been proved to be \( \frac{1}{N} \sum_{j=1}^{N} \gamma_{jq} \hat{\Sigma}_{joe}^{-1} \gamma_{jp}^{-1} \) + \( o_p(1) = \omega_{pq} + o_p(1) \) in result (a). So the first term is \( \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}(\omega_{pq}) + o_p(1) \). The second term is bounded in norm by

\[
C \sum_{q=1}^{K} \| \hat{\beta}_q - \beta_q \| \left( \frac{1}{N} \sum_{j=1}^{N} \| \gamma_{jq} \Sigma_{joe}^{-1} \|^{2} \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \| \frac{1}{T} \sum_{t=1}^{T} f_t v_{jtp} \|^{2} \right)^{1/2}
\]

which is \( o_p(\| \hat{\beta} - \beta \|) \). Given these results, we have

\[
\sum_{p=1}^{K} (\hat{\beta}_p - \beta_p) \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t \bar{x}_{jtp} \hat{\gamma}_{jp}^{-1} \gamma_{jp}^{-1} = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}(\omega_{pq}) + o_p(\| \hat{\beta} - \beta \|)
\]

Notice \( (I - A)' P \rightarrow I_r \) by Corollary A.1(a), then (c) follows. \( \square \)

**Lemma A.11.** Under Assumptions A-D, we have

(a) \( J_{p1} = O_p(N^{-1/2}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \)

(b) \( J_{p2} = o_p(\| \hat{\beta} - \beta \|) \)

for all \( p \), where \( J_{p1} \) and \( J_{p2} \) are defined in (A.43) and (A.44), respectively.

**Proof of Lemma A.11.** Consider (a). By definition, \( J_{p1} \) is composed of 15 terms which we denote by \( b_1, b_2, \ldots, b_{15} \). We put aside the 2nd, 3rd and 10th terms temporarily.

Consider \( b_1 \). Notice that \( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{joe}^{-1} \lambda_j \gamma_{jp}^{-1} \) = \( O_p(1) \) by (A.45). So \( b_1 \) is \( O_p(N^{-1/2}T^{-1}) \) + \( o_p(\| \hat{\beta} - \beta \|) \) by Corollary A.2 and \( \hat{G} = O_p(N^{-1}) \). Consider \( b_4 \). The term \( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{joe}^{-1} \lambda_j \gamma_{jp}^{-1} f_t v_{jtp} \) is equal to

\[
\frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{joe}^{-1} (\lambda_j - \lambda_j) f_t v_{jtp} - \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{joe} - \Sigma_{joe} \lambda_j f_t v_{jtp} + \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{joe}^{-1} \lambda_j f_t v_{jtp}
\]
The first term of the above expression is bounded in norm by

\[ C\left( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{ij}^{-1} \| \hat{\lambda}_j - \lambda_j \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} f_{it} v_{jtp} \right\|^2 \right)^{1/2} \]

Note that \( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{ij}^{-1} \| \hat{\lambda}_j - \lambda_j \|^2 \) is bounded by \( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii}^{-1} \| \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 \).

So the first term is \( O_p(T^{-1}) + o_p(\| \hat{\beta} - \beta \|) \) by Proposition A.2. The second term is bounded in norm by

\[ C\left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ij} - \Sigma_{ij} \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} f_{it} v_{jtp} \right\|^2 \right)^{1/2} \]

Note that \( \frac{1}{N} \sum_{i=1}^{N} (\hat{\Sigma}_{ij} - \Sigma_{ij})^2 \) is bounded by \( \frac{1}{N} \sum_{i=1}^{N} \| \Sigma_{ii} - \Sigma_{ii} \|^2 \). So the second term is also \( O_p(T^{-1}) + o_p(\| \hat{\beta} - \beta \|) \) by Proposition A.2. The third term is \( O_p(N^{-1/2}T^{-1/2}) \). Given these results, we have \( \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{ij}^{-1} f_{it} f_{jtp} = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(\| \hat{\beta} - \beta \|) \) so \( b_4 \) is \( O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \) by Corollary A.2.

Using the result in \( b_4 \), we have \( b_5 = O_p(N^{-3/2}T^{-1/2}) + O_p(N^{-1}T^{-1}) + o_p(\| \hat{\beta} - \beta \|) \). Consider \( b_6 \). Term \( \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iit}^{-1} \xi_t f'_i \), which is equal to \( \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iit}^{-1} \xi_t f'_i \), can be proved to be \( O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(\| \hat{\beta} - \beta \|) \) similarly as \( \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{ij}^{-1} f_{it} f_{jtp} \). So \( b_6 \) is \( O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \) by Corollary A.2.

Since

\[ \left\| \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} (\hat{\lambda}_i - \lambda_i) \gamma_{ip} \hat{G} \right\| \leq \left( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \| \hat{\lambda}_i - \lambda_i \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \| \gamma_{ip} \|^2 \right)^{1/2} \| \hat{G} \|_N \]

we have \( \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} (\hat{\lambda}_i - \lambda_i) \gamma_{ip} \hat{G} = O_p(T^{-1/2}) + o_p(\| \hat{\beta} - \beta \|) \). Given this result, together with \( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \lambda_i \gamma_{ip} \), by the similar arguments, we can prove that \( b_8 = O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \), \( b_9 = O_p(N^{-3/2}T^{-1/2}) + O_p(N^{-1}T^{-1}) + o_p(\| \hat{\beta} - \beta \|) \) and \( b_11 = O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \).

Consider the term \( \frac{1}{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{jje}^{-1} \hat{f}_j \gamma_{jtp} \hat{e}_{jt} \), which can be proved to be \( O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(\| \hat{\beta} - \beta \|) \) similarly as the term \( \frac{1}{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{jje}^{-1} \hat{f}_j \gamma_{jtp} \hat{e}_{jt} \). Given this result, we have \( b_7 = O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \) by Corollary A.2.

Notice that \( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} (\hat{\gamma}_{ip} - \gamma_{ip}) \hat{\lambda}_i \) is bounded in norm by

\[ C\left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\gamma}_{ip} - \gamma_{ip} \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \| \hat{\lambda}_i \|^2 \right)^{1/2} \]
which is $O_p(T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|)$. Given this result, it follows that $b_{12}$ is $O_p(N^{-1}T^{-1/2}) + o_p(\|\hat{\beta} - \beta\|)$.

The term $\frac{1}{N} \sum_{i=1}^{N} \hat{\Gamma}_i \Sigma_{ii}^{-1} (\Sigma_{ii} - \Sigma_{ii}) \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i'$ is bounded in norm by

$$C \left( \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Gamma}_i\|^4 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \|\Sigma_{ii} - \Sigma_{ii}\|^2 \right)^{1/2}$$

Notice that $\frac{1}{N} \sum_{i=1}^{N} \|\hat{\Gamma}_i\|^4 \leq 2^4 \left( \frac{1}{N} \sum_{i=1}^{N} \|\Gamma_i\|^4 + \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Gamma}_i - \Gamma_i\|^4 \right)$. Using (A.29), it is easy to check that $\frac{1}{N} \sum_{i=1}^{N} \|\hat{\Gamma}_i - \Gamma_i\|^4 = o_p(1)$. So we have $\frac{1}{N} \sum_{i=1}^{N} \hat{\Gamma}_i \Sigma_{ii}^{-1} (\Sigma_{ii} - \Sigma_{ii}) \Sigma_{ii}^{-1} \hat{\Gamma}_i' = O_p(T^{-1/2}) + o_p(\|\hat{\beta} - \beta\|)$. From this, we have $b_{15} = O_p(N^{-1}T^{-1/2}) + o_p(\|\hat{\beta} - \beta\|)$. The terms $b_{13}$ and $b_{14}$ can be proved to be $O_p(N^{-1}T^{-1/2}) + o_p(\|\hat{\beta} - \beta\|)$ similarly as $b_{15}$.

We now consider the 2nd, 3rd and 10th terms. Using the first column of (A.29), the 3rd term can be verified to be $O_p(T^{-3/2}) + O_p(N^{-1/2}T^{-1}) + o_p(\|\hat{\beta} - \beta\|)$. For the 2nd and 10th terms, substituting the first column of (A.29) into the 2nd term, we obtain an expression which is composed of 10 terms. The first of which is canceled out with the 10th term. The remaining expression can be proved to be $O_p(T^{-3/2}) + O_p(N^{-1/2}T^{-1}) + o_p(\|\hat{\beta} - \beta\|)$. So we have that the 2nd, 3rd and 10th terms are $O_p(T^{-3/2}) + O_p(N^{-1/2}T^{-1}) + o_p(\|\hat{\beta} - \beta\|)$. This completes the proof of (a).

The proof of (b) is quite similar to the results in Lemmas A.6, A.8 and A.10 and hence omitted.

**Lemma A.12.** Under Assumptions A-D, we have

(a) $\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{tie}^{-1} e_{it} v_{itp} = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|)$

(b) $\text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \Sigma_{jj}^{-1} (\Sigma_{jj} - \Sigma_{jj}) \Sigma_{jj}^{-1} \hat{\Gamma}_i' \hat{G} \right]$

$= O_p(N^{-1}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|)$

(c) $\text{tr} \left[ \hat{H} \hat{\Sigma}_{ee}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_t e_t' - \Sigma_{ee}] \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \hat{H} \hat{v}_p \right]$

$= O_p(N^{-1}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|)$

(d) $\text{tr} \left[ \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \Sigma_{ii}^{-1} \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \Sigma_{jj}^{-1} [\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})] \gamma_{jp} \right]$

$= O_p(N^{-1}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|)$
Proof of Lemma A.12. Consider (a). The left hand side of (a) is equal to

\[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iie}^{-1} e_{it} v_{itp} - \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iie} - \Sigma_{iie} e_{it} v_{itp} \]

The first term is \( O_p(NT^{-1/2}) \) and the second term is bounded in norm by

\[ C \left( \frac{1}{N} \sum_{i=1}^{N} (\hat{\Sigma}_{iie} - \Sigma_{iie})^2 \right)^{1/2} \left( \frac{1}{T} \sum_{t=1}^{T} \left| e_{it} v_{itp} \right|^2 \right)^{1/2} \]

which is \( O_p(T^{-1}) + O_p(T^{-1/2} \| \hat{\beta} - \beta \|) \) by Proposition A.2. Given these results, (a) follows.

The remaining three results can be proved similarly as Lemma C.1(d) in [10] and hence omitted.

Lemma A.12 is used to derive Theorem 2.1. Given Theorem 2.1, we can strengthen the results in Lemma A.12, which are stated in the following lemma. These results are helpful to derive Theorem 2.2.

Lemma A.13. Under Assumptions A-D, we have

(a) \[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iie}^{-1} e_{it} v_{itp} \]

\[ = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iie} e_{it} v_{itp} + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \]

(b) \[ \text{tr} \left[ \hat{H} \hat{\Sigma}_{ee}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{ie} e_{it} - E(\varepsilon_{ie} e_{it})] \hat{\Sigma}_{iie}^{-1} \hat{H} \hat{\Sigma}_{ej} \right] \]

\[ = O_p(N^{-1/2}T^{-1}) + O_p(N^{-1}T^{-1/2}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \]

(c) \[ \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{H} \hat{\Sigma}_{ij}^{-1} e_{jt} e_{it} v_{itp} - E(\varepsilon_{it} e_{jt} e_{it}) \hat{\Sigma}_{iie}^{-1} \hat{\Sigma}_{jii} \hat{\Sigma}_{jep} \right] \]

\[ = O_p(N^{-1/2}T^{-1}) + O_p(N^{-1}T^{-1/2}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \]

(d) \[ \text{tr} \left[ \hat{H} \sum_{i=1}^{N} \hat{\Sigma}_{ii}^{-1} \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} [\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})] \hat{\Sigma}_{jej}^{-1} \hat{\Sigma}_{jep} \right] \]

\[ = O_p(N^{-1/2}T^{-1}) + O_p(N^{-1}T^{-1/2}) + O_p(T^{-3/2}) + o_p(\| \hat{\beta} - \beta \|) \]
Consider (a). Using (A.46), we show

\[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \frac{\hat{\Sigma}_{iie} - \Sigma_{iie}}{\Sigma_{iie}} e_{it} v_{itp} = O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\|\hat{\beta} - \beta\|)
\]

The left hand side is equal to

\[
- \frac{1}{NT^2} \sum_{i=1}^{N} \sum_{t=1}^{T} \sum_{s=1}^{T} \frac{1}{\Sigma_{iie}} \frac{1}{\Sigma_{iie}} (c_{iie}^2 - \Sigma_{iie}) e_{it} v_{itp} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \text{Res} \left\{ \frac{\hat{\Sigma}_{iie} - \Sigma_{iie}}{\Sigma_{iie}} \right\} e_{it} v_{itp}
\]

where \(\text{Res} \{\hat{\Sigma}_{iie} - \Sigma_{iie}\}\) is the right hand side expression of (A.34) excluding the first term. We separate the first term from the remaining expressions because the first term is \(O_p(T^{-1/2})\) and the remainings are all \(O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1})\). Consider the 2nd term of the above expression. By the boundedness of \(\hat{\Sigma}_{iie}\) and \(\Sigma_{iie}\), the 2nd term is bounded in norm by

\[
C \left( \frac{1}{N} \sum_{j=1}^{N} \|\text{Res} \{\hat{\Sigma}_{iie} - \Sigma_{iie}\}\|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left( \frac{1}{T} \sum_{t=1}^{T} e_{it} v_{itp} \right)^2 \right)^{1/2}
\]

Term \(\frac{1}{N} \sum_{j=1}^{N} \|\text{Res} \{\hat{\Sigma}_{iie} - \Sigma_{iie}\}\|^2\) is \(O_p(T^{-2}) + O_p(N^{-1}T^{-1})\) which can be verified term by term. Since this process is quite easy, we omit it. Thus the 2nd term is \(O_p(T^{-3/2}) + O_p(N^{-1/2}T^{-1})\) in view of \(\frac{1}{T} \sum_{t=1}^{T} e_{it} v_{itp} = O_p(T^{-1/2})\).

The 1st term can be rewritten as

\[
\frac{1}{NT^2} \sum_{i=1}^{N} \sum_{t=1}^{T} \sum_{s=1}^{T} \frac{1}{\Sigma_{iie}} (c_{iie}^2 - \Sigma_{iie}) e_{it} v_{itp} - \frac{1}{NT^2} \sum_{i=1}^{N} \sum_{t=1}^{T} \sum_{s=1}^{T} (\hat{\Sigma}_{iie} - \Sigma_{iie}) (c_{iie}^2 - \Sigma_{iie}) e_{it} v_{itp}
\]

The first expression is \(O_p(N^{-1/2}T^{-1})\) and the second is bounded in norm by

\[
C \left( \frac{1}{N} \sum_{i=1}^{N} (\hat{\Sigma}_{iie} - \Sigma_{iie})^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left( \frac{1}{T^2} \sum_{t=1}^{T} \sum_{s=1}^{T} (c_{iie}^2 - \Sigma_{iie}) e_{it} v_{itp} \right)^2 \right)^{1/2}
\]

The above expression is \(O_p(T^{-3/2})\). Thus the second expression is \(O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2})\). So (a) follows.

Consider (b). Notice \(\hat{v}_p = v_p + o_p(1)\) by (A.45) and \(\hat{H}_N = Q^{-1} + o_p(1)\), then it suffices to consider the term

\[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Sigma}_{iie}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon_{jt}' - E(\varepsilon_{it} \varepsilon_{jt}')] \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j
\]
which can be rewritten as

\[
\frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}_{ii}^{-1} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \hat{\Sigma}_{ij}^{-1} \hat{\Gamma}_j'
\]

\[
+ \frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i \hat{\Sigma}_{ii}^{-1} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \hat{\Sigma}_{ij}^{-1} (\hat{\Gamma}_j - \Gamma_j)'
\]

\[
+ \frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i (\hat{\Sigma}_{ii}^{-1} - \Sigma_{ii}^{-1}) \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \hat{\Sigma}_{ij}^{-1} \Gamma_j'
\]

\[
+ \frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i \Sigma_{ii}^{-1} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \Sigma_{ij}^{-1} \Gamma_j'
\]

(A.47)

\[
+ \frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i \Sigma_{ii}^{-1} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \Sigma_{ij}^{-1} \Gamma_j'
\]

The last term of (A.47) is \(O_p(N^{-1}T^{-1/2})\). Consider the fourth term of (A.47), which can be written as

\[
\frac{1}{N} \sum_{j=1}^{N} \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Gamma_i \Sigma_{ii}^{-1} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] (\hat{\Sigma}_{jj}^{-1} - \Sigma_{jj}^{-1}) \Gamma_j' \right)
\]

By the boundedness of \(\hat{\Sigma}_{jj}, \Sigma_{jj}, \Gamma_j\), the above is bounded in norm by

\[
C \left( \frac{1}{N} \sum_{j=1}^{N} \left\| \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Gamma_i \Sigma_{ii}^{-1} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \right\|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \left\| \hat{\Sigma}_{jj} - \Sigma_{jj} \right\|^2 \right)^{1/2}
\]

which is \(O_p(N^{-1/2}T^{-1})\) since \(\frac{1}{N} \sum_{j=1}^{N} \left\| \hat{\Sigma}_{jj} - \Sigma_{jj} \right\|^2 = O_p(T^{-1})\) by Theorem 2.1.

Consider the third term of (A.47), which is equal to

\[
\frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i (\hat{\Sigma}_{ii}^{-1} - \Sigma_{ii}^{-1}) \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \Sigma_{ij}^{-1} \Gamma_j'
\]

\[
+ \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i (\hat{\Sigma}_{ii}^{-1} - \Sigma_{ii}^{-1}) \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] (\hat{\Sigma}_{jj}^{-1} - \Sigma_{jj}) \Gamma_j'
\]
The first expression is the same as the 4th term and hence $O_p(N^{-1/2}T^{-1})$. The second expression is bounded in norm by

$$
\left( \frac{1}{N} \sum_{i=1}^{N} \| \Gamma_i \|^2 \cdot \| \hat{\Sigma}_{ii}^{-1} - \Sigma_{ii}^{-1} \|^2 \right) \left( \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \left\| \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \right\|^2 \right)^{1/2}
$$

Term $\frac{1}{N} \sum_{i=1}^{N} \| \Gamma_i \|^2 \cdot \| \hat{\Sigma}_{ii}^{-1} - \Sigma_{ii}^{-1} \|^2$ is bounded in norm by $C \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 = O_p(T^{-1})$ by the boundedness of $\Sigma_{ii}, \hat{\Sigma}_{ii}$ and $\Gamma_i$. So the second expression is $O_p(T^{-3/2})$. Thus the third term of (A.47) is $O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2})$.

Consider the second term of (A.47), which is equal to

$$
\frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} \Gamma_i \hat{\Sigma}_{ii}^{-1} \sum_{t=1}^{T} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \hat{\Sigma}_{ij}^{-1} (\hat{\Gamma}_j - \Gamma_j)(\hat{\Gamma}_j - \Gamma_j)'
$$

By the boundedness of $\hat{\Sigma}_{ij}, \Gamma_i$, the second expression is bounded in norm by

$$
C \left( \frac{1}{N} \sum_{i=1}^{N} \| \Sigma_{ii} - \hat{\Sigma}_{ii} \|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1} \cdot \| \hat{\Gamma}_j - \Gamma_j \|^2 \right)^{1/2}
$$

Terms $\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2$ and $\frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1} \| \| \hat{\Gamma}_j - \Gamma_j \|^2$ are both $O_p(T^{-1})$ by Theorem 2.1. So the second expression is $O_p(T^{-3/2})$. The first expression can be written as

$$
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Gamma_i \hat{\Sigma}_{ii}^{-1} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \hat{\Sigma}_{ij}^{-1} (\hat{\Gamma}_j - \Gamma_j)(\hat{\Gamma}_j - \Gamma_j)'
$$

which is bounded in norm by

$$
C \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Gamma_i \hat{\Sigma}_{ii}^{-1} [\varepsilon_{it} \varepsilon'_{jt} - E(\varepsilon_{it} \varepsilon'_{jt})] \right\|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1} \| \| \hat{\Gamma}_j - \Gamma_j \|^2 \right)^{1/2}
$$

which is $O_p(N^{-1/2}T^{-1})$ since $\frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1} \| \| \hat{\Gamma}_j - \Gamma_j \|^2 = O_p(T^{-1})$ by Theorem 2.1. Given these two results, we have that the second term of (A.47) is $O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2})$. 


Now consider the 1st term of (A.47), which is equal to
\[
\frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}^{-1}_{ii} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it}\varepsilon_{jt}' - E(\varepsilon_{it}\varepsilon_{jt}')] \hat{\Sigma}^{-1}_{jj}' (\hat{\Gamma}_j - \Gamma_j)'
\]
\[
\frac{1}{N^2T} \sum_{i=1}^{N} \sum_{j=1}^{N} (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}^{-1}_{ii} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it}\varepsilon_{jt}' - E(\varepsilon_{it}\varepsilon_{jt}')] \hat{\Sigma}^{-1}_{jj}'
\]

The second expression is the same as the 2nd term of (A.47) and hence \(O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2})\). The first expression is bounded in norm by
\[
\left( \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Gamma}_i - \Gamma_i\| \cdot \|\hat{\Sigma}^{-1}_{ii}\| \right) \left( \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \frac{1}{T} \sum_{t=1}^{T} \|\varepsilon_{it}\varepsilon_{jt}' - E(\varepsilon_{it}\varepsilon_{jt}')\|^2 \right)^{1/2}
\]
which is \(O_p(T^{-3/2})\). Thus the 1st term of (A.47) is \(O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2})\).

Summarizing all the results, we obtain (b).

Notice \(\frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2 \leq \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2 = O_p(T^{-1})\) and \(\frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{ii} - \Sigma_{ii}\| \lambda_i - \lambda_i\|^2 \leq \frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{ii}^{-1}\| \cdot \|\hat{\Gamma}_i - \Gamma_i\|^2 = O_p(T^{-1})\). So results (c) and (d) can be proved in the same way as result (b). This completes the proof of Lemma A.13.

\(\square\)

PROOFS OF THEOREM 2.2. Consider (A.42). Terms on the left hand side are summarized in Lemma A.10. The first three and the fifth terms on the right hand side are shown in Lemma A.13. The fourth and sixth are also covered by Lemma A.10. Terms \(J_{p1}, J_{p2}\) are given in Lemma A.11. Given these results, we have
(A.48)
\[
\Omega(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iet} v_{itx} + O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1}T^{-1})
\]
This proves Theorem 2.2. \(\square\)

APPENDIX A.3: PROOF OF THE ALTERNATIVE EXPRESSION OF THEOREM 2.2

In this section, we show that the asymptotic expression of (2.9) is equal to the one in Theorem 2.2. We first introduce some notations. Let
\[
\bar{\Gamma}_p = \begin{pmatrix} \mu_{1p} & \gamma_{1p}' \\ \mu_{2p} & \gamma_{2p}' \\ \vdots & \vdots \\ \mu_{Np} & \gamma_{Np}' \end{pmatrix}, \quad \bar{F} = \begin{pmatrix} f_1' \\ f_2' \\ \vdots \\ f_T' \end{pmatrix}, \quad \bar{V}_p = \begin{pmatrix} v_{11p} & v_{12p} & \cdots & v_{1Tp} \\ v_{21p} & v_{22p} & \cdots & v_{2Tp} \\ \vdots & \vdots & \ddots & \vdots \\ v_{N1p} & v_{N2p} & \cdots & v_{NTp} \end{pmatrix}
\]
Then, by the second equation of (2.1), the \( p \)-th regressor can be written as:

\[
(A.49) \quad X_p = \bar{F}_p \bar{F}' + V_p, \quad (p = 1, 2, \ldots, K)
\]

Now consider term \( \frac{1}{NT} \text{tr}[\bar{M}X_p \bar{M}(\bar{F})X'_q] \), where \( p, q = 1, 2, \ldots, K \). By (A.49), this term is equal to

\[
\frac{1}{NT} \text{tr}[\bar{M}X_p \bar{M}(\bar{F})X'_q] = \frac{1}{NT} \text{tr}[\bar{M}V_p \bar{M}(\bar{F})V'_q] = \frac{1}{NT} \text{tr}[\bar{V}'q \Sigma^{-1} \bar{V}_p]
\]

\[
- \frac{1}{NT} \text{tr}[\Lambda' \Sigma^{-1} \bar{V}_p V'_q \Sigma^{-1} \Lambda (\Lambda' \Sigma^{-1} \Lambda)^{-1}] - \frac{1}{NT} \text{tr}[\bar{F}' \Sigma^{-1} \bar{V}_p \bar{F}(\bar{F}' \bar{F})^{-1}]
\]

\[
(A.50)
\]

The first term on the right hand side of (A.50) is equal to \( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} v_{ipt} v_{itq} \), which is equal to

\[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} [v_{ipt} v_{itq} - \Sigma^{(p,q)}_{iix}] + \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iiie} \Sigma^{(p,q)}_{iix} = \frac{1}{N} \sum_{i=1}^{N} \Sigma_{iie} \Sigma^{(p,q)}_{iix} + o_p(1)
\]

because \( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} [v_{ipt} v_{itq} - \Sigma^{(p,q)}_{iix}] = O_p(N^{-1/2}T^{-1/2}) \).

Consider the 2nd term of (A.50), which is equal to \( \text{tr}[\frac{1}{NT} \Lambda' \Sigma^{-1} \bar{V}_p V'_q \Sigma^{-1} \Lambda (\Lambda' \Sigma^{-1} \Lambda)^{-1}] \). The term inside the trace operator is equal to

\[
\left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \frac{1}{\Sigma_{iie} \Sigma_{jj}} \lambda_i \lambda_j \sum_{t=1}^{T} v_{ipt} v_{jtq} \right) \left( \frac{1}{N} \Lambda' \Sigma^{-1} \Lambda \right)^{-1}
\]

which can be rewritten as

\[
\left( \frac{1}{N^2 T} \sum_{i=1}^{N} \sum_{j=1}^{N} \frac{1}{\Sigma_{iie} \Sigma_{jj}} \lambda_i \lambda_j \sum_{t=1}^{T} [v_{ipt} v_{jtq} - E(v_{ipt} v_{jtq})] \right) \left( \frac{1}{N} \Lambda' \Sigma^{-1} \Lambda \right)^{-1}
\]

\[
+ \left( \frac{1}{N^2} \sum_{i=1}^{N} \frac{1}{\Sigma_{iie}^{(p,q)}} \right) \left( \frac{1}{N} \Lambda' \Sigma^{-1} \Lambda \right)^{-1}
\]

The first expression is \( O_p(N^{-1/2}T^{-1/2}) \) and the second is \( O_p(N^{-1}) \). So the 2nd term of (A.50) is \( O_p(N^{-1/2}T^{-1/2}) + O_p(N^{-1}) \).

Consider the 3rd term of (A.50), which is \( \text{tr}[\frac{1}{NT} \bar{F}' \Sigma^{-1} \bar{V}_p \bar{F}(\bar{F}' \bar{F})^{-1}] \). The term inside the trace operator is equal to

\[
\left( \frac{1}{NT^2} \sum_{i=1}^{N} \sum_{t=1}^{T} \sum_{s=1}^{T} \sum_{l=1}^{T} \sum_{j=1}^{T} f_{is} f_{lt} v_{ipt} v_{isq} \right) \left( \frac{1}{T} \bar{F}' \bar{F} \right)^{-1}
\]
where $\tilde{f}_t = (1, f'_t)'$. The above expression can be rewritten as

$$
\left( \frac{1}{NT^2} \sum_{i=1}^{N} \sum_{t=1}^{T} \sum_{s=1}^{T} \tilde{f}_s \tilde{f}_t [v_{ipt}v_{isq} - E(v_{ipt}v_{isq})] \right) \left( \frac{1}{T} \mathbf{F}' \mathbf{F} \right)^{-1} + \frac{1}{NT} \sum_{i=1}^{N} \Sigma^{(p,q)}_{iix}
$$

The first expression is $O_p(N^{-1/2}T^{-1})$ and the second is $O_p(T^{-1})$. So the third term of (A.50) is $O_p(N^{-1/2}T^{-1}) + O_p(T^{-1})$.

Consider the last term of (A.50). Notice

$$
1 \frac{1}{NT} \Lambda^c \Sigma^{-1} v_{p} \mathbf{F}
$$

is equal to

$$
1 \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \frac{1}{\Sigma_{iix}} \lambda_i \tilde{f}_t v_{ipt}
$$

which is $O_p(N^{-1/2}T^{-1/2})$. Thus the last term of (A.50) is $O_p(N^{-1}T^{-1})$.

Given these results, we have

$$
\frac{1}{NT} \text{tr} \left[ \tilde{M} X_p \mathcal{M}(\mathbf{F}) \right] = \frac{1}{N} \sum_{i=1}^{N} \Sigma_{iix}^{-1} \Sigma_{iix} + o_p(1)
$$

Then it follows

$$
\frac{1}{NT} \left( \begin{array}{ccc}
\text{tr} \left[ \tilde{M} X_1 \mathcal{M}(\mathbf{F}) X'_1 \right] & \cdots & \text{tr} \left[ \tilde{M} X_1 \mathcal{M}(\mathbf{F}) X'_K \right] \\
\vdots & \vdots & \vdots \\
\text{tr} \left[ \tilde{M} X_K \mathcal{M}(\mathbf{F}) X'_1 \right] & \cdots & \text{tr} \left[ \tilde{M} X_K \mathcal{M}(\mathbf{F}) X'_K \right]
\end{array} \right)
= \frac{1}{N} \sum_{i=1}^{N} \Sigma_{iix}^{-1} \Sigma_{iix} + o_p(1)
$$

(A.51)

Now we turn attention to $\frac{1}{NT} \text{tr} \left[ \tilde{M} X_p \mathcal{M}(\mathbf{F}) e' \right]$, where $p = 1, 2, \cdots K$. By (A.49), we have

$$
\frac{1}{NT} \text{tr} \left[ \tilde{M} X_p \mathcal{M}(\mathbf{F}) e' \right] = \frac{1}{NT} \text{tr} \left[ \tilde{M} \Sigma_{ee}^{-1} \mathcal{V} e' \right] = \frac{1}{NT} \text{tr} \left[ e' \Sigma_{ee}^{-1} \mathcal{V} \right]
$$

$$
- \frac{1}{NT} \text{tr} \left[ \Lambda \Sigma_{ee}^{-1} \mathcal{V} \mathcal{V} e' \Lambda (\Lambda \Sigma_{ee}^{-1} \Lambda)^{-1} \right] - \frac{1}{NT} \text{tr} \left[ \Lambda \Sigma_{ee}^{-1} \mathcal{V} \mathcal{F} \mathcal{F} \mathcal{F}^{-1} \right]
$$

$$
+ \frac{1}{NT} \text{tr} \left[ \mathcal{F} e' \Sigma_{ee}^{-1} \Lambda (\Lambda \Sigma_{ee}^{-1} \Lambda)^{-1} \Lambda \Sigma_{ee}^{-1} \mathcal{V} \mathcal{F} \mathcal{F} \mathcal{F}^{-1} \right]
$$

The first term on the right hand side is equal to $\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \frac{1}{\Sigma_{iix}} v_{ipt} e_{it}$, which is $O_p(N^{-1/2}T^{-1/2})$. The second term can be proved to be $O_p(N^{-1}T^{-1/2})$, similarly as the 2nd term of (A.50). The third term can be proved to be
\( O_p(N^{-1/2}T^{-1}) \) similarly as the 3rd term of (A.50). The last term can be proved to be \( O_p(N^{-1}T^{-1}) \) similarly as the 4th term of (A.50). So we have

\[
\frac{1}{NT} \text{tr}[\tilde{M}X_pM(F)e'] = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \frac{1}{\Sigma_{ite}} v_{ipt} e_{it} + O_p(N^{-1/2}T^{-1}) + O_p(N^{-1}T^{-1/2})
\]

Then it follows

\[
\begin{pmatrix}
\text{tr}[\tilde{M}X_1M(F)e'] \\
\vdots \\
\text{tr}[\tilde{M}X_KM(F)e']
\end{pmatrix} = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \frac{1}{\Sigma_{ite}} v_{ipt} e_{it} + O_p(N^{-1/2}T^{-1}) + O_p(N^{-1}T^{-1/2})
\]

Combining (A.51) and (A.52), we obtain the same asymptotic expression as in Theorem 2.2. Taking limit on both sides of (A.51), we obtain Corollary 2.1.

It is interesting to note that the limiting distribution in Corollary 2.1 can be obtained by a generalized principal components method. Rewrite the model as

\[
Y_t = \mathbf{X}_t \beta + \Lambda f_t + \epsilon_t,
\]

where \( Y_t = (y_{1t}, y_{2t}, \cdots, y_{Nt})' \), \( \mathbf{X}_t \) is an \( N \times K \) matrix with its \((i,k)\) element \( x_{ikt} \). Recall that \( \Sigma_{ee} \) is known. Consider the following objective function

\[
SSR(\beta, \Lambda, F) = \sum_{t=1}^{T} (Y_t - \mathbf{X}_t \beta - \Lambda f_t)' \Sigma_{ee}^{-1} (Y_t - \mathbf{X}_t \beta - \Lambda f_t)
\]

where \( X \) is a three dimensional data matrix such that \( X\beta = X_1\beta_1 + \cdots + X_K\beta_K \) (with \( X_k \) being \( N \times T \) for \( k \leq K \)); \( Y^\dagger = \Sigma_{ee}^{-1/2} Y \), and \( X^\dagger \beta = X_1^\dagger \beta_1 + \cdots + X_K^\dagger \beta_K \) with \( X_k^\dagger = \Sigma_{ee}^{-1/2} X_k \); \( \Lambda^\dagger = \Sigma_{ee}^{-1/2} \Lambda \). If \( \alpha \) is also present, we use \( \tilde{Y}_t \) and \( \tilde{X}_t \), etc. This objective function is similar to that of [8], which uses \( \Sigma_{ee} = I_N \). If \( \Sigma_{ee} \) is known, we can treat \( Y^\dagger \) and \( X^\dagger \) as the data and use the same estimation method as in [8]. Using the asymptotic representation in [8], it is not difficult to verify that the estimator obtained by minimizing the above objective function has the same asymptotic representation as in Corollary 2.1. Because \( \Sigma_{ee} \) is unknown, this leads naturally to an iterated two-step procedure. The first step estimates \((\beta, \Lambda, F)\) the same way as in [8]. The second step constructs an estimate of \( \Sigma_{ee} \) based on the residuals and then reestimate \((\beta, \Lambda, F)\); these two steps are iterated for a number of times. Despite the iteration, this is a two-step procedure (or a sequential procedure). In contrast, the maximum likelihood procedure is a joint procedure (simultaneous maximization). Also, the maximum likelihood procedure has much better finite sample properties.
In addition to Table 1, we define the following notations, which will be used in the subsequent proof.

The symbols used in the Appendix B

\[
\begin{align*}
\psi_{ip}^g &= N^{-1} \sum_{i=1}^{N} \lambda_i \Sigma_{iie}^{-1} \gamma_{ip}^g \\
\omega_{ip}^g &= N^{-1} \sum_{i=1}^{N} \psi_i \Sigma_{iie}^{-1} h_{ip} \\
\Pi_{\lambda\psi} &= N^{-1} \sum_{i=1}^{N} \lambda_i \Sigma_{iie}^{-1} \psi_i \\
\Pi_{\psi\psi} &= N^{-1} \sum_{i=1}^{N} \psi_i \Sigma_{iie}^{-1} \psi_i \\
F_t &= N^{-1} \sum_{i=1}^{N} \hat{z}_{it} \Sigma_{iie}^{-1} \lambda_i \\
\mathcal{H}_t &= N^{-1} \sum_{i=1}^{N} \hat{z}_{it} \Sigma_{iie}^{-1} \psi_i \\
\mathcal{G}_{tp} &= N^{-1} \sum_{i=1}^{N} \hat{z}_{it} \Sigma_{iie}^{-1} \gamma_{ip} \\
\omega_{pq}^{gg} &= N^{-1} \sum_{i=1}^{N} \gamma_{ip} \Sigma_{iie}^{-1} \gamma_{iq} \\
\omega_{pq}^{hh} &= N^{-1} \sum_{i=1}^{N} \gamma_{ip} \Sigma_{iie}^{-1} h_{ip} \\
G_{1N} &= N \cdot G_1 \\
G_{2N} &= N \cdot G_2
\end{align*}
\]

where $G_1$ and $G_2$ are defined in the main context. Note that not all variables in the right column have “hat”.

We then derive (3.9) in the main text:

\[
\frac{1}{N} \hat{G}_2 \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma} \hat{\Sigma}_{i}^{-1} \left( \hat{B} M_{zz}^g \hat{B}' - \hat{\Sigma}_{zz}^g \right) I_{K+1}^1 \hat{\psi}_j' = 0
\]

The first order condition on $\Gamma$ gives

\[
\hat{G} \hat{\Sigma}_{zz}^{-1} (I_N \otimes \hat{B}) M_{zz} (I_N \otimes \hat{B}') - \hat{\Sigma}_{zz} = W'
\]

Notice $\hat{G} \hat{\Sigma}_{zz}^{-1} = \hat{G} \hat{\Sigma}_{zz}^{-1}$ and $\hat{\Sigma}_{zz}^{-1} = \hat{\Sigma}_{zz}^{-1} - \hat{\Sigma}_{zz}^{-1} \hat{G} \hat{\Sigma}_{zz}^{-1}$, we have

\[
\hat{G} \hat{\Sigma}_{zz}^{-1} (I_N \otimes \hat{B}) M_{zz} (I_N \otimes \hat{B}') - \hat{\Sigma}_{zz} - \hat{\Sigma}_{zz}^{-1} \hat{G} \hat{\Sigma}_{zz}^{-1} = W'
\]

However, $\hat{G} \hat{\Sigma}_{zz}^{-1} (I_N \otimes \hat{B}) M_{zz} (I_N \otimes \hat{B}') - \hat{\Sigma}_{zz} \hat{\Sigma}_{zz}^{-1} \hat{G} = 0$ by (3.4). So the above equation can be simplified as

\[
\hat{G} \hat{\Sigma}_{zz}^{-1} (I_N \otimes \hat{B}) M_{zz} (I_N \otimes \hat{B}') - \hat{\Sigma}_{zz} \hat{\Sigma}_{zz}^{-1} = W'
\]

By $\frac{1}{N} W' = 0$, we have

\[
\frac{1}{N} \hat{G}_2 \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma} \hat{\Sigma}_{i}^{-1} \left( \hat{B} M_{zz}^g \hat{B}' - \hat{\Sigma}_{zz}^g \right) I_{K+1}^1 \hat{\psi}_j' = 0
\]
APPENDIX B1: PROOF OF CONSISTENCY

Again, for consistency, we use the superscript “*” to denote the true parameters.

**Proposition B.1.** Let \( \hat{\theta} = (\hat{\beta}, \hat{\Gamma}, \hat{\Sigma}_{ee}) \) be the solution by maximizing (3.2). Under Assumptions A-E, together with the identification conditions \( IZ \), when \( N, T \to \infty \), we have

\[
\hat{\beta} - \beta \overset{p}{\to} 0
\]

\[
\frac{1}{N} \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i)' \overset{p}{\to} 0
\]

\[
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 \overset{p}{\to} 0
\]

To prove Proposition B.1, we need the following lemmas.

**Lemma B.1.** Under Assumptions A-E,

(a) \( \hat{G} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i' \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t' e_{jt} \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j = O_p(T^{-1/2}) \)

(b) \( \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_t' \right) \hat{\Pi}_{\lambda\psi} = O_p(T^{-1/2}) \)

(c) \( \hat{G} \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \sum_{t=1}^{T} [e_{jt} e_{jt} - E(e_{jt} e_{jt})] \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j = O_p(T^{-1/2}) \)

(d) \( \hat{G} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i' \left( \frac{1}{T} \sum_{t=1}^{T} f_t' (\hat{\beta} - \beta^*)' \hat{H}_t \right) = O_p(\| \hat{\beta} - \beta^* \|) \)

(e) \( \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta^*)' \hat{H}_t \right) = o_p(\| \hat{\beta} - \beta^* \|) \)

(f) \( \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^*) f_t' \right) \hat{\Pi}_{\lambda\psi} = O_p(\| \hat{\beta} - \beta^* \|) \)

(g) \( \hat{G}_N \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^*) e_{jt} \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j \right) = o_p(\| \hat{\beta} - \beta^* \|) \)

(h) \( \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^*) (\hat{\beta} - \beta^*)' \hat{H}_t \right) = O_p(\| \hat{\beta} - \beta^* \|^2) \)
\begin{equation}
(i) \quad \frac{1}{N} \hat{\sum} \sum_{j=1}^{N} \hat{\lambda}_j \hat{\sum}^{-1}_{j\epsilon} \hat{\sum}^{-1}_{j\epsilon} \hat{\psi}_j = o_p \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\sum}_{ii} - \sum_{ii}^\star \| \right)
\end{equation}

PROOF OF LEMMA B.1. Consider (a). The left hand side of (a) is bounded, by the Cauchy-Schwarz inequality, in norm by

\begin{equation}
C \| \hat{H}^{1/2} \| \cdot \left( \sum_{i=1}^{N} \left\| \hat{H}^{1/2} \hat{\Gamma}_i \hat{\sum}^{-1/2}_{ii} \right\|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \left\| \hat{\Gamma}_i \right\|^2 \right)^{1/2} \cdot \left( \sum_{j=1}^{N} \left\| \hat{\sum}^{-1/2}_{j\epsilon} \hat{\psi}_j \right\|^2 \right)^{1/2} \left( \frac{1}{T} \sum_{t=1}^{T} \| \hat{f}_t e_j \| \right)^{1/2}
\end{equation}

Note that \( \sum_{i=1}^{N} \| \hat{H}^{1/2} \hat{\Gamma}_i \hat{\sum}^{-1/2}_{ii} \|^2 = r \) by (A.1), \( \frac{1}{N} \sum_{i=1}^{N} \left\| \hat{\sum}^{-1/2}_{ii} \hat{\psi}_j \right\|^2 = O_p(1) \) by the boundedness of \( \hat{\sum}_{ii} \). \( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Gamma}_i \|^2 = O_p(T^{-1}) \) and \( \| (\hat{H} + I)^{-1} \| \leq \sqrt{r} \). Further notice that \( \| \hat{H}^{1/2} \| (\sum_{j=1}^{N} \left\| \hat{\sum}^{-1/2}_{j\epsilon} \hat{\psi}_j \right\|^2)^{1/2} \) is equal to

\begin{equation}
\left[ \text{tr}(\hat{H}) \text{tr} \left( \sum_{i=1}^{N} \hat{\psi}_i \hat{\sum}^{-1}_{j\epsilon} \hat{\psi}_j \right) \right]^{1/2} \leq 1
\end{equation}

by the definition of \( \hat{H} \). Given these results, we have (a).

Using the arguments in proving Lemmas A.5 and A.6, we can prove the remaining results similarly as (a). To save space, the detailed proofs are omitted.

We point out that Lemmas A.3-A.6 continue to hold in the present context because the proofs don’t involve the special structure of the factor loadings and the identification conditions.

PROOF FOR PROPOSITION B.1. The proof of Proposition 2.1, from the beginning to equation (A.24) except equations (A.28), continues to hold in the present context because this part of proof doesn’t involve the prior restrictions as well as the identification conditions. In addition, (2.7) holds for the same reason. The following proof are based on these results.

By (A.24), together with \( \hat{M}_{ff} = M_{ff} = I_T \), we have

\begin{equation}
I = (I - A)'(I - A) + o_p(1)
\end{equation}

The above result, in combination with (A.16) indicates that

\begin{equation}
\| N^{1/2} \hat{H}^{1/2} \| = O_p(1)
\end{equation}
Noticing \( \frac{1}{T} \hat{\Gamma} \hat{\Sigma}_{\varepsilon \varepsilon}^{-1} \hat{\Gamma} = \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{\varepsilon i e}^{-1} \hat{\lambda}_i \hat{\lambda}_i' + \frac{1}{T} \sum_{t=1}^{T} \hat{\varepsilon}_{it} \hat{\Sigma}_{\varepsilon i e}^{-1} \hat{\varepsilon}_{it}' \), we have \( \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{\varepsilon i e}^{-1} \hat{\lambda}_i \hat{\lambda}_i' = O_p(1) \), which implies

\[
(\text{B.3}) \quad \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{\varepsilon i e}^{-1} \hat{\psi}_i \hat{\psi}_i' = O_p(1)
\]

Now consider the first order condition (3.5). Noticing \( \hat{H} = \hat{G} + \hat{H} \hat{G} \), by (A.20), equation (3.5) is equivalent to

\[
(\text{B.4}) \quad \tilde{\psi}_j - \psi_j^* = -\hat{H}_1 N \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} (\hat{\Gamma}_i - \Gamma_i^*) \lambda_j + \hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} \Gamma_i^{**} \frac{1}{T} \sum_{t=1}^{T} f_{jt}^* e_{jt}
\]

\[
+ \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_{jt}^{**} \right) \lambda_j^* + \hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})]
\]

\[
- \hat{G}_1 N \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} \Gamma_i^{**} \frac{1}{T} \sum_{t=1}^{T} f_{jt}^* \hat{\lambda}_j - \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t \hat{x}_{jt} \right) (\beta - \hat{\beta})
\]

\[
- \hat{G}_1 N \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} \Gamma_i^{**} \frac{1}{T} \sum_{t=1}^{T} \hat{i}_{it} (\beta - \beta^*) e_{jt}
\]

\[
+ \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta^*) \hat{x}_{jt} \right) - \hat{G}_1 \hat{\lambda}_j \hat{\Sigma}_{\varepsilon j}^{-1} (\hat{\Sigma}_{\varepsilon j} - \Sigma_j^{**})
\]

\[
+ \hat{H}_1 \hat{G} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} (\hat{\Gamma}_i - \Gamma_i^* \lambda_j^* + \hat{H}_1 \hat{G} \hat{H}^{-1} (\lambda_j - \hat{\lambda}_j)
\]

where we neglect the smaller order term \( \hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} \hat{\varepsilon}_i \hat{\varepsilon}_j \). Furthermore, equation (3.9) can be written as

\[
(\text{B.5}) \quad \hat{H}_2 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} (\hat{\Gamma}_i - \Gamma_i^* \lambda_j \hat{\psi}_j = \hat{G}_2 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_{jt}^{**} \right) \hat{\Pi}_{\lambda \psi}
\]

\[
+ \frac{1}{NT} \hat{G}_2 \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})] \hat{\Sigma}_{\varepsilon j e}^{-1} \hat{\psi}_j'
\]

\[
+ \hat{G}_2 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\varepsilon i}^{-1} \Gamma_i^{**} \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_{jt}^* e_{jt} \hat{\Sigma}_{\varepsilon j e}^{-1} \hat{\psi}_j' - \hat{G}_2 N \left( \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta^*) \hat{\lambda}_j \right)
\]

\[
- \hat{G}_2 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta^*) f_{jt}^{**} \right) \hat{\Pi}_{\lambda \psi} + \hat{H}_2 \hat{G} \hat{H}^{-1} (\lambda_j - \hat{\lambda}_j) \hat{\Sigma}_{\varepsilon j e}^{-1} \hat{\psi}_j'
\]
\[ -G_{2N} \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \xi_t(\beta - \beta^*) e_{jt} \hat{\Sigma}^{-1}_{jej} \hat{\psi}'_j \right) + \hat{H}_2 \hat{G} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}_i - \Gamma_i) \hat{\Pi}_{\lambda} \]

\[ + \hat{G}_{2N} \left( \frac{1}{T} \sum_{t=1}^{T} \xi_t(\beta - \beta^*) (\beta - \beta^*)' \hat{H}_t \right) - \frac{1}{N} \hat{G}_2 \sum_{j=1}^{N} \lambda_j \hat{\Sigma}_{jej}^{-1} (\hat{\Sigma}_{jej} - \Sigma_{jej}) \hat{\Sigma}_{jej}^{-1} \hat{\psi}'_j \]

\[ - \hat{G}_2 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} \hat{\Gamma}_i \frac{1}{T} \sum_{t=1}^{T} f_t^* (\beta - \beta^*)' \hat{H}_t \]

where \( \hat{\Pi}_{\lambda} = \frac{1}{N} \sum_{j=1}^{N} \lambda_j^* \hat{\Sigma}^{-1}_{jej} \hat{\psi}'_j \) and \( \hat{H}_t = \frac{1}{N} \sum_{j=1}^{N} f_t^* \hat{\Sigma}^{-1}_{jej} \hat{\psi}'_j \).

We continue to use notation \( A = (\hat{\Gamma} - \Gamma)' \Sigma_{jej}^{-1} \hat{\Gamma} \hat{H} \). Except the 6th and 8th terms on the right hand side of (B.5), all the remaining are given in Lemma B.1. However, the 6th term is of smaller order than the left hand side and hence negligible. For the 8th term, it is bounded in norm by

\[ CN^{-1/2} \| \hat{H}^{-1/2} \| \left( \sum_{i=1}^{N} \left\| \hat{\Sigma}_{jej}^{-1/2} \right\|^2 \right)^{1/2} \left( \frac{1}{N} \sum_{i=1}^{N} \| \lambda_i^* \|^2 \right)^{1/2} \| \hat{G} \hat{H}^{-1} \| \]

which is \( o_p(1) \) by \( \hat{H} = o_p(1) \). So we can write equation (B.5) alternatively, in terms of \( A \), as

\[ A_{12}' \frac{1}{N} \sum_{j=1}^{N} \psi_j^* \hat{\Sigma}^{-1}_{jej} \hat{\psi}'_j = o_p(1) \]

where \( A_{12}' = \hat{H}_2 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}^y - \Gamma^y)' \) is the (1, 2)th submatrix of \( A \).

Consider (B.4). First post-multiplying \( \hat{\Sigma}^{-1}_{jej} \hat{\psi}'_j \), then taking summation over index \( j \) from 1 to \( N \), then dividing both sides by \( N \), and by the similar arguments in proving Lemma B.1, we have

\[ \frac{1}{N} \sum_{j=1}^{N} (\hat{\psi}_j - \psi_j^*) \hat{\Sigma}_{jej}^{-1} \psi_{j'} = -A_{11} \frac{1}{N} \sum_{j=1}^{N} \psi_j^* \hat{\Sigma}_{jej}^{-1} \psi_{j'} + o_p(1) \]

which can be written alternatively as

\[ \frac{1}{N} \sum_{j=1}^{N} \hat{\psi}_j \hat{\Sigma}_{jej}^{-1} \psi_{j'} = (I - A_{11}) \frac{1}{N} \sum_{j=1}^{N} \psi_j^* \hat{\Sigma}_{jej}^{-1} \psi_{j'} + o_p(1) \]

Similarly, post-multiplying (B.4) by \( \hat{\Sigma}_{jej}^{-1} \hat{\psi}'_j \) instead of \( \hat{\Sigma}_{jej}^{-1} \psi_{j'} \), we have

\[ \frac{1}{N} \sum_{j=1}^{N} \hat{\psi}_j \hat{\Sigma}_{jej}^{-1} \hat{\psi}'_j = (I - A_{11}) \frac{1}{N} \sum_{j=1}^{N} \psi_j^* \hat{\Sigma}_{jej}^{-1} \hat{\psi}'_j + o_p(1) \]

Combining the preceding two equations, we have

\[ \frac{1}{N} \sum_{j=1}^{N} \hat{\psi}_j \hat{\Sigma}_{jej}^{-1} \psi_{j'} = (I - A_{11}) \frac{1}{N} \sum_{j=1}^{N} \psi_j^* \hat{\Sigma}_{jej}^{-1} \psi_{j'} (I - A_{11})' + o_p(1) \]
Notice both \( \frac{1}{N} \sum_{j=1}^{N} \hat{\psi}_j \hat{\Sigma}^{-1} \hat{\psi}'_j \) and \( \frac{1}{N} \sum_{j=1}^{N} \psi_j^* \hat{\Sigma}^{-1} \psi_j^* \) are of full rank, then \( I - A_{11} \) is of full rank. Given this result, in combination with \( \frac{1}{N} \sum_{j=1}^{N} \hat{\psi}_j \hat{\Sigma}^{-1} \hat{\psi}'_j = (I - A_{11}) \frac{1}{N} \sum_{j=1}^{N} \psi_j^* \hat{\Sigma}^{-1} \psi_j^* + o_p(1) \), we have \( \frac{1}{N} \sum_{j=1}^{N} \hat{\psi}_j \hat{\Sigma}^{-1} \hat{\psi}'_j \) is of full rank. Given this result, in combination with (B.6), we have \( A_{12} = o_p(1) \). By \( A_{12} = o_p(1) \), together with (B.1), we have \( A_{21} = o_p(1) \) and \( (I - A_{11}) (I - A_{11})' = I + o_p(1) \). Given these results, in conjunction with the identification conditions \( IZ_2 \), assuming the column signs are known, by Lemma A.2, we have \( A_{11} = o_p(1) \) and \( A_{22} = o_p(1) \). Then \( A = o_p(1) \). The remaining proof is same as Proposition 2.1 and hence omitted. This completes the proof for consistencies of the estimates.

**Corollary B.1.** Under Assumptions A-E, together with the identification conditions \( IZ \), we have

(a) \[ \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}^{-1} \hat{\Gamma} - \frac{1}{N} \Gamma^* \Sigma^{* -1} \Gamma^* = o_p(1); \quad \frac{1}{N} (\hat{\Gamma}' - \Gamma^*) \Sigma^{* -1} \hat{\Gamma} = o_p(1) \]

(b) \[ \frac{1}{N} \sum_{i=1}^{N} \hat{\psi}_i \hat{\Sigma}^{-1} \psi_i' - \frac{1}{N} \sum_{i=1}^{N} \psi_i^* \Sigma^{-1} \psi_i^* = o_p(1) \]

\[ \frac{1}{N} \sum_{i=1}^{N} (\hat{\psi}_i - \psi_i^*) \hat{\Sigma}^{-1} \psi_i' = o_p(1) \]

(c) \[ \frac{1}{N} \sum_{i=1}^{N} \hat{\gamma}_{ix} \hat{\Sigma}^{-1} \gamma_{ix}' - \frac{1}{N} \sum_{i=1}^{N} \gamma_{ix}^* \Sigma^{* -1} \gamma_{ix}^* = o_p(1) \]

\[ \frac{1}{N} \sum_{i=1}^{N} (\hat{\gamma}_{ix} - \gamma_{ix}^*) \hat{\Sigma}^{-1} \gamma_{ix}' = o_p(1) \]

**Proof of Corollary B.2.** Results (a) and (b) have already been proved in the proof of Proposition B.1. Notice that

\[ \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}^{-1} \hat{\Gamma} = \frac{1}{N} \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}^{-1} \lambda_i' + \frac{1}{N} \sum_{i=1}^{N} \hat{\gamma}_{ix} \hat{\Sigma}^{-1} \gamma_{ix}' \]

a similar expression also holds for \( \frac{1}{N} \Gamma^* \Sigma^{* -1} \Gamma^* \). Using results (a) and (b), the first part of (c) follows immediately. The second part of (c) can be proved in the same way. So Corollary B.2 follows.

**APPENDIX B2: PROOF OF THE CONVERGENCE RATES**

From now on, we drop the superscript "*" from the true parameters. Symbols with a hat represents the MLE estimators. Those without a hat
denote the true parameters. To derive the convergence rates, we need the following lemma.

**Lemma B.2.** Let $\mathcal{E} = O_p(T^{-1/2}) + O_p((\frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2)^{1/2}) + O_p(\frac{1}{N} \sum_{i=1}^{N} \|\hat{\Sigma}_{ii}^{-1}\|\|\hat{\Gamma}_i - \Gamma_i\|^2) + O_p(\|\hat{\beta} - \beta\|)$. Under Assumptions A-E, we have

$$A = \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i)\hat{\Sigma}_{ii}^{-1}\hat{\Gamma}_i'\hat{H} = \mathcal{E}$$

**Proof of Lemma B.2.** By equation (3.4), (A.21) continues to hold in the present context. As remarked at the beginning of Appendix A, Lemmas A.5 and A.6 still hold. By Lemmas A.5 and A.6, in combination with Corollary B.2(a), we have

$$A + A' - AA' = \mathcal{E}$$

where $\mathcal{E}$ is defined earlier, which emphasize the order of magnitude. We partition the matrix $A$ as

$$A = \begin{bmatrix} A_{11} & A_{12} \\ A_{21} & A_{22} \end{bmatrix}$$

Then equation (B.8) is equivalent to

(B.9) \[ A_{11} + A_{11}' - A_{11}A_{11}' - A_{12}A_{12}' = \mathcal{E} \]

(B.10) \[ A_{12} + A_{21}' - A_{11}A_{21}' - A_{12}A_{22}' = \mathcal{E} \]

(B.11) \[ A_{21} + A_{12}' - A_{21}A_{12}' - A_{22}A_{12}' = \mathcal{E} \]

(B.12) \[ A_{22} + A_{22}' - A_{22}A_{12}' - A_{21}A_{21}' = \mathcal{E} \]

Again $\mathcal{E}$ signifies the order of magnitude. Equation (B.11) is equal to

(B.13) \[ A_{21}(I - A_{11}') + (I - A_{22})A_{12}' = \mathcal{E} \]

Lemma B.1 shows that the right hand side of (B.5) is $\mathcal{E}$. The left hand side of (B.5) is

$$A_{12}' \frac{1}{N} \sum_{j=1}^{N} \psi_j \hat{\Sigma}_{jj}^{-1} \psi_j' = A_{12}' \left[ \frac{1}{N} \sum_{j=1}^{N} \psi_j \hat{\Sigma}_{jj}^{-1} \psi_j' + o_p(1) \right]$$

by Corollary B.2(b). Because $\frac{1}{N} \sum_{j=1}^{N} \psi_j \Sigma_{jj}^{-1} \psi_j' > 0$, we have $A_{12} = \mathcal{E}$. Notice $I - A_{11} \overset{p}{\to} I$ and $I - A_{22} \overset{p}{\to} I$ by $A \overset{p}{\to} 0$. Equation (B.13) implies $A_{21} = \mathcal{E}$. 
Given $A_{12} = \mathcal{E}$ and $A_{21} = \mathcal{E}$, and $A_{11} \overset{p}{\to} 0$, $A_{22} \overset{p}{\to} 0$, equations (B.9) and (B.12) imply

\[
A_{11} + A'_{11} = \mathcal{E} \\
A_{22} + A'_{22} = \mathcal{E}
\]

Let $\hat{Q} = \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}$ (recall in Assumption C, $Q$ is defined as $\lim_{N \to \infty} \frac{1}{N} \hat{\Gamma}' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}$),

\[
\hat{Q}_{11} = \frac{1}{N} \hat{\Gamma}^g \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^g, \quad \hat{Q}_{12} = \frac{1}{N} \hat{\Gamma}^g \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^h, \quad \hat{Q}_{21} = \frac{1}{N} \hat{\Gamma}^h \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^g, \quad \text{and} \quad \hat{Q}_{22} = \frac{1}{N} \hat{\Gamma}^h \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^h.
\]

By the definition of $\hat{Q}$ and $\hat{H}_N (H_N = N \hat{\hat{H}})$, we have

\[
\hat{H}_N = \begin{bmatrix}
\hat{Q}_{11} & \hat{Q}_{12} \\
\hat{Q}_{21} & \hat{Q}_{22}
\end{bmatrix}
\]

This gives

\[
\hat{H}_N = \begin{bmatrix}
\hat{Q}_{11}^{-1} + \hat{Q}_{12}^{-1} \hat{Q}_{12} \hat{Q}_{21}^{-1} \hat{Q}_{11} - \hat{Q}_{11}^{-1} \hat{Q}_{12} \hat{Q}_{22}^{-1} \\
-\hat{Q}_{22}^{-1} \hat{Q}_{21} \hat{Q}_{11}^{-1} & \hat{Q}_{22}^{-1}
\end{bmatrix}
\]

where $\hat{Q}_{22}^{-1} = \hat{Q}_{22} - \hat{Q}_{21} \hat{Q}_{11}^{-1} \hat{Q}_{12}$. Furthermore, let $A = \frac{1}{N}(\hat{\Gamma} - \Gamma)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}$ and $A_{11} = \frac{1}{N}(\hat{\Gamma}^g - \Gamma^g)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^g$, $A_{12} = \frac{1}{N}(\hat{\Gamma}^g - \Gamma^g)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^h$, $A_{21} = \frac{1}{N}(\hat{\Gamma}^h - \Gamma^h)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^g$, $A_{22} = \frac{1}{N}(\hat{\Gamma}^h - \Gamma^h)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^h$. From $A = \hat{A} H_N$, we have

\[
A_{12} = -A_{11} \hat{Q}_{11}^{-1} \hat{Q}_{12} \hat{Q}_{22}^{-1} + A_{12} \hat{Q}_{22}^{-1}
\]

\[
A_{11} = A_{11} (\hat{Q}_{11}^{-1} + \hat{Q}_{12}^{-1} \hat{Q}_{12} \hat{Q}_{21}^{-1} (\hat{Q}_{21}^{-1})) + A_{12} \hat{Q}_{22}^{-1} \hat{Q}_{21} \hat{Q}_{11}^{-1}
\]

So we have $A_{11} = A_{11} \hat{Q}_{11}^{-1} - A_{12} \hat{Q}_{21} \hat{Q}_{11}^{-1}$. By Corollary B.2(a), we have $\hat{Q}_{21} \overset{p}{\to} Q_{21}, \hat{Q}_{11} \overset{p}{\to} Q_{11}$. Given $A_{12} = \mathcal{E}$, we have

(B.14)

\[
A_{11} = A_{11} \hat{Q}_{11}^{-1} + \mathcal{E}
\]

Substituting (B.14) into $A_{11} + A'_{11} = \mathcal{E}$, we have

(B.15)

\[
A_{11} \hat{Q}_{11}^{-1} + \hat{Q}_{11} A'_{11} = \mathcal{E}
\]

By similar arguments, we also have

\[
A_{22} \hat{Q}_{22}^{-1} + \hat{Q}_{22} A'_{22} = \mathcal{E}
\]

The identification condition $IZ2$ is equal to

\[
\text{ndiag}\left\{ \frac{1}{N} (\hat{\Gamma}^g - \Gamma^g)' \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}^g + \frac{1}{N} \hat{\Gamma}^g \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma}^g - \Gamma^g) \right\}
\]

(B.16)

\[
= \text{ndiag}\left\{ \frac{1}{N} (\hat{\Gamma}^g - \Gamma^g)' \hat{\Sigma}_{ee}^{-1} (\hat{\Gamma}^g - \Gamma^g) - \frac{1}{N} \hat{\Gamma}^g (\hat{\Sigma}_{ee}^{-1} - \Sigma_{ee}^{-1}) \Gamma^g \right\}
\]
Notice \( \frac{1}{N}(\Gamma^g - \Gamma^g)\hat{\Sigma}_{\epsilon e}^{-1}(\Gamma^g - \Gamma^g) \) is \( O_p\left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{\epsilon i}^{-1} \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 \right) \), and \( \frac{1}{N} \Gamma^g(\hat{\Sigma}_{\epsilon e}^{-1} - \Sigma_{\epsilon e}^{-1}) \Gamma^g \) is \( O_p\left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{\epsilon i} - \Sigma_{\epsilon i} \|^{21/2} \right) \). So (B.16) can be written, in terms of \( A \), as

\[
\text{ndiag}(A_{11} + A'_{11}) = \mathcal{E}
\]

This equation together with (B.15) implies \( A_{11} = \mathcal{E} \) by using the arguments of Proposition A.1. Matrix \( A_{22} \) can be proved to be \( \mathcal{E} \) similarly as \( A_{11} \). Given the results of \( A_{11}, A_{22}, A_{12}, A_{21} \), Lemma B.2 follows.

**Proposition B.2.** Under Assumptions A-E, together with the identification conditions \( IZ \), we have

\[
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{\epsilon i}^{-1} \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 = O_p(T^{-1}) + O_p(\| \hat{\beta} - \beta \|^2)
\]

\[
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{\epsilon i} - \Sigma_{\epsilon i} \|^2 = O_p(T^{-1}) + O_p(\| \hat{\beta} - \beta \|^2)
\]

**Proof of Proposition B.2.** The first order condition (3.6) is equal to

\[
\hat{\gamma}_{jx} - \gamma_{jx} = -A'\hat{\gamma}_{jx} + \hat{H} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{\epsilon i}^{-1} \hat{\Gamma}_i' \frac{1}{T} \sum_{t=1}^{T} f_t v'_{jtx} - \hat{H} \hat{\gamma}_{jx} \hat{\Sigma}_{\epsilon jx}^{-1} \hat{\Sigma}_{\epsilon jx} - \hat{\Sigma}_{\epsilon jx}
\]

\[
(B.17)
\]

\[
-\hat{H} \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta) f_t' \right) \hat{\gamma}_{jx} - \hat{H} \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta) v'_{jtx} \right)
\]

where \( A = (\hat{\Gamma} - \Gamma)'\hat{\Sigma}_{\epsilon e}^{-1} \hat{\Gamma} \hat{H} \). We use \( q_1, q_2, \cdots, q_7 \) to denote the 7 terms on the right hand side. Given \( \hat{\Sigma}_{\epsilon ix}^{-1} \) is bounded in a compact set by Assumption D, we have

\[
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{\epsilon ix}^{-1} \cdot \| \hat{\gamma}_{ix} - \gamma_{ix} \|^2 \leq C \frac{1}{N} \sum_{i=1}^{N} \| \hat{\gamma}_{ix} - \gamma_{ix} \|^2
\]

\[
\leq 7C \frac{1}{N} \sum_{i=1}^{N} (\| q_1 \|^2 + \| q_2 \|^2 + \cdots + \| q_7 \|^2)
\]
We can derive the bound for each term of the above equation. This process is similar to the derivation of (A.31) in Appendix A. Thus we state the result without writing out the details.

\[
\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{itx}^{-1}|| \cdot ||\hat{\gamma}_{ix} - \gamma_{ix}||^2 = O_p(T^{-1}) + O_p\left(\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii} - \Sigma_{ii}||^2\right) + O_p(\|\hat{\beta} - \beta\|^2)
\]

Also, By (B.4) we can derive a similar result for \(\hat{\psi}_j - \psi_j\),

\[
\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{itx}^{-1}|| \cdot ||\hat{\psi}_i - \psi_i||^2 = O_p(T^{-1}) + O_p\left(\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii} - \Sigma_{ii}||^2\right) + O_p(\|\hat{\beta} - \beta\|^2)
\]

Combining the above results and noticing that

\[
\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii}^{-1}|| \cdot ||\hat{\Gamma}_i - \Gamma_i||^2 \leq \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{itx}^{-1}|| \cdot ||\hat{\gamma}_{ix} - \gamma_{ix}||^2 + \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{itx}^{-1}|| \cdot ||\hat{\psi}_i - \psi_i||^2
\]

we have

\[
(B.18) \quad \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii}^{-1}|| \cdot ||\hat{\Gamma}_i - \Gamma_i||^2 = O_p(T^{-1}) + O_p\left(\frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{ii} - \Sigma_{ii}||^2\right) + O_p(\|\hat{\beta} - \beta\|^2)
\]

Next, we study the first order condition (3.7). Using the fact that

\[
\hat{B}M_{ij} \hat{B}' - \hat{\Sigma}_{ij} = \Gamma_i\Gamma_j + \Gamma_i' \frac{1}{T} \sum_{t=1}^{T} f_t e_{jt} + \frac{1}{T} \sum_{t=1}^{T} \varepsilon_u f_t' \Gamma_j + \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_u e_{jt}' - E(\varepsilon_u e_{jt}')] - \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{\psi}_j + \frac{1}{T} \sum_{t=1}^{T} \varepsilon_u (\hat{\beta} - \beta)' \hat{\psi}_j
\]

\[
= -\frac{1}{T} \sum_{t=1}^{T} \hat{i}_u(\hat{\beta} - \beta)\varepsilon_{jt}' + \frac{1}{T} \sum_{t=1}^{T} \hat{i}_u(\hat{\beta} - \beta)(\hat{\beta} - \beta)' \hat{\psi}_j - \hat{\Gamma}_j (\hat{\psi}_j - 1\{i = j\}) (\hat{\Sigma}_{ij} - \Sigma_{ij})
\]

we have

\[
\hat{\Sigma}_{je} - \Sigma_{je} = \frac{1}{T} \sum_{t=1}^{T} (\varepsilon_{jt}' - \Sigma_{je}) + (\hat{\lambda}_j - \lambda_j)'(\hat{\lambda}_j - \lambda_j) - 2\lambda_j' \hat{G}(\hat{\lambda}_j - \lambda_j)
\]

\[
- 2(\hat{\lambda}_j - \lambda_j)' \frac{1}{T} \sum_{t=1}^{T} f_t e_{jt} + 2\lambda_j' \hat{G} \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i)' \frac{1}{T} \sum_{t=1}^{T} f_t e_{jt} + 2\lambda_j' \hat{G} \frac{1}{T} \sum_{t=1}^{T} f_t e_{jt}
\]
\[-2\lambda_j^* \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} \frac{1}{T} \sum_{t=1}^T [\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})] + 2\lambda_j^* \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}_i - \Gamma_i)' \lambda_j \]

\[-2\lambda_j^* \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^T \hat{\chi}_t f'_t \right) \lambda_j + 2\lambda_j^* \hat{G} \hat{\lambda}_j \hat{\Sigma}^{-1}_{je} (\hat{\Sigma}_{je} - \Sigma_{je}) \]

\[+ \mathcal{V}_{j1} \]

where

\[\mathcal{V}_{j1} = -2\lambda_j^* \frac{1}{T} \sum_{t=1}^T f_t \hat{x}_{jt} (\hat{\beta} - \beta) + \frac{1}{T} \sum_{t=1}^T \hat{x}_{jt} (\hat{\beta} - \beta)(\hat{\beta} - \beta)' \hat{x}'_{jt} \]

\[+ 2\lambda_j^* \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} f'_i \sum_{t=1}^T f_t \hat{x}_{jt} (\hat{\beta} - \beta) + 2\lambda_j^* \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^T \hat{\chi}_t \hat{x}_{jt} \right) (\hat{\beta} - \beta) \]

\[+ 2\lambda_j^* \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^T \hat{\xi}_t (\hat{\beta} - \beta) f'_t \right) \lambda_j + 2\lambda_j^* \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^T \hat{\xi}_t (\hat{\beta} - \beta) e_{jt} \right) \]

\[-2 \frac{1}{T} \sum_{t=1}^T e_{jt} \hat{x}_{jt} (\hat{\beta} - \beta) - 2\lambda_j^* \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^T \hat{F}'_t (\hat{\beta} - \beta)(\hat{\beta} - \beta)' \hat{x}'_{jt} \right) \]

We also have

\[\hat{\Sigma}_{jx} - \Sigma_{jx} = \frac{1}{T} \sum_{t=1}^T (v_{jtx} v_{jtx} - \Sigma_{jx}) + (\hat{\gamma}_{jx} - \gamma_{jx})' (\hat{\gamma}_{jx} - \gamma_{jx}) \]

\[+ \gamma_{jx}^* \sum_{i=1}^N (\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}^{-1}_{ii} \hat{\Gamma}'_i \hat{\gamma}_{jx} + \hat{\gamma}_{jx}' \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}_i - \Gamma_i)' \gamma_{jx} - (\hat{\gamma}_{jx} - \gamma_{jx})' \hat{G} \hat{\gamma}_{jx} \]

\[-\hat{\gamma}_{jx}' \hat{G} (\hat{\gamma}_{jx} - \gamma_{jx}) - (\hat{\gamma}_{jx} - \gamma_{jx})' \frac{1}{T} \sum_{t=1}^T f_t v_{jtx} - \frac{1}{T} \sum_{t=1}^T v_{jtx} f'_t (\hat{\gamma}_{jx} - \gamma_{jx}) \]

\[+ \hat{\gamma}_{jx}' \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}_i - \Gamma_i)' \frac{1}{T} \sum_{t=1}^T f_t v_{jtx} + \frac{1}{T} \sum_{t=1}^T v_{jtx} f'_t (\hat{\gamma}_{jx} - \gamma_{jx}) \]

\[+ \hat{\gamma}_{jx}' \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}_i - \Gamma_i)' \frac{1}{T} \sum_{t=1}^T f_t v_{jtx} + \frac{1}{T} \sum_{t=1}^T v_{jtx} f'_t (\hat{\gamma}_{jx} - \gamma_{jx}) \]

\[+ \hat{\gamma}_{jx}' \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}_i - \Gamma_i)' \frac{1}{T} \sum_{t=1}^T f_t v_{jtx} + \frac{1}{T} \sum_{t=1}^T v_{jtx} f'_t (\hat{\gamma}_{jx} - \gamma_{jx}) \]

\[+ \hat{\gamma}_{jx}' \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}^{-1}_{ii} (\hat{\Gamma}_i - \Gamma_i)' \frac{1}{T} \sum_{t=1}^T f_t v_{jtx} + \frac{1}{T} \sum_{t=1}^T v_{jtx} f'_t (\hat{\gamma}_{jx} - \gamma_{jx}) \]

\[- \frac{1}{T} \sum_{t=1}^T [v_{jtx} e_{it} - v_{jtx} e'_{it}] \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}'_i \hat{G} \hat{\gamma}_{jx} + (\hat{\Sigma}_{jx} - \Sigma_{jx}) \hat{\Sigma}^{-1}_{jx} \hat{\gamma}'_{jx} \hat{G} \hat{\gamma}_{jx} \]
\begin{align*}
+ \gamma'_{jx} \hat{G}_{jx} \hat{\Sigma}_{jx}^{-1} (\hat{\Sigma}_{jx} - \Sigma_{jx}) - \gamma'_{jx} \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f'_t \right) \gamma_{jx} - \gamma'_{jx} \left( \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\chi}_t' \right) \hat{G}_N \gamma_{jx} \\
+ \mathcal{V}_{j2}
\end{align*}

where

\begin{align*}
\mathcal{V}_{j2} &= \gamma'_{jx} \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) f'_t \right) \gamma_{jx} + \gamma'_{jx} \left( \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta) \hat{\xi}_t' \right) \hat{G}_N \gamma_{jx} \\
&+ \mathcal{V}_{jx} \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{v}_{jx} (\hat{\beta} - \beta) \hat{\xi}_t' \right) + \left( \frac{1}{T} \sum_{t=1}^{T} v_{jx} (\hat{\beta} - \beta) \hat{\xi}_t' \right) \hat{G}_N \gamma_{jx}
\end{align*}

Similarly as in Appendix A, we can prove

\begin{align*}
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 &= O_p(T^{-1}) + o_p \left( \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii}^{-1} \| \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 \right) + O_p(\| \hat{\beta} - \beta \|^2)
\end{align*}

Moreover,

\begin{align*}
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{iix} - \Sigma_{iix} \|^2 \\
&= \frac{1}{N} \sum_{i=1}^{N} (\hat{\Sigma}_{iix} - \Sigma_{iix})^2 + \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{iix} - \Sigma_{iix} \|^2
\end{align*}

We have

\begin{align*}
\frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jj} - \Sigma_{jj} \|^2 &= O_p(T^{-1}) + o_p \left( \frac{1}{N} \sum_{j=1}^{N} \| \hat{\Sigma}_{jj}^{-1} \| \cdot \| \hat{\Gamma}_j - \Gamma_j \|^2 \right) \\
&+ O_p(\| \hat{\beta} - \beta \|^2)
\end{align*}

(B.19)

Substituting (B.19) into (B.18), we get

\begin{align*}
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{i}^{-1} \| \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 &= O_p(T^{-1}) + O_p(\| \hat{\beta} - \beta \|^2)
\end{align*}

Substituting the above result into (B.19), we have

\begin{align*}
\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{i} - \Sigma_{i} \|^2 &= O_p(T^{-1}) + O_p(\| \hat{\beta} - \beta \|^2)
\end{align*}

This completes the proof of Proposition B.2.

**Proposition B.3.** Under Assumptions A-E, we have

\begin{align*}
\hat{\beta} - \beta &= O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1})
\end{align*}
To prove Proposition B.3, we need the following lemmas.

**Lemma B.3.** Under Assumptions A-E, together with the identification conditions IZ, we have

(a) \( \vartheta_{pq} = \text{tr}(\omega_p) - \text{tr}(\nu'_pQ^{-1}v_q) + \Omega_{pq} + o_p(1) \)

(b) \( \text{tr}\left[ \hat{G}_N \frac{1}{T} \sum_{t=1}^{T} \xi_t (\hat{\beta} - \beta) f'_t \hat{v}_p \right] = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}[v_pQ^{-1}v_q] + o_p(\|\hat{\beta} - \beta\|) \)

(c) \( \text{tr}\left[ \hat{G}_1 \Gamma^x_{\varepsilon \varepsilon} \Gamma^1_{1} \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{G}_p' \right] = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}(\omega_{pq}^g) + o_p(\|\hat{\beta} - \beta\|) \)

(d) \( \text{tr}\left[ (I - A)' \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{H}_N \hat{v}_t' \right] = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}[v_pQ^{-1}v_q] + o_p(\|\hat{\beta} - \beta\|) \)

(e) \( \text{tr}\left[ \hat{H}_N \frac{1}{T} \sum_{t=1}^{T} \xi_t (\hat{\beta} - \beta)' f'_t (I - A) \hat{v}_t' \right] = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}[v_pQ^{-1}v_q] + o_p(\|\hat{\beta} - \beta\|) \)

(f) \( \text{tr}\left[ \hat{G}_2 \Gamma^x_{\varepsilon \varepsilon} \Gamma^1_{1} \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{H}_t \hat{\Pi}^{-1}_{\psi \psi} \vartheta_p \right] = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}\left[ \hat{G}_p' \hat{\Pi}^{-1}_{\psi \psi} \vartheta_p \right] + o_p(\|\hat{\beta} - \beta\|) \)

where the symbols \( \omega_{pq}, \Pi_{\psi \psi}, \hat{\Pi}_{\psi \psi}, \nu_p, \hat{v}_p, \hat{H}_t, \hat{G}_p' \) and \( \xi_t \) are all defined in Table 1 and 2, and

\[
\vartheta_{pq} = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{itp} \hat{x}_{itp} \hat{x}_{itq} - \frac{1}{NT} \sum_{t=1}^{T} \left( \sum_{i=1}^{N} \hat{\Sigma}_{itp} \hat{x}_{itp} \hat{\lambda}_j \right) \hat{G} \left( \sum_{j=1}^{N} \hat{\Sigma}_{itp}^{-1} \hat{\lambda}_j \hat{x}_{itq} \right).
\]

**Proof of Lemma B.3.** The proof of this lemma is quite similar to that of Lemma A.10. To save space, we only prove (b) and (f).

Consider (b). First we show \( \hat{v}_p = v_p + o_p(1) \). Notice

\[
\frac{1}{N} \sum_{j=1}^{N} \lambda_j \hat{\Sigma}_{jpe}^{-1} \gamma'_j p = \frac{1}{N} \sum_{j=1}^{N} \lambda_j \Sigma_{jpe}^{-1} \gamma'_j p + \frac{1}{N} \sum_{j=1}^{N} \lambda_j (\Sigma_{jpe}^{-1} - \Sigma_{jpe}^{-1}) \gamma'_j p
\]

The second term is bounded in norm by \( C \left( \frac{1}{N} \sum_{i=1}^{N} (\hat{\Sigma}_{iie} - \Sigma_{iie})^2 \right)^{1/2} \), which is \( o_p(1) \) by Proposition B.1. By the definition of \( v_p \) we have \( \hat{v}_p = v_p + o_p(1) \).

However, \( \frac{1}{T} \sum_{t=1}^{T} \xi_t (\hat{\beta} - \beta)' f'_t \) is equal to \( \sum_{q=1}^{K} v_q (\hat{\beta}_q - \beta_q) + o_p(\|\hat{\beta} - \beta\|) \), which is given in result (b) in Lemma A.10. By \( \hat{G}_N - Q^{-1} \xrightarrow{p} 0 \), we have (b).
Consider (f). First, we can show \( \hat{q}_p = q_p + o_p(1) \) similarly as \( \hat{v}_p = v_p + o_p(1) \). Then for the term \( \hat{\Pi}_{\psi\psi} \), by its definition, it’s equal to

\[
\frac{1}{N} \sum_{i=1}^{N} \psi_i \Sigma_{tie}^{-1} \psi_i' + \frac{1}{N} \sum_{i=1}^{N} \psi_i \Sigma_{tie}^{-1} (\hat{\psi}_i - \psi_i)' + \frac{1}{N} \sum_{i=1}^{N} \psi_i (\Sigma_{tie}^{-1} - \Sigma_{tie}) \hat{\psi}_i'
\]

The second expression is bounded by \( C[\frac{1}{N} \sum_{i=1}^{N} (\Sigma_{tie}^{-1} - \Sigma_{tie})^2]^{1/2} \), and the third is bounded by \( C[\frac{1}{N} \sum_{i=1}^{N} (\Sigma_{tie}^{-1} - \Sigma_{tie})^2]^{1/2} \). By Proposition B.1, these two terms are both \( o_p(1) \). It follows that \( \Pi_{\psi\psi} = \Pi_{\psi\psi} + o_p(1) \).

Now consider the term \( \frac{1}{T} \sum_{t=1}^{T} f_t(\hat{\beta} - \beta)' \hat{\Gamma}_t \). By the definition of \( \hat{\Gamma}_t \), it can be written as

\[
\sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t \hat{x}_{jq} \hat{\Sigma}_{ie}^{-1} \psi_j' \right) + \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t v_{jq} \hat{\Sigma}_{ie}^{-1} \hat{\psi}_j' \right)
\]

where \( \hat{x}_{jq} \) can be replaced with \( v_{jq} \) since \( \frac{1}{T} \sum_{t=1}^{T} f_t = 0 \). Notice \( \frac{1}{N} \sum_{j=1}^{N} \gamma_{jq} \Sigma_{ie}^{-1} \hat{\psi}_j' = \frac{1}{N} \sum_{j=1}^{N} \sum_{q=1}^{K} \gamma_{jq} \Sigma_{ie}^{-1} \psi_j' + o_p(1) \), which can be proved similarly as \( \Pi_{\psi\psi} = \Pi_{\psi\psi} + o_p(1) \), and \( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t v_{jq} \hat{\Sigma}_{ie}^{-1} \hat{\psi}_j' = o_p(1) \). Given these results, we have

\[
\frac{1}{T} \sum_{t=1}^{T} f_t(\hat{\beta} - \beta)' \hat{\Gamma}_t = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \left( \frac{1}{N} \sum_{j=1}^{N} \gamma_{jq} \Sigma_{ie}^{-1} \psi_j' \right) + o_p(\|\hat{\beta} - \beta\|)
\]

The above result, together with \( \hat{C}_2 \hat{\Gamma}_t \hat{\Sigma}_{ie}^{-1} \Gamma_t \xrightarrow{P} I_2 + o_p(1) \), \( \hat{q}_p = q_p + o_p(1) \) and \( \Pi_{\psi\psi} = \Pi_{\psi\psi} + o_p(1) \), we have (f).

**Lemma B.4.** Under Assumptions A-E, together with the identification conditions \( IZ \), we have

(a) \[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{ie}^{-1} \hat{\gamma}_{ip} h_t c_{it} = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ie}^{-1} \gamma_{ip} h_t c_{it} + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|)
\]
\[(b) \quad \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{\pi i}^{-1} e_{it} v_{itp} = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{\pi i}^{-1} e_{it} v_{itp} + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|) \]

\[(c) \quad \hat{G}_2 \hat{\Sigma}_{cc}^{-1} \Gamma e_j e_j = \frac{1}{NT} \sum_{j=1}^{N} h_i e_j \hat{\Sigma}_{ji}^{-1} \psi_i \hat{\Pi}^{-1} \psi_i \hat{\eta}_p \]

\[(d) \quad \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ji}^{-1} \frac{1}{T} \sum_{t=1}^{T} [e_{it} v_{itp} - E(e_{it} v_{itp})] \hat{\Sigma}_{\pi i}^{-1} \hat{\lambda}_i \hat{G} = O_p(N^{-1}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|) \]

\[(e) \quad \hat{G}_1 \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_j \hat{\Sigma}_{ji}^{-1} \frac{1}{T} \sum_{t=1}^{T} [e_{it} e_{jt} - E(e_{it} e_{jt})] \hat{\Sigma}_{ji}^{-1} \hat{\gamma}_j \hat{G} = O_p(N^{-1}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|) \]

\[(f) \quad \hat{H} \frac{1}{T} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ji}^{-1} [e_{it} e_{jt} - E(e_{it} e_{jt})] \hat{\Sigma}_{ji}^{-1} \hat{\gamma}_j \hat{H} = O_p(N^{-1}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|) \]

\[(g) \quad \frac{1}{NT} \hat{G}_2 \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ji}^{-1} \frac{1}{T} \sum_{t=1}^{T} [e_{it} e_{jt} - E(e_{it} e_{jt})] \hat{\Sigma}_{ji}^{-1} \hat{\gamma}_j \hat{G} \quad = O_p(N^{-1}T^{-1/2}) + O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|) \]

**Proof of Lemma B.4.** The proofs of the results in this lemma are quite similar to those of Lemma A.12. To save the space, we prove (c) as an illustration.

Notice \(\hat{G}_2 \hat{\Sigma}_{cc}^{-1} \Gamma = I_2 + o_p(1)\) by Corollary B.2(a) and \(\frac{1}{N} \sum_{j=1}^{N} \psi_j \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j = \frac{1}{N} \sum_{j=1}^{N} \psi_j \Sigma_{ji}^{-1} \psi_j + o_p(1)\) by Lemma B.3(f). Term \(\frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t e_j \hat{\Sigma}_{ji}^{-1} \psi_j = O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|).\) Term \(\frac{1}{N} \sum_{j=1}^{N} \psi_j \hat{\Sigma}_{ji}^{-1} \hat{\gamma}_j \) is \(\frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t e_j \hat{\Sigma}_{ji}^{-1} \psi_j = O_p(T^{-1}) + o_p(\|\hat{\beta} - \beta\|).\) Given this result, (c) follows.

**Proof of Proposition B.3.** By (3.8), equation (A.41) still holds in the present context. Consider the first term on the right hand side of (A.41). Notice \(\hat{M}_{ij} = I_r\) and \(\hat{\lambda}_j = (\hat{\psi}_j, 0_{r \times 1})'\). The first term is equal to \(-\text{tr}[\frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{\pi i}^{-1} (\hat{\psi}_i - \psi_i)']\), using (B.4) to replace \(\hat{\psi}_i - \psi_i\), we can rewrite (A.41) as

\[(B.20) \quad \sum_{q=1}^{K} \partial_{pq}(\hat{\beta}_q - \beta_q) = -\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \gamma_{ip} f_t \hat{\Sigma}_{\pi i}^{-1} \hat{\lambda}_i \hat{G}_N \hat{\lambda}_t \]
where $\hat{\gamma}_p^g$, $\hat{G}_p^g$, $\hat{\lambda}_t$ and $\hat{\xi}_t$ are defined in Tables 1 and 2; $\hat{G} = (\hat{\Gamma} - \Gamma')\hat{\Sigma}^{-1}(\hat{\Gamma} - \Gamma)$. All terms except the 4th and 14th term can be proved to be $O_p(N^{-1/2}T^{-1/2}) + O_p(\|\hat{\beta} - \beta\|)$. The 4th and 14th terms are each $O_p(T^{-1/2})$ by Lemma B.2. But they share common components that are offset each other. To see this, by $\hat{G} = \hat{H} - \hat{H}\hat{G}$, the sum of the 4th and 14th terms is

\begin{equation}
\text{tr}\left[\hat{G}_1 N \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \hat{\gamma}_p^g \right] + \text{tr}\left[\hat{G}_1 N \sum_{i=1}^N \hat{\xi}_t (\hat{\beta} - \beta) \hat{G}_p^g \right]
\end{equation}

(B.21)
\[
\begin{align*}
&= \operatorname{tr}[\hat{\Lambda}^t \hat{\varphi}_p] - \operatorname{tr}\left[\hat{G}_2 \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i) \frac{1}{N} \sum_{j=1}^N \lambda_j \hat{\Sigma}_{ji}^{-1} \gamma_{jp}^t \right] \\
&\quad + \operatorname{tr}[\hat{\varphi}_p \hat{\Lambda}^t] + \operatorname{tr}\left[\frac{1}{N} \sum_{i=1}^N \gamma_{ip} \hat{\Sigma}_{ii}^{-1} (\hat{\lambda}_i - \lambda_i)^t \hat{\Lambda}^t \right] \\
&= \operatorname{tr}[(A + A') \hat{\varphi}_p] - \operatorname{tr}[\hat{H} \hat{\Lambda} \hat{\varphi}_p] + \operatorname{tr}\left[\frac{1}{N} \sum_{i=1}^N \gamma_{ip} \hat{\Sigma}_{ii}^{-1} (\hat{\lambda}_i - \lambda_i)^t \hat{\Lambda}^t \right] \\
&\quad - \operatorname{tr}[A \hat{G} \hat{\varphi}_p] - \operatorname{tr}\left[\hat{G}_2 \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i) \frac{1}{N} \sum_{j=1}^N \lambda_j \hat{\Sigma}_{ji}^{-1} \gamma_{jp}^t \right]
\end{align*}
\]

where \( A = (\hat{\Gamma} - \Gamma)^t \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \hat{H} \). The first term involves \( A + A' \), However, by (A.21) and noting \( \hat{M}_{ff} = M_{ff} = I_r \), we have

\[
(B.22) \quad A + A' = A' A + (I - A)^t \left[ \frac{1}{T} \sum_{t=1}^T \hat{f}_t \hat{\chi}^t \right] \hat{H}_N
\]

\[
\begin{align*}
&+ \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^T \hat{\chi}_t \hat{f}_t \right] (I - A) + \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^T \sum_{i=1}^N \sum_{j=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \epsilon_{i,j,t} \hat{\Sigma}_{jj}^{-1} \hat{\Gamma}_j^t \right] \hat{H}_N \\
&\quad - (I - A)^t \left[ \frac{1}{T} \sum_{t=1}^T \hat{f}_t (\hat{\beta} - \beta)^t \hat{\xi}_t \right] \hat{H}_N - \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^T \hat{\chi}_t (\hat{\beta} - \beta)^t \hat{\xi}_t \right] \hat{H}_N \\
&\quad - \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^T \hat{\xi}_t (\hat{\beta} - \beta)^t \hat{f}_t \right] (I - A) - \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^T \hat{\xi}_t (\hat{\beta} - \beta)^t \hat{\chi}_t \right] \hat{H}_N \\
&\quad + \hat{H}_N \left[ \frac{1}{T} \sum_{t=1}^T \hat{\xi}_t (\hat{\beta} - \beta) (\hat{\beta} - \beta)^t \hat{\xi}_t \right] \hat{H}_N - \hat{H}_N \hat{\Gamma}_t \hat{\Sigma}_{ee}^{-1} (\hat{\Sigma}_{ee} - \Sigma_{ee}) \hat{\Sigma}_{ee}^{-1} \hat{\Gamma}_t
\end{align*}
\]

The last term of (B.21) is equal to \( \operatorname{tr}[\hat{G}_2 \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i^q - \Gamma_i^q)^t \hat{\varphi}_p] \), which involves \( \hat{G}_2 \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i^q - \Gamma_i^q)^t \). This term is implicitly given in (B.5). To see this, notice that the 8th term of (B.5) from the right hand side to the left, then the left hand side now is

\[
\hat{H}_2 \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i)^t \hat{\Pi}_{\lambda \psi} - \hat{H}_2 \hat{G} \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i^q - \Gamma_i^q)^t \hat{\Pi}_{\lambda \psi}
\]

Notice \( \hat{G}_2 = \hat{H}_2 - \hat{H}_2 \hat{G} \) by \( \hat{G} = \hat{H} - \hat{H} \hat{G} \). The above expression is

\[
\hat{G}_2 \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i)^t \hat{\Pi}_{\lambda \psi} = \hat{G}_2 \sum_{i=1}^N \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i^q - \Gamma_i^q)^t \frac{1}{N} \sum_{j=1}^N \psi_j \hat{\Sigma}_{ji}^{-1} \hat{\varphi}_j^t
\]
Substituting (B.22) and (B.23) into (B.21), we obtain an alternative expression:

\[ \hat{G}_2 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i^g - \Gamma_i^g)^{\prime} \hat{\Pi}_{\psi\psi} = \hat{G}_{2N} \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)^{\prime} \hat{\Upsilon}_t \right) \hat{\Pi}_{\psi\psi}^{-1} \]

\[ + \frac{1}{NT} \hat{G}_2 \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ji}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\hat{\xi}_t e_{jt} - E(\hat{\xi}_t e_{jt})] \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j^{\prime} \hat{\Pi}_{\psi\psi}^{-1} \]

\[ + \hat{G}_2 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \Gamma_i \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t e_{jt} \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j^{\prime} \right) \hat{\Pi}_{\psi\psi}^{-1} - \hat{G}_{2N} \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) g_t^{\prime} \right) \]

\[ - \hat{G}_{2N} \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)^{\prime} \hat{\Upsilon}_t \right) \hat{\Pi}_{\psi\psi}^{-1} + \hat{H}_2 \hat{G} \hat{H}^{-1} \frac{1}{N} \sum_{j=1}^{N} (\hat{\lambda}_j - \lambda_j) \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j^{\prime} \hat{\Pi}_{\psi\psi}^{-1} \]

Substituting (B.22) and (B.23) into (B.21), we obtain an alternative expression of \( \text{tr}[\hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\Gamma_i - \Gamma_i)^{\prime} \hat{\psi}_i^{\prime}] + \text{tr}[\frac{1}{N} \sum_{i=1}^{N} \gamma_i e_{ip} \hat{\Lambda}^{\prime}] \). Then equation (B.20) can be rewritten as

\[ \varphi_{p1} (\hat{\beta}_1 - \beta_1) + \varphi_{p2} (\hat{\beta}_2 - \beta_2) + \cdots + \varphi_{pK} (\hat{\beta}_K - \beta_K) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{ii}^{-1} e_{it} e_{it} \]

\[ + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{ii}^{-1} e_{it} v_{itp} - \text{tr} \left[ \hat{G}_2 \hat{\Sigma}_{ee}^{-1} \Gamma \left( \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t e_{jt} \hat{\Sigma}_{ji}^{-1} \hat{\psi}_j^{\prime} \right) \hat{\Pi}_{\psi\psi}^{-1} \hat{\psi}_p^{\prime} \right] \]

\[ + \text{tr} \left[ \hat{G}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) f_t^{\prime} \right) \hat{\Pi}_{\psi\psi}^{-1} \hat{\psi}_p^{\prime} \right] + \text{tr} \left[ \hat{G}_1 \hat{\Sigma}_{ee}^{-1} \Gamma \left( \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta) \hat{\psi}_t^{\prime} \right) \right] \]

\[ - \text{tr} \left[ (I - A)^{\prime} \left( \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta) \hat{\psi}_t^{\prime} \right) \hat{H}_N \hat{\psi}_p^{\prime} \right] \]

\[ - \text{tr} \left[ \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) f_t^{\prime} (I - A) \hat{\psi}_p^{\prime} \right) \right] \]

\[ + \text{tr} \left[ \hat{G}_2 \hat{\Sigma}_{ee}^{-1} \Gamma \left( \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)^{\prime} \hat{\Upsilon}_t \right) \hat{\Pi}_{\psi\psi}^{-1} \hat{\psi}_p^{\prime} \right] \]

(B.24)
where the symbols \( \hat{v}_p, \hat{v}_p', \hat{\theta}_p, \hat{\xi}_t, \hat{H}_{\psi\psi}, \hat{H}_t \) and \( \mathcal{G}_{tp}^g \) are defined in Tables 1 and 2 and \( S_{p1}, S_{p2}, S_{p3} \) are given by

\[
S_{p1} = \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i - \lambda_i)' f_i v_{itp} A' \right] + \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i)' v_{itp} A' \right] \]

\[
- \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i - \lambda_i)' f_i v_{itp} \right] + \text{tr} \left[ \hat{H}_1 A' \right] \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i)' v_{itp} \bigg] \right] \]

\[
+ \text{tr} \left[ \hat{G}_1 \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i - \lambda_i)' \hat{H}_N \hat{G}_N \right] + \text{tr} \left[ \hat{G}_1 \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i - \lambda_i)' \hat{H}_N \hat{G}_N \right] \]

\[
- \text{tr} \left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i - \lambda_i)' \hat{H}_N \hat{G}_N \right] + \text{tr} \left[ \hat{G}_1 \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1}(\hat{\lambda}_i - \lambda_i)' \hat{H}_N \hat{G}_N \right] \]

\[
- \text{tr} \left[ \hat{H}_1 \hat{G}_N \hat{H}_t \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_i' \right] - \text{tr} \left[ \hat{v}_p A' A \right] \]

\[
+ \text{tr} \left[ \hat{v}_p \left( \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\chi}_t' \right) \hat{H}_N \hat{G}_N \right] \]

\[
- \text{tr} \left[ \hat{v}_p A' \left( \frac{1}{T} \sum_{t=1}^{T} f_t \hat{\chi}_t' \right) \hat{H}_N \right] - \text{tr} \left[ \hat{v}_p A' \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_i' \right) A \right] \]

\[
+ \text{tr} \left[ \hat{H}_2 \hat{G}_N \hat{H}_t \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_i' \hat{\psi}_j' \hat{\Pi}_{\psi \psi} \hat{\psi}_j \hat{\psi}_j' \hat{\Pi}_{\psi \psi} \hat{\psi}_j \hat{\psi}_j' \right] \]

\[
+ \text{tr} \left[ \frac{1}{N} \hat{G}_2 \frac{1}{N} \sum_{j=1}^{N} \hat{\lambda}_j \hat{\Sigma}_j^{-1}(\hat{\Sigma}_j - \Sigma_j)^{-1} \hat{\psi}_j' \hat{\psi}_j \hat{\Pi}_{\psi \psi} \hat{\psi}_j \hat{\psi}_j' \right] \]

(B.25)
and

\[ S_{p2} = \text{tr} \left[ \hat{G}_1 N \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{\epsilon}_{itp}^{g'} \right] + \text{tr} \left[ \hat{G}_1 N \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta) \sum_{j=1}^{N} e_{jt} \hat{\Sigma}^{-1}_{ji} \epsilon_{jp}^{g'} \right] \]

\[-\text{tr} \left[ \hat{G}_1 N \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta) (\hat{\beta} - \beta)' \hat{\epsilon}_{itp}^{g'} \right] - \text{tr} \left[ \hat{v}_p \hat{H}_N N \frac{1}{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{\epsilon}_t \hat{H}_N \right] \]

\[-\text{tr} \left[ \hat{v}_p \hat{H}_N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta) \hat{\epsilon}_t' \right) \hat{H}_N \right] + \text{tr} \left[ \hat{v}_p \hat{H}_N N \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta) (\hat{\beta} - \beta)' \hat{\epsilon}_t \hat{H}_N \right] \]

\[ \text{tr} \left[ \hat{G}_2 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta)' \hat{\epsilon}_t \right) \hat{H}_N \right] \]

\[ + \text{tr} \left[ \hat{G}_2 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\epsilon}_t (\hat{\beta} - \beta)' \hat{\epsilon}_t \right) \hat{H}_N \right] \]

(B.26)

The expressions \( S_{p1} \) and \( S_{p2} \) are dealt with in Lemma B.5 below. The last four terms on the right hand side of (B.24) and \( \Pi_{pq} \) are summarized in Lemma B.3. The first three terms on the right hand side of (B.24) and \( S_{p3} \) are dealt with in Lemma B.4. Given these results, we have

\[ \mathcal{P}_{p1}(\hat{\beta}_1 - \beta_1) + \mathcal{P}_{p2}(\hat{\beta}_2 - \beta_2) + \cdots + \mathcal{P}_{pk}(\hat{\beta}_k - \beta_k) \]
Under Assumptions A-E, together with the identification conditions IZ, we have

\[
\begin{align*}
    \mathcal{P}(\hat{\beta} - \beta) &= \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} e_{it}v_{itp} + \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} \beta_{h} h_{it} e_{it} \\
    &- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} \psi_{i} \Pi_{iit}^{-1} \phi_{i} h_{it} e_{it} + O_{p}(T^{-1}) + O_{p}(N^{-1}T^{-1/2})
\end{align*}
\]

for any \( p = 1, 2, \ldots, K \). The above result is equivalent to

\[
(B.28) \quad \mathcal{P}(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} e_{it}v_{itx} + \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} \beta_{x} h_{it} e_{it} \\
- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} \psi_{i} \Pi_{iit}^{-1} \phi_{i} h_{it} e_{it} + O_{p}(T^{-1}) + O_{p}(N^{-1}T^{-1/2})
\]

which implies that \( \hat{\beta} - \beta = O_{p}(N^{-1/2}T^{-1/2}) + O_{p}(T^{-1}) \). This completes the proof of Proposition B.3.

\[ \square \]

**Corollary B.2.** Under Assumptions A-E, together with the identification conditions IZ, we have

\[
\begin{align*}
    \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{iit}^{-1}|| \cdot ||\hat{\Gamma}_{i} - \Gamma_{i}||^2 &= O_{p}(T^{-1}) \\
    \frac{1}{N} \sum_{i=1}^{N} ||\hat{\Sigma}_{iit} - \Sigma_{iit}||^2 &= O_{p}(T^{-1})
\end{align*}
\]

This corollary is a direct result of Propositions B.2 and B.3.

**Lemma B.5.** Under Assumptions A-E, together with the identification conditions IZ, we have

\[ (a) \quad S_{p1} = O_{p}(N^{-1}T^{-1/2}) + O_{p}(T^{-3/2}) + o_{p}(||\hat{\beta} - \beta||) \\
(b) \quad S_{p2} = o_{p}(||\hat{\beta} - \beta||) \]

where \( S_{p1} \) and \( S_{p2} \) are defined in (B.25) and (B.26), respectively.

**Proof of Lemma B.5.** Using Proposition B.2, we can prove, just like Lemma A.11, all the terms in \( S_{p1} \) are \( O_{p}(N^{-1}T^{-1/2}) + O_{p}(T^{-3/2}) + o_{p}(||\hat{\beta} - \beta||) \). It is worth pointing out that, to prove \( \text{tr}[\frac{1}{N} \Sigma_{i=1}^{N} \Sigma_{iit}^{-1} \phi_{i} \varphi_{i} \hat{\lambda}_{i} - \lambda_{i} A'] \) and \( \text{tr}[\frac{1}{N} \Sigma_{i=1}^{N} \Sigma_{iit}^{-1} \phi_{i} \varphi_{i} A' A] \) to be \( O_{p}(N^{-1}T^{-1/2}) + O_{p}(T^{-3/2}) + o_{p}(||\hat{\beta} - \beta||) \), we need to strengthen Proposition B.2 to \( A \equiv \Sigma_{i=1}^{N} (\hat{\Gamma}_{i} - \Gamma_{i}) \Sigma_{iit}^{-1} \hat{\Gamma}_{i} \hat{H} = \)
O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(∥\hat{\beta} - \beta∥). Since our identification conditions are similar as IC3 in [10], this result can be proved by the same way as in their paper. We omit the details.

Result (b) is easier to prove. The details are omitted. □

APPENDIX B3: PROOF OF THEOREM 3.1 AND ITS ALTERNATIVE EXPRESSION

Note that (B.28) is close to Proposition 3.1 except that the remainder term $O_p(T^{-1}) + O_p(N^{-1}T^{-1/2})$ needs to be strengthened to $O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2})$. The strengthened results are stated in the following lemma.

**Lemma B.6.** Under Assumptions A-E, together with the identification conditions IZ, we have

(a) $\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iit}^{-1} h_t e_{it}$

\[= \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} h_t e_{it} + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(∥\hat{\beta} - \beta∥)\]

(b) $\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iit}^{-1} e_{it} v_{itp}$

\[= \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iit}^{-1} e_{it} v_{itp} + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(∥\hat{\beta} - \beta∥)\]

(c) $\hat{G}_2 \hat{\Sigma}_e^{-1} \hat{\Gamma} \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t e_{jt} \hat{\Sigma}_{jj}^{-1} \hat{\psi}_j^{-1} \hat{\Pi}_{\psi_j}^{-1} \hat{\beta}_p = \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} h_t e_{jt} \Sigma_{jj}^{-1} \psi_j^{-1} \Pi_{\psi_j}^{-1} \hat{\beta}_p$

\[+ O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(∥\hat{\beta} - \beta∥)\]

(d) $\frac{1}{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\gamma}_t \hat{\Sigma}_{jj}^{-1} \frac{1}{T} \sum_{t=1}^{T} (\varepsilon_{jt} v_{itp} - E(\varepsilon_{jt} v_{itp})) \hat{\Sigma}_{iit}^{-1} \hat{\psi}_t \hat{\beta}$

\[= O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(∥\hat{\beta} - \beta∥)\]

(e) $\hat{G}_1 \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\gamma}_t \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} (\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})) \hat{\Sigma}_{jj}^{-1} \hat{\psi}_t \hat{\beta}$

\[= O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(∥\hat{\beta} - \beta∥)\]

(f) $H \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\gamma}_t \hat{\Sigma}_{ii}^{-1} (\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})) \hat{\Sigma}_{jj}^{-1} \hat{\psi}_t \hat{\beta}$

\[= O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(∥\hat{\beta} - \beta∥)\]

\[\hat{\beta}^{-1} \hat{\Sigma}_{ii}^{-1} (\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})) \hat{\Sigma}_{jj}^{-1} \hat{\psi}_t \hat{\beta} = O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(∥\hat{\beta} - \beta∥)\]
Similarly we can rewrite the first equation as

\[ (g) \quad \frac{1}{NT} \hat{G}_2 \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{I}_{ij} \Sigma_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} (\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})) \hat{\Gamma}^{-1} \psi_i \beta_p = O_p(N^{-1/2}T^{-1}) + O_p(N^{-1}T^{-1}) + O_p(T^{-3/2}) + o_p(\|\hat{\beta} - \beta\|) \]

**Proof of Lemma B.6.** The whole proof of Lemma B.6 is very similar to that of Lemma A.13, we omit the proof to avoid repetition.

**Proof of Proposition 3.1.** Given (B.24) and Lemmas B.3, B.5 and B.6, the proof of Proposition 3.1 is almost the same as that of Theorem 2.2.

**Proof of Theorem 3.1.** To prove Theorem 3.1, we first show that for any \( \Gamma \) and \( f_1 \) we can always transform them into new \( \Gamma^* \) and \( f_1^* \), which satisfy the identification conditions IZ. Consider Model (3.1), which we write out below for ease of reading:

\[
y_{it} = \alpha_i + x_{it1} \beta_1 + x_{it2} \beta_2 + \cdots + x_{itK} \beta_K + \psi_i g_t + \epsilon_{it} \\
x_{itk} = \mu_{ik} + \gamma_{ik} \gamma_{ik} + \gamma_{ik} h_t + v_{itk}
\]

Now we rewrite the second equation of the above model as the following way

\[
x_{itk} = (\mu_{ik} + \gamma_{ik} \gamma_{ik} + \gamma_{ik} h_t) + (\gamma_{ik} \gamma_{ik} + \gamma_{ik} (\hat{\Gamma}^T \hat{\Gamma})^{-1} \hat{\Gamma}^T) \gamma_{ik} h_t + v_{itk}
\]

\[
= (\mu_{ik} + \gamma_{ik} \gamma_{ik} + \gamma_{ik} (\hat{\Gamma}^T \hat{\Gamma})^{-1} \hat{\Gamma}^T) \gamma_{ik} h_t + v_{itk}
\]

\[
= (\mu_{ik} + \gamma_{ik} \gamma_{ik} + \gamma_{ik} (\hat{\Gamma}^T \hat{\Gamma})^{-1} \hat{\Gamma}^T) \gamma_{ik} h_t + v_{itk}
\]

\[ \times \left\{ Q_1' (\hat{\Gamma}^T \hat{\Gamma})^{-1/2} g_t, + \left\{ \gamma_{ik} (\hat{\Gamma}^T \hat{\Gamma})^{-1} \hat{\Gamma}^T \right\}^{1/2} Q_2 \right\}
\]

\[ \times \left\{ Q_2 \left( \hat{\Gamma}^T \hat{\Gamma}^T - (\hat{\Gamma}^T \hat{\Gamma})^{-1} (\hat{\Gamma}^T \hat{\Gamma})^{-1} \right)^{-1/2} (\hat{\Gamma}^T \hat{\Gamma}^T - (\hat{\Gamma}^T \hat{\Gamma})^{-1} (\hat{\Gamma}^T \hat{\Gamma})^{-1} \gamma_{ik} h_t) \right\}
\]

where \( Q_1 \) is an \( r_1 \times r_1 \) and \( Q_2 \) is an \( r_2 \times r_2 \) orthogonal matrix defined below. Similarly we can rewrite the first equation as

\[ y_{it} = \alpha_i + \sum_{k=1}^{K} x_{itk} \beta_k + \psi_i g_t + \epsilon_{it} \]

\[ = (\alpha_i + \psi_i g_t) + \sum_{k=1}^{K} x_{itk} \beta_k + \psi_i g_t + \epsilon_{it} \]

\[ = (\alpha_i + \psi_i g_t) + \sum_{k=1}^{K} x_{itk} \beta_k + \left\{ \psi_i (\hat{\Gamma}^T \hat{\Gamma})^{1/2} Q_1 \right\} \left\{ Q_1' (\hat{\Gamma}^T \hat{\Gamma})^{-1/2} g_t \right\} \]
We first introduce $\Gamma$ eigenvalues in descending order and $Q$.

Now consider term (B.34)

Then the second equation of (3.1) can be written as Theorem 3.1 state.

Let $\Gamma^g, \Gamma^h, \Gamma$ be defined similarly as $\Gamma^g, \Gamma^h, \Gamma$ in the main text. If matrix $Q_1$ is chosen to be the eigenvector matrix of $\Gamma^{g+}\Sigma_{\varepsilon^{-1}}\Gamma^g$ with the associated eigenvalues in descending order and $Q_2$ to be the eigenvector matrix of $\Gamma^{h+}\Sigma_{\varepsilon^{-1}}\Gamma^h$, we can easily verify $\Gamma^*$ and $f^*_t = (g^*_t, h^*_t)'$ satisfy IZ.

Since $(\Gamma^*, f^*_t, \beta)$ satisfy IZ, by Proposition 3.1, we have

$$P^*(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{\varepsilon t} e_{it} v_{itx} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{\varepsilon t} h^*_t e_{it}$$

$$- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{\varepsilon t} \psi \psi^{-1}
\left(\frac{1}{N} \sum_{j=1}^{N} \psi_j \Sigma^{-1}_{\varepsilon t} h^*_t e_{it}\right)$$

$$+ O_p(T^{-3/2}) + O_p(N^{-1}T^{-1/2}) + O_p(N^{-1/2}T^{-1})$$

Substituting (B.29)—(B.33) into the above expression, we obtain the result as Theorem 3.1 state.

**Deriving the alternative expression of $\hat{\beta} - \beta$.** We first introduce the following notations for ease of exposition. Let

$$\Gamma^g_p = \begin{pmatrix} \mu_{1p} & \gamma^g_{3p} \\ \mu_{2p} & \gamma^g_{2p} \\ \vdots & \vdots \\ \mu_{Np} & \gamma^g_{Np} \end{pmatrix} \quad \Gamma^h_p = \begin{pmatrix} \gamma^h_{1p} \\ \gamma^h_{2p} \\ \vdots \\ \gamma^h_{Np} \end{pmatrix} \quad H = \begin{pmatrix} h_1' \\ h_2' \\ \vdots \\ h_T' \end{pmatrix}$$

Then the second equation of (3.1) can be written as

(B.34) $$X_p = \Gamma^g_p \psi' + \Gamma^h_p H' + \nu_p, \quad p = 1, 2, \ldots, K$$

Now consider term $\frac{1}{NT} \text{tr}(\bar{M} X_q M(\bar{G}) X_p')$, where $p, q = 1, 2, \ldots, K$. By (B.34), this term is equal to

$$\frac{1}{NT} \text{tr}(\bar{M} \Gamma^h_p H' M(\bar{G}) H_p') + \frac{1}{NT} \text{tr}(\bar{M} \Gamma^g_p \psi' M(\bar{G}) \psi_p')$$
By the similar way to deal with $\frac{1}{NT} \text{tr}[\mu_{q}H'\mathcal{M}(\overline{\mathcal{G}})\Gamma_{p}]$ in Theorem 2.2, we can prove the last term of (B.35) is equal to $\frac{1}{N} \sum_{i=1}^{N} \Sigma_{i}^{-1} \Sigma_{iix} + o_p(1)$. Consider the second term of (B.35), which is equal to

$$
\frac{1}{NT} \text{tr}\left[ \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ite}^{-1} h'^{t} v_{iq} \right] - \text{tr}\left[ \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ite}^{-1} h'^{t} v_{iq} \right) (\overline{\mathcal{G}}^{-1}) \mathcal{G} \mathcal{H} \right]
$$

The four terms of the above expression are all $O_p(N^{-1/2}T^{-1/2})$ by $(\Gamma_{p}^{-1}) \mathcal{H}$, $O(1)$ and $(\overline{\mathcal{G}}^{-1}) \mathcal{H} = O(1)$. So the second term of (B.35) is $O_p(N^{-1/2}T^{-1/2})$. The third term can be proved to be $O_p(N^{-1/2}T^{-1/2})$ similarly as the second term. Given these results, we have

$$
\frac{1}{NT} \text{tr}[\mu_{q}H'\mathcal{M}(\overline{\mathcal{G}})X_{p}'] = \frac{1}{NT} \text{tr}[\mu_{q}H'\mathcal{M}(\overline{\mathcal{G}})\Gamma_{p}] + \frac{1}{N} \sum_{i=1}^{N} \Sigma_{i}^{-1} \Sigma_{iix} + o_p(1)
$$

Then it follows

$$
\frac{1}{NT} \begin{pmatrix}
\text{tr}[\mu_{X_{1}}\mathcal{M}(\overline{\mathcal{G}})X_{1}'] & \cdots & \text{tr}[\mu_{X_{1}}\mathcal{M}(\overline{\mathcal{G}})X_{K}'] \\
\vdots & \ddots & \vdots \\
\text{tr}[\mu_{X_{K}}\mathcal{M}(\overline{\mathcal{G}})X_{1}'] & \cdots & \text{tr}[\mu_{X_{K}}\mathcal{M}(\overline{\mathcal{G}})X_{K}']
\end{pmatrix} = \mathcal{P} + o_p(1)
$$

Now we turn attention to $\frac{1}{NT} \text{tr}[\mu_{q}H'\mathcal{M}(\overline{\mathcal{G}})e']$, which is equal to

$$
\frac{1}{NT} \text{tr}[\mu_{q}H'\mathcal{M}(\overline{\mathcal{G}})e'] + \frac{1}{NT} \text{tr}[\mu_{p}H'\mathcal{M}(\overline{\mathcal{G}})e']
$$

The first term of the above expression is equal to

$$
\text{tr}\left[ \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ite}^{-1} h'^{t} e_{it} \right] - \text{tr}\left[ \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ite}^{-1} h'^{t} e_{it} \right) (\overline{\mathcal{G}}^{-1}) \mathcal{G} \mathcal{H} \right]
$$

$$
- \text{tr}\left[ \left( \Gamma_{p}^{-1} \Sigma_{iix} \right) (\Psi' \Sigma_{iix}^{-1} \Psi)^{-1} \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{ite}^{-1} h'^{t} e_{it} \right) (\overline{\mathcal{G}}^{-1}) \mathcal{G} \mathcal{H} \right]
$$
Using \( h_t - \frac{\eta}{(\bar{G}^T \bar{G})^{-1} \bar{G}' \bar{H}} \), we can rewrite the above expression as

\[
\text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \frac{h_t}{\gamma_{ip}} \hat{h}_t e_{it} \right] - \text{tr} \left[ \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \frac{h_t}{\gamma_{ip}} \hat{g}_te_{it} \right) (\bar{G}' \bar{G})^{-1} \bar{G}' \bar{H} \right]
\]

\[
- \text{tr} \left[ \left( \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \frac{h_t}{\gamma_{ip}} \hat{g}_te_{it} \right) (\bar{G}' \bar{G})^{-1} \bar{G}' \bar{H} \right]
\]

which is equivalent to

\[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \frac{h_t}{\gamma_{ip}} (\hat{h}_t - \bar{H}' \bar{G}(\bar{G}' \bar{G})^{-1} \bar{g}_t)e_{it}
\]

\[
- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \frac{h_t}{\gamma_{ip}} (\hat{h}_t - \bar{H}' \bar{G}(\bar{G}' \bar{G})^{-1} \bar{g}_t)e_{it}
\]

From this result and

\[
\frac{1}{NT} \text{tr} \left[ \hat{M}' \hat{p} \hat{M}(\bar{G}) e' \right] = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} e_{it} v_{itp} + O_p(N^{-1/2}T^{-1}) + O_p(N^{-1/2}T^{-1})
\]

which has been shown in Appendix A.3, we have

(B.38) \[
\text{tr} \left[ \frac{1}{NT} \hat{M} X_p \hat{M}(\bar{G}) e' \right] = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} e_{it} v_{itp}
\]

\[
+ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \frac{h_t}{\gamma_{ip}} (\hat{h}_t - \bar{H}' \bar{G}(\bar{G}' \bar{G})^{-1} \bar{g}_t)e_{it}
\]

\[
- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{iie} \frac{h_t}{\gamma_{ip}} (\hat{h}_t - \bar{H}' \bar{G}(\bar{G}' \bar{G})^{-1} \bar{g}_t)e_{it}
\]

+ \text{O}_p(N^{-1/2}T^{-1}) + O_p(N^{-1/2}T^{-1})

Combining (B.36) and (B.38), we can see that the alternative asymptotic expression of \( \hat{\beta} - \beta \) is equivalent to the one in Theorem 3.1. In addition, Corollary 3.1 is an immediate result of (B.36).
APPENDIX C: PROOFS OF RESULTS FOR MODELS WITH TIME-INVARIANT AND COMMON REGRESSORS

APPENDIX C1: PROOF OF CONSISTENCY

Again, for consistency, we use the superscript \( \ast \) to denote the true parameters.

**Proposition C.1.** Let \( \hat{\theta} = (\hat{\beta}, \hat{\Gamma}, \Sigma_{ee}, \hat{M}_{ff}) \) be the solution by maximizing (4.2). Under Assumptions A-D plus F, together with the identification conditions IO, when \( N, T \to \infty \), we have

\[
\hat{\beta} - \beta^\ast \overset{p}{\to} 0
\]

\[
\frac{1}{N} \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i^\ast) \Sigma_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i^\ast)' \overset{p}{\to} 0
\]

\[
\frac{1}{N} \sum_{i=1}^{N} \|\Sigma_{ii} - \Sigma_i^\ast\|^2 \overset{p}{\to} 0
\]

\[
\|\hat{M}_{ff} - M_{ff}^\ast\| \overset{p}{\to} 0
\]

**Proof of Proposition C.1.** First note that \( \hat{\beta} - \beta^\ast \overset{p}{\to} 0 \) and \( \frac{1}{N} \sum_{i=1}^{N} \|\Sigma_{ii} - \Sigma_i^\ast\|^2 \overset{p}{\to} 0 \) continue to hold as they do not rely on priori restrictions. To prove the remaining results, it is sufficient to prove

\[
A = (\hat{\Gamma}_i - \Gamma_i^\ast)' \Sigma_{ii}^{-1} \hat{\lambda}_i
\]

Equation (4.6) can be written as

\[
\hat{G}_2 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i^\ast \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t^* e_{jt} \hat{\Sigma}_{ge}^{-1} \hat{\lambda}_j + \hat{G}_2 N \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t f_t^* \hat{\Pi}_{\lambda}\lambda
\]

\[
+ \hat{G}_2 \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [e_{it} e_{jt} - E(e_{it} e_{jt})] \hat{\Sigma}_{ge}^{-1} \hat{\lambda}_j - \hat{G}_2 N \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^\ast)' f_t^* \hat{\Pi}_{\lambda}\lambda
\]

\[
- \hat{G}_2 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^\ast)' \hat{F}_t \right) - \hat{G}_2 \hat{\Gamma} \hat{\Sigma}_{ee}^{-1} \Gamma^\ast \left( \frac{1}{T} \sum_{t=1}^{T} f_t^*(\hat{\beta} - \beta^\ast)' \hat{F}_t \right)
\]

\[
- \hat{G}_2 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^\ast) e_{jt} \hat{\Sigma}_{ge}^{-1} \hat{\lambda}_j + \hat{G}_2 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta^\ast)' (\hat{\beta} - \beta^\ast)' \right) \hat{F}_t \right)
\]

\[
- \frac{1}{N} \hat{G}_2 \sum_{i=1}^{N} \hat{\lambda}_i \hat{\Sigma}_{ii}^{-1} (\Sigma_{ii} - \Sigma_i^\ast) \hat{\Sigma}_{ii}^{-1} \hat{\lambda}_i + \hat{G}_2 N \sum_{j=1}^{N} (\hat{\lambda}_j - \lambda_j^\ast) \hat{\Sigma}_{ge}^{-1} \hat{\lambda}_j
\]
where $\hat{\Pi}_{\lambda\lambda}$, $\hat{\xi}_t$ and $\hat{F}_t$ are defined in Tables 1 and 2. The last term involves $M_{ff} - M'_{ff}$. Notice (2.7) continues to hold, so (A.21) is applicable. By (A.21), we can rewrite (C.1) as

$$-\hat{G}_2 \sum_{i=1}^{N} \hat{\gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i') M_{ij}' \hat{\Pi}_{ll} - \hat{G}_2 \hat{H}^{-1} (M_{ff} - M'_{ff}) \hat{\Pi}_{ll} = 0$$

(C.2) $$-\hat{G}_2 \sum_{i=1}^{N} \hat{\gamma}_i \hat{\Sigma}_{ii}^{-1} \Gamma_t' \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_t e_{jt} \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_{jt} + \hat{G}_2 \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t f_t' \right) \hat{\Pi}_{ll}$$

$$+ \hat{G}_2 \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{ut} e_{jt} - E(\varepsilon_{ut} e_{jt})] \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_{jt} - I_2 \hat{A} M_{ff} \hat{A} \hat{\Pi}_{ll}$$

$$- \hat{G}_2 \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta)^' F_t \right) - \hat{G}_2 \hat{\Gamma} \hat{\Sigma}_{ee}^{-1} \Gamma^* \left( \frac{1}{T} \sum_{t=1}^{T} f_t' (\hat{\beta} - \beta)^' F_t \right)$$

$$- \hat{G}_2 \frac{1}{T} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\xi}_t (\hat{\beta} - \beta) e_{jt} \hat{\Sigma}_{jj}^{-1} \hat{\lambda}_{jt} + \hat{G}_2 \frac{1}{T} \sum_{j=1}^{N} \hat{\xi}_t (\hat{\beta} - \beta)^' \hat{F}_t$$

$$- \frac{1}{N} \hat{G}_2 \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\xi}_t (\hat{\beta} - \beta)^' \hat{F}_t$$

where $\hat{\lambda} = (\hat{\Gamma} - \Gamma^*) \hat{\Sigma}_{ee}^{-1} \hat{G}$ and $A = (\hat{\Gamma} - \Gamma^*) \hat{\Sigma}_{ee}^{-1} \hat{H}$. There are 22 terms on the left hand side of (C.2). We can prove that all the terms except the 11th
and 22th terms are $o_p(1)$, given the fact that $\hat{\beta} - \beta^* \overset{p}{\to} 0$ and $\frac{1}{N} \sum_{i=1}^N \|\hat{\Sigma}_{ii} - \Sigma_{ii}^*\|^2 \overset{p}{\to} 0$, whose proofs, as pointed out in Appendix B, involve no priori restrictions and still hold in present context. By the definitions of $A$ and $A^*$, the 11th term is equal to

$$
\hat{G}_2^\gamma \hat{\Sigma}_{ee}^{-1}(\hat{\Gamma} - \Gamma^*) M_{ff}^* A \hat{\Pi}_{\lambda \lambda} = \hat{H}_2^\gamma \hat{\Sigma}_{ee}^{-1}(\hat{\Gamma} - \Gamma^*) M_{ff}^* A \hat{\Pi}_{\lambda \lambda}
$$

where we have used $\hat{G} = \hat{H} - \hat{H} \hat{M}_{ff}^{-1} \hat{G}$. Since (A.21) continues to hold in Section 4, we have $A = O_p(1)$. By (A.16), we have $\frac{1}{N} \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} = O_p(1)$, it is easy to verify that $\hat{\Pi}_{\lambda \lambda} = \frac{1}{N} \sum_{j=1}^N \lambda^*_j \hat{\Sigma}_{ee}^{-1} \hat{\lambda}_j = O_p(1)$. Given $\hat{\Pi}_{\lambda \lambda} = O_p(1)$ and $A = O_p(1)$, together with $\hat{G} \hat{H}^{-1} = O_p(1)$, $M_{ff}^{-1} = O_p(1)$, $\hat{H} = o_p(1)$, we have

$$
I_2(I_r - A)' M_{ff}^* A \hat{\Pi}_{\lambda \lambda} \overset{p}{\to} 0
$$

However, term $\hat{\Pi}_{\lambda \lambda} = \frac{1}{N} \sum_{j=1}^N \lambda^*_j \hat{\Sigma}_{ee}^{-1} \hat{\lambda}_j$ can be proved to be invertible similarly as $\frac{1}{N} \sum_{j=1}^N \psi^*_j \hat{\Sigma}_{ee}^{-1} \psi_j$ in Appendix B. So we have

$$
I_2(I_r - A)' M_{ff}^* A \overset{p}{\to} 0
$$

From this result, in combination with $\hat{M}_{ff} = (I_r - A)'M_{ff}^*(I_r - A) + o_p(1)$, we obtain $A_{12} = o_p(1)$ and $A_{22} = o_p(1)$. These results, in combination with $\hat{M}_{ff} = (I_r - A)'M_{ff}^*(I_r - A) + o_p(1)$ again and $IO2$, we have $A_{21} = o_p(1)$ and $A_{11} = o_p(1)$. The remaining proof is the same as that of Proposition 2.1. This completes the proof of Proposition C.1. \hfill \Box

APPENDIX C2: PROOFS OF THE CONVERGENCE RATES AND
PROPOSITION 4.1

Now we drop "\***" from the true value of the parameters for notational simplicity. The following lemma is useful for deriving the rates of convergence.

**Lemma C.1.** Under Assumptions A-D plus $G$, in combination with the identification conditions $IO$, we have

$$
\sum_{i=1}^N(\hat{\Gamma}_i - \Gamma_i) \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i \hat{H} = O_p(T^{-1/2}) + O_p\left(\frac{1}{N} \sum_{i=1}^N \|\hat{\Sigma}_{ii} - \Sigma_{ii}\|^2\right)^{1/2} + O_p\left(\frac{1}{N} \sum_{i=1}^N \|\hat{\Sigma}_{ii}\| \cdot \|\hat{\Gamma}_i - \Gamma_i\|^2\right) + O_p(\|\hat{\beta} - \beta\|)
$$

Proof of Lemma C.1. The proof of Lemma C.1 is similar to that of Lemma B.2. We partition matrix $A$ into $A_{11}, A_{12}, A_{21}, A_{22}$ and prove each submatrix is $O_p(T^{-1/2}) + O_p((\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii}^2 \|^{1/2}) + O_p((\frac{1}{N} \sum_{i=1}^{N} \| \Sigma^{-1}_{ii} \| \cdot \| \hat{\Gamma}_i - \hat{\Gamma}_i \|)^2) + O_p(\| \hat{\beta} - \beta \|)$. The details are omitted.

Proposition C.2. Under Assumptions A-D plus $G$, together with the identification conditions IO, we have

$$\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} \| \cdot \| \hat{\Gamma}_i - \hat{\Gamma}_i \|^2 = O_p(T^{-1}) + O_p(\| \hat{\beta} - \beta \|^2)$$

$$\frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 = O_p(T^{-1}) + O_p(\| \hat{\beta} - \beta \|^2)$$

$$\| M_{ff} - M_{ff} \|^2 = O_p(T^{-1}) + O_p(\| \hat{\beta} - \beta \|^2)$$

Proof of Proposition C.2. The first order condition with respect to $\psi_j$ is identical to the one in the last section. By (B.4), we have

$$\hat{\psi}_j - \psi_j = -\hat{G}_1 \sum_{i=1}^{N} (\hat{\Gamma}_i - \Gamma_i)' M_{ff} \lambda_j + \hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i) M_{ff} \lambda_j + \frac{1}{T} \sum_{t=1}^{T} f_t e_{jt}$$

$$+ \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{\Gamma}_i M_{ff} \lambda_j \right) - \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{\Gamma}_i M_{ff} \lambda_j \right)$$

$$- \hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i \frac{1}{T} \sum_{t=1}^{T} f_t e_{jt} - \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{\Gamma}_i M_{ff} \lambda_j \right)$$

(C.3)

$$+ \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{\Gamma}_i M_{ff} \lambda_j \right) - \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\hat{\beta} - \beta)' \hat{\Gamma}_i M_{ff} \lambda_j \right)$$

The remaining proof is the same as that of Proposition B.2 and the details are hence omitted.

Proposition C.3. Under Assumption A-D plus $G$, we have

$$\hat{\beta} - \beta = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1})$$
To prove Proposition C.3, we need the following lemmas.

**Lemma C.2.** Under Assumptions A-D plus G, together with the identification conditions IO, we have

(a) \[ \text{tr}\left[ G_1 \sum_{i=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \hat{\Gamma}_i' \hat{\Sigma}_{ii} \hat{\Gamma}_i' \hat{\Sigma}_{ii} \hat{\Gamma}_i' \hat{\Sigma}_{ii} \right] \]

\[ = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}[I_1 M_{ff} \omega_{qp} I_1'] + o_p(||\hat{\beta} - \beta||) \]

(b) \[ \text{tr}\left[ I_1 (I - A)' \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{\xi}_t' \hat{H}_N \hat{\nu}_{\hat{p}}' \right] \]

\[ = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}[I_1 M_{ff} \nu_{q} Q^{-1} \nu_{p} I_1'] + o_p(||\hat{\beta} - \beta||) \]

(c) \[ \text{tr}\left[ \hat{G}_2 \hat{\Gamma} \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{F}_t \Pi_{\lambda \lambda}^{-1} \hat{\nu}_{\hat{p}}' \right] \]

\[ = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}[I_2 M_{ff} \nu_{q} Q^{-1} \nu_{p} I_2'] + o_p(||\hat{\beta} - \beta||) \]

(d) \[ \text{tr}\left[ \hat{G}_2 \hat{\Gamma} \hat{\Sigma}_{ee}^{-1} \hat{\Gamma} \frac{1}{T} \sum_{t=1}^{T} f_t (\hat{\beta} - \beta)' \hat{F}_t \Pi_{\lambda \lambda}^{-1} \hat{\nu}_{\hat{p}}' \right] \]

\[ = \sum_{q=1}^{K} (\hat{\beta}_q - \beta_q) \text{tr}[I_2 M_{ff} \nu_{q} Q^{-1} \nu_{p} I_2'] + o_p(||\hat{\beta} - \beta||) \]

where \( \Pi_{\lambda \lambda}, \Pi_{\lambda \lambda}^{-1}, \hat{\nu}_{\hat{p}}, \hat{\nu}_{\hat{p}}, \nu_{p}, \omega_{pq}, G_{tp}, \hat{\xi}_t \) are defined in Tables 1, 2 and 3. \( I_1 \) denotes the first \( r_1 \) rows and \( I_2 \) denotes the remaining \( r_2 \) rows of the identity matrix \( I_r \).

**Lemma C.3.** Under Assumptions A-D plus G, together with the identification conditions IO, we have

(a) \[ S_{p1} = o_p(N^{-1/2}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(b) \[ S_{p2} = o_p(||\hat{\beta} - \beta||) \]

where \( S_{p1} \) and \( S_{p2} \) are defined in (C.6) below.
Lemma C.4. Under Assumptions A-D plus G, together with the identification conditions IO, we have

\[(a)\] \[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iie}^{-1} h_{it} e_{it} = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

\[(b)\] \[
\frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iit}^{-1} e_{it} v_{itp} = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

\[(c)\] \[
\frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} h_{it} e_{jt} \hat{\Sigma}_{je}^{-1} \underbrace{\hat{\lambda}_j^{1} \hat{\Pi}_{\lambda\lambda}^{-1} \hat{\psi}_p^{1}} = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

\[(d)\] \[
\text{tr}\left[\hat{G}_1 \frac{1}{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_t \hat{\Sigma}_{ii}^{-1} \left( 1 \sum_{j=1}^{T} [\varepsilon_{it} \varepsilon_{jt} - E(\varepsilon_{it} \varepsilon_{jt})] \hat{\Sigma}_{je}^{-1} g_j' \right)\right]
= O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

\[(e)\] \[
\text{tr}\left[\frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_t \hat{\Sigma}_{ii}^{-1} \left( [\varepsilon_{jt} v_{itp} - E(\varepsilon_{jt} v_{itp})] \hat{\Sigma}_{iie}^{-1} \hat{\psi}_j \hat{G} \right) \right]
= O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

\[(f)\] \[
\text{tr}\left[\hat{H}_1 \frac{1}{T} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_t \hat{\Sigma}_{ii}^{-1} \left( [\varepsilon_{jt} \varepsilon_{jt} - E(\varepsilon_{jt} \varepsilon_{jt})] \hat{\Sigma}_{je}^{-1} \hat{G}_j' \hat{H}_p^{-1} \hat{h}_p \right) \right]
= O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

\[(g)\] \[
\text{tr}\left[\hat{G}_2 \frac{1}{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_t \hat{\Sigma}_{ii}^{-1} \left( [\varepsilon_{it} \varepsilon_{jt} - E(\varepsilon_{it} \varepsilon_{jt})] \hat{\Sigma}_{je}^{-1} \hat{\lambda}_j^{1} \hat{\Pi}_{\lambda\lambda}^{-1} \hat{\psi}_p^{1} \right) \right]
= O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

\[(h)\] \[
\text{tr}\left[\hat{G}_2 \frac{1}{T} \sum_{j=1}^{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_t \hat{\Sigma}_{ii}^{-1} \left( [\varepsilon_{it} \varepsilon_{jt} - E(\varepsilon_{it} \varepsilon_{jt})] \hat{\Sigma}_{je}^{-1} \hat{G}_j' \hat{H}_p^{-1} \hat{h}_p \right) \right]
= O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) + o_p(||\hat{\beta} - \beta||)
\]

The proofs of Lemmas C.2, C.3 and C.4 are quite similar as the counterparts in Appendix B.

Proof of Proposition C.3. As argued in Section 4, equation (2.5) still holds. So equation (A.41) continues to hold. Consider the first term on the right hand side of (A.41). Notice that \(-\text{tr}[\frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{iie}^{-1} (\hat{\psi}_i - \psi_i)' M_{il}'] = -\text{tr}[\frac{1}{N} \sum_{i=1}^{N} \gamma_{ip} \hat{\Sigma}_{iie}^{-1} (\hat{\psi}_i - \psi_i)']\). Using (C.3) to replace \(\hat{\psi}_i - \psi_i\) from the ex-
pression, we have
\[
\vartheta_{p1}(\beta_1 - \beta_1) + \vartheta_{p2}(\beta_2 - \beta_2) + \cdots + \vartheta_{pK}(\beta_K - \beta_K)
\]
\[
= \text{tr}\left[ \hat{G}_1 \sum_{i=1}^{N} \hat{f}_i \hat{\Sigma}_i^{-1} (\hat{f}_i - \Gamma_i)' M_{i} f_i \hat{\gamma}_p \right] - \text{tr}\left[ \hat{G}_1 \hat{f}_i \hat{\Sigma}_i^{-1} \Gamma_i \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_i e_{jt} \hat{\Sigma}_i^{-1} \hat{\gamma}_j \right] 
\]
\[
- \text{tr}\left[ \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_i \hat{\gamma}_p \right) \right] + \text{tr}\left[ \hat{G}_1 \sum_{i=1}^{N} \hat{f}_i \hat{\Sigma}_i^{-1} \Gamma_i \frac{1}{T} \sum_{t=1}^{T} f_i (\beta - \beta)' \hat{G}_i \right] 
\]
\[
+ \text{tr}\left[ \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta)' \hat{G}_i \right) \right] + \text{tr}\left[ \hat{G}_1 N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta)' \hat{G}_i \right) \right] 
\]
\[
+ \text{tr}\left[ \frac{1}{N} \hat{G}_1 \sum_{j=1}^{N} \hat{\lambda}_j \hat{\Sigma}_j^{-1} \left( \hat{\Sigma}_j - \hat{\Sigma}_j \hat{\gamma}_j \right) \right] + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1} \hat{\lambda}_i f_i v_{it} \epsilon_{it} 
\]
\[
+ \text{tr}\left[ \hat{H}_1 M_{i} f_i \hat{G}_i \frac{1}{N} \sum_{j=1}^{N} \hat{\lambda}_j \hat{\Sigma}_j^{-1} \hat{\gamma}_j \right] + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1} \hat{\gamma}_i \hat{G}_i \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_i \right] 
\]
\[
- \text{tr}\left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_i^{-1} \hat{\chi}_t v_{it} \hat{G}_i \hat{\Sigma}_i^{-1} \Gamma_i \right] - \text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_i^{-1} \hat{\gamma}_i \hat{G}_i \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_i \right] 
\]
\[
- \text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_i^{-1} \left( \hat{\gamma}_i - \hat{\gamma}_i \right) \hat{G}_i \right] + \text{tr}\left[ \frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_i^{-1} \left( \hat{\gamma}_i - \hat{\gamma}_i \right) P_{i}^{+} \hat{\Sigma}_i^{-1} \hat{\gamma}_i \right] 
\]
\[
- \text{tr}\left[ \hat{G}_1 \frac{1}{N} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\chi}_t \hat{\Sigma}_i^{-1} \left( e_{jt} - (e_{jt} e_{jt}) \hat{\Sigma}_i^{-1} \hat{\gamma}_j \right) \right] 
\]
\[
- \text{tr}\left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\chi}_t \hat{\Sigma}_j^{-1} [e_{jt} v_{it} - (e_{jt} v_{it})] \hat{\Sigma}_j^{-1} \hat{G}_j \right] 
\]
\[
(C.4) - \text{tr}\left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\chi}_t \hat{\Sigma}_j^{-1} [e_{jt} v_{it} - (e_{jt} v_{it})] \hat{\Sigma}_j^{-1} \hat{G}_j \right]
\]
The right hand side of the above expression has 22 terms. We pick out the first term \( \text{tr}[\hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \Sigma_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i) \hat{M}_{ff} \hat{v}_p^g] \) and the 15th term \( \text{tr}[\frac{1}{N} \sum_{i=1}^{N} \Sigma_{tie}^{-1} \gamma_{ip} \lambda_i' \hat{A}' \hat{M}_{ff}] \) for consideration since these two terms are each \( O_p(T^{-1/2}) \) due to Proposition C.1, which violates the claim of this proposition. But we show that the sum of the 1st and 15th terms satisfies the proposition. By \( \hat{G} = \hat{H} - \hat{H} \hat{M}_{ff}^{-1} \hat{G} \), term \( \hat{A} = (\hat{\Gamma} - \Gamma) \Sigma_{ee}^{-1} \hat{\Gamma} \hat{G} \) is equal to \( (\hat{\Gamma} - \Gamma) \Sigma_{ee}^{-1} \hat{\Gamma} \hat{H} - (\hat{\Gamma} - \Gamma) \Sigma_{ee}^{-1} \hat{H} \hat{M}_{ff}^{-1} \hat{G} = A - A \hat{M}_{ff}^{-1} \hat{G} \). So the 15th term \( \text{tr}[\frac{1}{N} \sum_{i=1}^{N} \Sigma_{tie}^{-1} \gamma_{ip} \lambda_i' \hat{A}' \hat{M}_{ff}] \) is equal to

\[
\text{tr}\left[\frac{1}{N} \sum_{i=1}^{N} \Sigma_{tie}^{-1} \gamma_{ip} \lambda_i' \hat{A}' \hat{M}_{ff}\right] = \text{tr}\left[M_{ff} A \frac{1}{N} \sum_{i=1}^{N} \lambda_i \Sigma_{tie}^{-1} \gamma_{ip}\right] \\
= \text{tr}\left[M_{ff} A \frac{1}{N} \sum_{i=1}^{N} (\hat{\lambda}_i - \lambda_i) \Sigma_{tie}^{-1} \gamma_{ip}\right] + \text{tr}\left[M_{ff} \hat{A} \hat{v}_p\right] \\
= \text{tr}\left[M_{ff} A \frac{1}{N} \sum_{i=1}^{N} (\hat{\lambda}_i - \lambda_i) \Sigma_{tie}^{-1} \gamma_{ip}\right] + \text{tr}\left[M_{ff} A \hat{M}_{ff}^{-1} \hat{G} \hat{v}_p\right] \\
\]

So the sum of the 1st and 15th terms can be written as

\[
(C.5) \quad \text{tr}\left[\hat{G}_1 \sum_{i=1}^{N} \hat{\Gamma}_i \Sigma_{ii}^{-1} (\hat{\Gamma}_i - \Gamma_i) \hat{M}_{ff} \hat{v}_p^g\right] + \text{tr}\left[\frac{1}{N} \sum_{i=1}^{N} \Sigma_{tie}^{-1} \gamma_{ip} \lambda_i' \hat{A}' \hat{M}_{ff}\right] \\
= \text{tr}\left[I_1 A' \hat{M}_{ff} \hat{v}_p^g\right] + \text{tr}\left[I_1 M_{ff} \hat{A} \hat{v}_p\right] + \text{tr}\left[I_2 M_{ff} \hat{A} \hat{v}_p^h\right] \\
- \text{tr}\left[H_1 \hat{M}_{ff}^{-1} A' \hat{M}_{ff} \hat{v}_p^g\right] - \text{tr}\left[M_{ff} A \hat{M}_{ff}^{-1} \hat{G} \hat{v}_p\right] + \text{tr}\left[M_{ff} A \frac{1}{N} \sum_{j=1}^{N} (\hat{\lambda}_j - \lambda_j) \Sigma_{jie}^{-1} \gamma_{jp}\right] \\
\]

The first two terms can be written as \( \text{tr}[I_1 (A'M_{ff} + M_{ff} A) \hat{v}_p^g] \). Under the identification condition \( IZ \), the first \( r_1 \) rows of \( \hat{M}_{ff} - M_{ff} \) is zero. Thus the expression of \( I_1 (A'M_{ff} + M_{ff} A) \) is implicitly given in (A.21). The third term involves \( I_2 M_{ff} A \). Notice the last term of the left hand side of (C.2) is \( I_2 M_{ff} A \hat{\Pi}_{\lambda \lambda} \). Shifting \( I_2 M_{ff} A \hat{\Pi}_{\lambda \lambda} \) from the left to the right, then post-multiplying \( -\hat{\Pi}_{\lambda \lambda} \), we obtain the expression of \( I_2 M_{ff} A \). Substituting these two expressions into (C.5), we can rewrite the first three terms of (C.5). This allows us to rewrite (C.4) as

\[
(C.6) \quad \vartheta_{p1}(\hat{\beta}_1 - \beta_1) + \vartheta_{p2}(\hat{\beta}_2 - \beta_2) + \cdots + \vartheta_{pK}(\hat{\beta}_K - \beta_K) \\
= \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{tie}^{-1} v_{iit}^e + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{tie}^{-1} h_{iit} \epsilon_{it} \\
\]
\[- \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} h_{tj} \epsilon_{j\lambda} \hat{\Sigma}_{yj}^{-1} \lambda_{j} \hat{v}_{p}^{\lambda} \]

\[+ \text{tr} \left( \hat{G}_{1} \sum_{i=1}^{N} \hat{\Gamma}_{i} \hat{\Sigma}_{y}^{-1} \hat{\Gamma}_{i} \left( \frac{1}{T} \sum_{t=1}^{T} f_{t} (\beta - \beta') \hat{G}_{ip}^{\beta} \right) \right) \]

\[- \text{tr} \left( I_{1} (I - A)' \left( \frac{1}{T} \sum_{t=1}^{T} f_{t} (\beta - \beta') \hat{\epsilon}_{t} \right) \hat{H}_{N} \hat{v}_{p}^{\beta} \right) \]

\[+ \text{tr} \left( \hat{G}_{2} \hat{\Gamma} \hat{\Sigma}_{y}^{-1} \hat{\Gamma} \left( \frac{1}{T} \sum_{t=1}^{T} f_{t} (\beta - \beta') \hat{\epsilon}_{t} \right) \hat{H}_{N} \hat{v}_{p}^{\beta} \right) \]

where

\[S_{p1} = \text{tr} \left[ \hat{G}_{1} \hat{M}_{ff}^{-1} \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_{tj} \hat{\Sigma}_{yj}^{-1} \gamma_{j} \hat{v}_{p}^{\gamma} \right] + \text{tr} \left[ I_{1} A' \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} f_{tj} \hat{\Sigma}_{yj}^{-1} \gamma_{j} \hat{v}_{p}^{\gamma} \right] \]

\[- \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{y}^{-1} (\hat{\lambda}_{i} - \lambda_{i}) f_{t} \hat{v}_{itp} + \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{y}^{-1} f_{t} \hat{\lambda}_{i} \hat{v}_{itp} A' \right] \]

\[+ \text{tr} \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{y}^{-1} f_{t} \hat{\lambda}_{i} \hat{v}_{itp} G \hat{M}_{ff}^{-1} \right] - \text{tr} \left[ I_{1} A' \frac{1}{T} \left( \sum_{t=1}^{T} f_{t} \hat{\lambda}_{t} \right) \hat{H}_{N} \hat{v}_{p}^{\beta} \right] \]

\[- \text{tr} \left[ \hat{H}_{1} N \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\lambda}_{t} f_{t} \right) \hat{A} \hat{v}_{p}^{\beta} \right] + \text{tr} \left[ \hat{H}_{1} N \hat{M}_{ff}^{-1} \hat{G}_{N} \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\lambda}_{t} f_{t} \right) \hat{v}_{p}^{\beta} \right] \]

\[- \text{tr} \left[ \frac{1}{N} \sum_{j=1}^{N} \gamma_{jp} \hat{\Sigma}_{yj}^{-1} (\hat{\lambda}_{j} - \lambda_{j}) G_{N} \left( \frac{1}{T} \sum_{t=1}^{T} \hat{\lambda}_{t} f_{t} \right) \right] - \text{tr} \left[ \hat{H}_{1} \hat{M}_{ff}^{-1} \hat{A}' \hat{M}_{ff} \hat{v}_{p}^{\beta} \right] \]

\[+ \text{tr} \left[ \hat{H} \hat{M}_{ff}^{-1} \hat{G}_{N} \frac{1}{T} \left( \sum_{t=1}^{T} \hat{\lambda}_{t} f_{t} \right) \hat{v}_{p}^{\beta} \right] - \text{tr} \left[ M_{ff} \hat{A} \frac{1}{N} \sum_{j=1}^{N} (\hat{\lambda}_{j} - \lambda_{j}) \hat{\Sigma}_{yj}^{-1} \gamma_{jp} \right] \]

\[- \text{tr} \left[ I_{2} \hat{A}' \left( \frac{1}{T} \sum_{t=1}^{T} f_{t} \hat{\lambda}_{t} \right) \hat{H}_{N} \hat{v}_{p}^{\beta} \right] + \text{tr} \left[ M_{ff} \hat{A} \frac{1}{N} \sum_{j=1}^{N} (\hat{\lambda}_{j} - \lambda_{j}) \hat{\Sigma}_{yj}^{-1} \gamma_{jp} \right] \]

\[- \text{tr} \left[ G_{2} \hat{M}_{ff}^{-1} \left( \frac{1}{T} \sum_{t=1}^{T} f_{t} \hat{\lambda}_{t} \right) \hat{H}_{N} \hat{v}_{p}^{\beta} \right] + \text{tr} \left[ I_{1} A' \hat{M}_{ff} \hat{A} \hat{v}_{p}^{\beta} \right] \]
\[-\text{tr}\left[G_{2N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t f_t^i\right) A^h_p\right] + \text{tr}\left[\frac{1}{N} \hat{G}_1 \sum_{j=1}^{N} \hat{\lambda}_j \Sigma_{iie}^{-1}(\hat{\Sigma}_{iie} - \Sigma_{iie}) \Sigma_{iie}^{-1} \gamma^g_{jp}\right]\]

\[+ \text{tr}\left[\hat{H}_1 \hat{M}_{ff}^{-1} \hat{G} \hat{H}^{-1} \hat{M}_{ff} \frac{1}{N} \sum_{j=1}^{N} (\hat{\lambda}_j - \lambda_j) \Sigma_{iie}^{-1} \gamma^g_{jp}\right] - \text{tr}\left[\frac{1}{N} \sum_{i=1}^{N} \hat{\gamma}_{ip} (\hat{\gamma}_{ip} - \gamma_{ip}) \hat{\lambda}^i \hat{G}\right]\]

\[+ \text{tr}\left[\hat{H}_1 \hat{M}_{ff}^{-1} \hat{G} \hat{H}^{-1} (\hat{M}_{ff} - \hat{M}_{ff}) \check{\nu}^h_p\right] - \text{tr}\left[\frac{1}{N} \sum_{i=1}^{N} \hat{\Sigma}_{iie}^{-1} \hat{\lambda}^i \hat{G} \hat{M}_{ff}^{-1} (\hat{M}_{ff} - \hat{M}_{ff}) \check{\nu}^h_p\right]\]

\[+ \text{tr}\left[\check{I}_2 \check{A} \hat{M}_{ff} \hat{A}^h_p\right] + \text{tr}\left[G_2 \sum_{i=1}^{N} \hat{\Gamma}_1 \hat{\Sigma}_{iie}^{-1}(\Sigma_{iie} - \Sigma_{iie}) \hat{\Sigma}_{iie}^{-1} \hat{\Gamma}_1 \hat{H} \hat{v}^p\right]\]

\[+ \text{tr}\left[H_2 \hat{M}_{ff}^{-1} \hat{G} \check{\Sigma}_{ee} \hat{M}_{ff} \hat{A}^h_p\right] + \text{tr}\left[I_2 \check{A} \frac{1}{NT} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{f}_t e_{jt} \hat{\Sigma}_{iie}^{-1} \hat{\lambda}_j \hat{\Pi}_{\lambda \lambda}^{-1} \hat{v}^h_p\right]\]

\[+ \text{tr}\left[\frac{1}{N} \sum_{i=1}^{N} \hat{\Gamma}_1 \hat{\Sigma}_{iie}^{-1}(\Sigma_{iie} - \Sigma_{iie}) \hat{\Sigma}_{iie}^{-1} \hat{\Gamma}_1 \hat{H} \hat{v}^p\right]\]

\[= \text{tr}\left[G_{1N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta) f_t^i\right) \hat{v}^g_p\right] - \text{tr}\left[\hat{H}_{1N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta) f_t^i\right) (I - A) \hat{v}^g_p\right]\]

\[+ \text{tr}\left[G_{2N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta) f_t^i\right) A^h_p\right] + \text{tr}\left[G_{1N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta) \hat{G}^g_{ip}\right)\right]\]

\[+ \text{tr}\left[G_{1} \frac{1}{T} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta) e_{jt} \Sigma_{iie}^{-1} \gamma^g_{jp}\right] - \text{tr}\left[G_{1N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta)(\beta - \beta) \hat{G}^g_{ip}\right)\right]\]

\[-\text{tr}\left[H_{1N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta) \hat{\xi}_t^i\right) \hat{H}_N \hat{v}^g_p\right] - \text{tr}\left[H_{1N}\left(\frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta) \hat{\xi}_t^i\right) \hat{H}_N \hat{v}^g_p\right]\]
+tr \left[ \hat{H}_1 N \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta)' \hat{\xi}_t \hat{H}_N \check{v}_p^g \right] + tr \left[ \hat{G}_2 N \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta)' \hat{F}_t \hat{\Pi}_{\hat{\lambda}_t}^{-1} \check{v}_p^h \right] \\
+tr \left[ \hat{G}_2 N \frac{1}{T} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta)e_{jt} \Sigma_{jte}^{-1} \hat{\xi}_t \hat{\Pi}_{\hat{\lambda}_t}^{-1} \check{v}_p^h \right] - tr \left[ \hat{G}_2 N \frac{1}{T} \sum_{t=1}^{T} \hat{\chi}_t (\beta - \beta)' \hat{F}_t \hat{\Pi}_{\hat{\lambda}_t}^{-1} \check{v}_p^h \right] \\
- tr \left[ \hat{G}_2 N \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta)(\beta - \beta)' \hat{F}_t \hat{\Pi}_{\hat{\lambda}_t}^{-1} \check{v}_p^h \right] - tr \left[ \hat{G}_2 N \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta) \hat{H}_N \check{v}_p^h \right] \\
+tr \left[ \hat{G}_2 N \frac{1}{T} \sum_{t=1}^{T} \hat{\xi}_t (\beta - \beta)(\beta - \beta)' \hat{H}_N \check{v}_p^h \right]

and

\begin{align*}
S_{p3} &= - tr \left[ \hat{G}_1 N \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \frac{1}{T} \sum_{t=1}^{T} [\varepsilon_{it} e_{jt} - E(\varepsilon_{it} e_{jt})] \hat{\Sigma}_{jte}^{-1} \check{v}_p^g \right] \\
&- tr \left[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{jt} v_{ipt} - E(\varepsilon_{jt} v_{ipt})] \hat{\Sigma}_{iie}^{-1} \lambda \hat{G} \right] \\
&+ tr \left[ \hat{H}_1 N \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} e'_{jt} - E(\varepsilon_{it} e'_{jt})] \hat{\Sigma}_{jte}^{-1} \hat{\Gamma}_j \hat{H} \check{v}_p^g \right] \\
&- tr \left[ \hat{G}_2 N \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} e'_{jt} - E(\varepsilon_{it} e'_{jt})] \hat{\Sigma}_{jte}^{-1} \hat{\Gamma}_j \hat{H} \check{v}_p^h \right] \\
&+ tr \left[ \hat{G}_2 N \frac{1}{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T} \hat{\Gamma}_i \hat{\Sigma}_{ii}^{-1} \varepsilon_{it} e'_{jt} - E(\varepsilon_{it} e'_{jt})] \hat{\Sigma}_{jte}^{-1} \hat{\Gamma}_j \hat{H} \check{v}_p^h \right]
\end{align*}

Consider (C.6). Terms $S_{p1}, S_{p2}$ are dealt with in Lemma C.3. The terms on the left hand side and the 4th-7th terms on the right hand side are summarized in Lemma C.2. The first three terms on the right hand side and $S_{p3}$ are given in Lemma C.4. Using the results in Lemma C.2, C.3 and C.4, we have

\[ Q(\beta - \beta) = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1}) \]

This leads to Proposition C.3.

**Corollary C.1.** Under Assumptions A-D plus G, together with the identification conditions IO, we have

\[ \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii}^{-1} \| \cdot \| \hat{\Gamma}_i - \Gamma_i \|^2 = O_p(T^{-1}) \]
Under the assumptions of Theorem 4.1, we have

\[ \frac{1}{N} \sum_{i=1}^{N} \| \hat{\Sigma}_{ii} - \Sigma_{ii} \|^2 = O_p(T^{-1}) \]

\[ \| \hat{M}_{ff} - M_{ff} \|^2 = O_p(T^{-1}) \]

This corollary is an immediate result of Proposition C.2 and C.3.

To prove Proposition 4.1, we need the following lemma.

**Lemma C.5.** Under the assumptions of Theorem 4.1, we have

(a) \[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iie} h_{te} = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iie} h_{te} + O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(b) \[ \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\Sigma}_{iie} = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma_{iie} + O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(c) \[ \frac{1}{NT} \sum_{j=1}^{T} \sum_{t=1}^{T} h_{te} \hat{\Sigma}_{jte} = \frac{1}{NT} \sum_{j=1}^{T} \sum_{t=1}^{T} \Sigma_{jte} + O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(d) \[ \text{tr} \left[ \hat{G} - \frac{1}{2} E \left( \hat{\Sigma}_{iie}^2 \right) \right] = O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(e) \[ \text{tr} \left[ \hat{\Sigma}_{iie} - \frac{1}{2} E \left( \hat{\Sigma}_{iie}^2 \right) \right] = O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(f) \[ \text{tr} \left[ \hat{H} - \frac{1}{2} E \left( \hat{H}^2 \right) \right] = O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(g) \[ \text{tr} \left[ \hat{G} - \frac{1}{2} E \left( \hat{G}^2 \right) \right] = O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]

(h) \[ \text{tr} \left[ \hat{G} - \frac{1}{2} E \left( \hat{G}^2 \right) \right] = O_p(T^{-1/2}) + O_p(T^{-1}) + O_p(T^{-3/2}) + o_p(||\hat{\beta} - \beta||) \]
\[ = O_p(N^{-1/2}T^{-1/2} + O_p(N^{-1/2}T^{-1}) + O_p(T^{-3/2}) + o_p(\|\hat{\beta} - \beta\|) \]

**Proof of Lemma C.5.** Proof of Lemma C.5 is quite similar to that of Lemma A.13 and hence is omitted.

**Proofs of Proposition 4.1.** Given (C.6) and Lemmas C.2, C.3 and C.5, the proof of Proposition 4.1 is identical to that of Proposition 3.1. The details are omitted.

**APPENDIX C3: PROOFS OF THEOREMS 4.1 AND 4.3**

**Proof of Theorem 4.1.** To prove Theorem 4.1, we first transform the parameters set \((\Gamma, f_t, \beta)\) into \((\Gamma^*, f^*_t, \beta)\) which satisfies the identification condition IO. For ease of reading, we rewrite model (4.1) below

\[ y_{it} = \alpha_i + x_{it1}\beta_1 + x_{it2}\beta_2 + \cdots + x_{itK}\beta_K + \psi_i'g_{it} + \phi_i'h_{it} + \epsilon_{it} \]

\[ x_{itk} = \mu_{ik} + \gamma_{itk}^g g_{it} + \gamma_{itk}^h h_{it} + v_{itk} \]

The first equation of the above can be rewritten as

\[ y_{it} = (\alpha_i + \psi_i'\hat{g} + \phi_i'\hat{h}) + \sum_{k=1}^{K} x_{itk}\beta_k + \left\{ [\psi_i' + \phi_i'\hat{H}'\hat{G}'\hat{G}^{-1}(\hat{G}'\hat{G})^{-1/2}Q] \right\} \]

\[ \times \left\{ Q'(\hat{G}'\hat{G})^{-1/2}\hat{g}_t \right\} + \phi_i'\hat{H}_t - \hat{H}'\hat{G}'(\hat{G}'\hat{G})^{-1}\hat{g}_t \} + \epsilon_{it} \]

and the second equation can be rewritten as

\[ x_{itk} = (\mu_{ik} + \gamma_{itk}^g \hat{g} + \gamma_{itk}^h \hat{h}) + \left\{ [\gamma_{itk}^g + \gamma_{itk}^h \hat{H}'\hat{G}'(\hat{G}'\hat{G})^{-1}(\hat{G}'\hat{G})^{-1/2}Q] \right\} \]

\[ \times \left\{ Q'(\hat{G}'\hat{G})^{-1/2}\hat{g}_t \right\} + \gamma_{itk}^h \hat{h}_t - \hat{H}'\hat{G}'(\hat{G}'\hat{G})^{-1}\hat{g}_t \} + v_{itk} \]

where \(Q\) is an \(r_1 \times r_1\) orthogonal matrix which is defined below. Let

(C.7) \[ \psi_i'' = [\psi_i' + \phi_i'\hat{H}'\hat{G}'\hat{G}^{-1}(\hat{G}'\hat{G})^{-1/2}Q] \]

(C.8) \[ \gamma_{itk}'' = [\gamma_{itk}^g + \gamma_{itk}^h \hat{H}'\hat{G}'(\hat{G}'\hat{G})^{-1}(\hat{G}'\hat{G})^{-1/2}Q] \]

(C.9) \[ \gamma_{itk}'' = \gamma_{itk}^h \]

(C.10) \[ g_i^* = Q'(\hat{G}'\hat{G})^{-1/2}\hat{g}_t \]

(C.11) \[ h_i^* = \hat{h}_t - \hat{H}'\hat{G}'(\hat{G}'\hat{G})^{-1}\hat{g}_t \]

Let \(\Gamma^{g*}, \Gamma^{h*}, \Gamma^*\) be defined similarly as \(\Gamma^g, \Gamma^h, \Gamma\). If we choose \(Q\) to be the eigenvector matrix of \(\Gamma^g\Sigma_{\Gamma^g}^{-1}\Gamma^g\) with the associated eigenvalues in descending order, we can easily verify that the parameters set \((\Gamma^*, f^*_t, \beta)\) satisfy the
identification conditions IO. Then by Proposition 4.1, we have

\[ Q^*(\hat{\beta} - \beta) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{ite} \epsilon_{it} v_{it} + \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{ite} h^* \epsilon_{it} \]

\[ - \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} \Sigma^{-1}_{ite} \lambda^*_{i\lambda} \Pi^{-1}_{i\lambda} \left( \frac{1}{N} \sum_{j=1}^{N} \lambda^*_{j\lambda} \Sigma^{-1}_{jtx} h^* \right) h^* \epsilon_{it} \]

\[ + O_p(T^{-3/2}) + O_p(N^{-1} T^{-1/2}) + O_p(N^{-1/2} T^{-1}) \]

Substituting (C.7)–(C.11) into the above expression, we have the same asymptotic expression as stated in Theorem 4.1.

The proof of the alternative expression in Theorem 4.1 is the same with that of Theorem 3.1 and hence omitted.

**Corollary 4.2** is a consequence of the alternative expression of \( \hat{\beta} - \beta \).  

**Proof of Theorem 4.3.** To prove Theorem 4.3, notice the first equation of (4.7) can always be written as

\[ y_{it} = \sum_{k=1}^{K} x_{itk} \beta_{k} + \left\{ \left( \psi_i + \phi_i [H'M(D)G[G'M(D)G]^{-1}] [G'M(D)G]^{1/2} \right) \right\} \]

\[ \times \left\{ Q'[G'M(D)G]^{-1/2} (g_t - G'D'D)^{-1} d_t \right\} \]

\[ + \phi_i \left\{ h_t - H'H(D'D)^{-1} d_t - H'M(D)G[G'M(D)G]^{-1} (g_t - G'D'D)^{-1} d_t \right\} \]

\[ + \left\{ \kappa_i + \phi_i G'D(D'D)^{-1} + \phi_i H'H(D'D)^{-1} \right\} d_t + e_{it} \]

The second equation of (4.7) can always be written as

\[ x_{itk} = \left\{ \left( \gamma_{ik} + \gamma_{ik} H'H(D)G[G'M(D)G]^{-1} \right) \right\} \]

\[ \times \left\{ Q'[G'M(D)G]^{-1/2} (g_t - G'D(D')^{-1} d_t) \right\} \]

\[ + \gamma_{ik} \left\{ h_t - H'H(D')^{-1} d_t - H'M(D)G[G'M(D)G]^{-1} (g_t - G'D(D')^{-1} d_t) \right\} \]

\[ + \left\{ \gamma_{ik} + \gamma_{ik} G'D'(D')^{-1} + \gamma_{ik} H'H'(D')^{-1} \right\} d_t + v_{itk} \]

Let

\[ \psi_i^{st} = (\psi_i + \phi_i [H'M(D)G[G'M(D)G]^{-1}] [G'M(D)G]^{1/2} Q \]

\[ g_t^{st} = Q'[G'M(D)G]^{-1/2} (g_t - G'D(D')^{-1} d_t) \]

\[ \gamma_{ik}^{st} = (\gamma_{ik} + \gamma_{ik} [H'H(D)G[G'M(D)G]^{-1}] [G'M(D)G]^{1/2} R \]
Under Assumptions A-D, in combination with the parameters (\( \Gamma^* \), \( f^* \), \( \beta \)) satisfying the identification condition \( \text{IO} \). Using the same method in deriving Theorem 4.1, we can prove Theorem 4.3. The details are omitted.

\[ h_t^* = h_t - \mathbb{H}' \mathbb{D}'(\mathbb{D}' \mathbb{D})^{-1} dt - \mathbb{H}' \mathbb{M}(\mathbb{D}) \mathbb{G}[\mathbb{G}' \mathbb{M}(\mathbb{D}) \mathbb{G}]^{-1} (gt - \mathbb{G}' \mathbb{D}'(\mathbb{D}' \mathbb{D})^{-1} dt) \]

\[ \eta_t^* = \eta_t + \gamma_t g' \mathbb{D}'(\mathbb{D}' \mathbb{D})^{-1} + \gamma_t h'_t \mathbb{H}' \mathbb{D}(\mathbb{D}' \mathbb{D})^{-1} \]

\[ \gamma_{ik}^* = \gamma_t g' \mathbb{D}'(\mathbb{D}' \mathbb{D})^{-1} + \gamma_t h'_t \mathbb{H}' \mathbb{D}(\mathbb{D}' \mathbb{D})^{-1} \]

After appropriately choosing the orthogonal matrix \( Q \), we can make the parameters \( \Gamma^* \), \( f^* \), \( \beta \) satisfy the identification condition \( \text{IO'} \). Using the same method in deriving Theorem 4.1, we can prove Theorem 4.3. The details are omitted.

APPENDIX C4: PROOFS OF THEOREM 4.2 AND 4.4

The following proposition is useful to derive Theorem 4.2 and 4.4.

**Proposition C.4.** Under Assumptions A-D, in combination with the identification conditions IO and IO', we have

\[ A \equiv \sum_{i=1}^{N}(\hat{\Gamma}_i - \Gamma_i)^{-1} \hat{\Gamma}_i' \hat{H} = \text{O}_p(N^{-1/2}T^{-1/2}) + \text{O}_p(T^{-1}) \]

**Proof of Proposition C.4.** Using Corollary C.1 and (C.2), we can prove that \( A_{12} \) and \( A_{22} \) are both \( \text{O}_p(N^{-1/2}T^{-1/2}) + \text{O}_p(T^{-1}) \). However, the identification condition IO1 implies that \( \hat{M}_{gh} = M_{gh} = 0 \). Given this result, by (A.21), we have \( A_{12} = \text{O}_p(N^{-1/2}T^{-1/2}) + \text{O}_p(T^{-1}) \). Now we only need to prove \( A_{11} = \text{O}_p(N^{-1/2}T^{-1/2}) + \text{O}_p(T^{-1}) \). This result can be proved by the same method as proving \( (\hat{\Lambda} - \Lambda)^{-1} \hat{\Sigma}_{ee}^{-1} \hat{\Lambda} H = \text{O}_p(N^{-1/2}T^{-1/2}) + \text{O}_p(T^{-1}) \) under \( \text{IC3} \) in [10]. So Proposition C.4 follows.

Theorem 4.2 can be viewed as a special case of Theorem 4.4. We only focus on proof of Theorem 4.4.

**Proof of Theorem 4.4.** The formula to estimate \( f_t \) is

\[ \hat{f}_t = (\hat{\Lambda}' \hat{\Sigma}_{ee}^{-1} \hat{\Lambda})^{-1} \hat{\Lambda}' \hat{\Sigma}_{ee}^{-1} (\hat{\Psi}_t - \tilde{\Psi}_t) \]

where \( \hat{f}_t = (\hat{g}'_t, \hat{h}'_t) \), \( \hat{\Lambda} = (\hat{\Psi}, \hat{\Phi}) \) and \( \tilde{\Psi}_t \) is the \( t \) th column of the matrix \( Y \mathbb{M}(\mathbb{D}) \), \( \tilde{\Psi}_t \) is an \( N \times K \) matrix with its \( k \) th column equal to the \( t \) th column of the matrix \( X_k \mathbb{M}(\mathbb{D}) \). Given the above equation, we have

(C.12)

\[ \hat{h}_t = \left[ \hat{\Phi}' \hat{\Sigma}_{ee}^{-1/2} \mathbb{M}(\hat{\Sigma}_{ee}^{-1/2} \hat{\Psi}) \hat{\Sigma}_{ee}^{-1/2} \hat{\Phi} \right]^{-1} \left[ \Phi' \hat{\Sigma}_{ee}^{-1/2} \mathbb{M}(\hat{\Sigma}_{ee}^{-1/2} \hat{\Psi}) \hat{\Sigma}_{ee}^{-1/2} (\tilde{\Psi}_t - \tilde{\Psi}_t) \right] \]

The first equation of (4.7) can be written as

\[ Y = X_1 \beta_1 + \cdots + X_K \beta_K + \Psi \mathbb{G}' + \Phi \mathbb{H}' + K \mathbb{D}' + e \]
Post-multiplying $\mathcal{M}(\mathbb{D})$ on both sides, by $\mathcal{G}' \mathbb{D} = 0, \mathcal{H}' \mathbb{D} = 0$, we have

$$Y \mathcal{M}(\mathbb{D}) = X_1 \mathcal{M}(\mathbb{D})\beta_1 + \cdots + X_K \mathcal{M}(\mathbb{D})\beta_K + \Psi \mathcal{G}' + \Phi \mathcal{H}' + e \mathcal{M}(\mathbb{D})$$

So we have

$$\tilde{Y}_t = \tilde{X}_t \beta + \Psi g_t + \Phi h_t + e_t - \bar{e}_t$$

where $\bar{e}_t = (\sum_{s=1}^T e_{is} d_s')(\sum_{s=1}^T d_s d_s')^{-1} d_t$ and $\tilde{e}_t = (\bar{e}_1, \bar{e}_2, \ldots, \bar{e}_N)'$. Substituting the above equation into (C.12), we have

$$\sqrt{N}(\hat{h}_t - h_t) = -\hat{\lambda}^{-1} \left[ \frac{1}{N} \Phi' \hat{\Sigma}^{-1/2} \mathcal{M}(\hat{\Sigma}^{-1/2} \hat{\psi}) \hat{\Sigma}^{-1/2} \hat{\Sigma}_t \sqrt{N}(\hat{\beta} - \beta) \right]$$

(C.13)

$$-\hat{\lambda}^{-1} \left[ \sqrt{N} \frac{1}{N} \Phi' \hat{\Sigma}^{-1/2} \mathcal{M}(\hat{\Sigma}^{-1/2} \hat{\psi}) \hat{\Sigma}^{-1/2} \hat{\Sigma}_t \hat{\theta}_t \right]$$

$$+ \hat{\lambda}^{-1} \left[ \sqrt{N} \frac{1}{N} \Phi' \hat{\Sigma}^{-1/2} \mathcal{M}(\hat{\Sigma}^{-1/2} \hat{\psi}) \hat{\Sigma}^{-1/2} (\epsilon_t - \bar{e}_t) \right]$$

where $\hat{\lambda} = \frac{1}{N} \Phi' \hat{\Sigma}^{-1/2} \mathcal{M}(\hat{\Sigma}^{-1/2} \hat{\psi}) \hat{\Sigma}^{-1/2} \Phi$. Using the consistency result we have proved, term $\frac{1}{N} \Phi' \hat{\Sigma}^{-1/2} \mathcal{M}(\hat{\Sigma}^{-1/2} \hat{\psi}) \hat{\Sigma}^{-1/2} \Phi$ can be showed to be

$$\frac{1}{N} \sum_{i=1}^N \phi_i \Sigma^{-1} \phi_i' - \left( \frac{1}{N} \sum_{i=1}^N \phi_i \Sigma^{-1} \psi_i' \right) \left( \frac{1}{N} \sum_{i=1}^N \psi_i \Sigma^{-1} \psi_i' \right)^{-1} \left( \frac{1}{N} \sum_{i=1}^N \psi_i \Sigma^{-1} \phi_i' \right) + O_p(1)$$

Consider the first term on the right hand side of (C.13). Term $\frac{1}{N} \Phi' \hat{\Sigma}^{-1/2} \mathcal{M}(\hat{\Sigma}^{-1/2} \hat{\psi}) \hat{\Sigma}^{-1/2} \hat{\Sigma}_t$ is equal to

$$\frac{1}{N} \sum_{i=1}^N \phi_i \Sigma^{-1} \xi_{it} - \left( \frac{1}{N} \sum_{i=1}^N \phi_i \Sigma^{-1} \psi_i' \right) \left( \frac{1}{N} \sum_{i=1}^N \psi_i \Sigma^{-1} \psi_i' \right)^{-1} \left( \frac{1}{N} \sum_{i=1}^N \psi_i \Sigma^{-1} \phi_i' \right)$$

where $\xi_{it}$ is the ith row of matrix $\hat{\Sigma}_t$. By the consistency result we have proved, the above expression is equal to

$$\frac{1}{N} \sum_{i=1}^N \phi_i \Sigma^{-1} \xi_{it} - \left( \frac{1}{N} \sum_{i=1}^N \phi_i \Sigma^{-1} \psi_i' \right) \left( \frac{1}{N} \sum_{i=1}^N \psi_i \Sigma^{-1} \psi_i' \right)^{-1} \left( \frac{1}{N} \sum_{i=1}^N \psi_i \Sigma^{-1} \phi_i' \right) + O_p(1)$$

which is $O_p(1)$. From this, in combination with $\hat{\beta} - \beta = O_p(N^{-1/2} T^{-1/2}) + O_p(T^{-3/2})$, the first term of (C.13) is equal to $O_p(T^{-1/2}) + O_p(N^{1/2} T^{-3/2})$. Now consider the term $\frac{1}{N} \Phi' \hat{\Sigma}^{-1/2} \mathcal{M}(\hat{\Sigma}^{-1/2} \hat{\psi}) \hat{\Sigma}^{-1/2} (\hat{\psi} - \psi) g_t$, which is equal to

$$\frac{1}{N} \sum_{i=1}^N \phi_i \Sigma^{-1} (\hat{\psi}_i - \psi_i)'$$
So we have

\[-\left(\frac{1}{N}\sum_{i=1}^{N} \phi_i \hat{\Sigma}_{ii}^{-1} \hat{\psi}_i'\right) \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \hat{\Sigma}_{ii}^{-1} \hat{\psi}_i'\right)^{-1} \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \hat{\Sigma}_{ii}^{-1} (\hat{\psi}_i - \psi_i)'\right) g_t\]

Using (C.3) to replace \(\hat{\psi}_i - \psi_i\) from the above expression and using the result in Proposition C.4, we can show that

\[\frac{1}{N}\sum_{i=1}^{N} \phi_i \hat{\Sigma}_{ii}^{-1} (\hat{\psi}_i - \psi_i)' = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1})\]

\[\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \hat{\Sigma}_{ii}^{-1} (\hat{\psi}_i - \psi_i)' = O_p(N^{-1/2}T^{-1/2}) + O_p(T^{-1})\]

From these, together with \(\frac{1}{N}\sum_{i=1}^{N} \phi_i \hat{\Sigma}_{ii}^{-1} \hat{\psi}_i' = \frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} \psi_i' + o_p(1)\) and \(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \hat{\Sigma}_{ii}^{-1} \hat{\psi}_i' = \frac{1}{N}\sum_{i=1}^{N} \psi_i \Sigma_{ii}^{-1} \psi_i' + o_p(1)\), we have that the second term of the right hand side of (C.13) is \(O_p(T^{-1/2}) + O_p(N^{1/2}T^{-1})\).

Now consider the third term. We first consider \(\frac{1}{N}\Phi \hat{\Sigma}_{ee}^{-1/2} M(\hat{\Sigma}_{ee}^{-1/2} \hat{\psi}) \hat{\Sigma}_{ee}^{-1/2} e_t\), which is equal to

\[\frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} e_t - \left(\frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} \hat{\psi}_i'\right) - \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \Sigma_{ii}^{-1} \hat{\psi}_i'\right)^{-1} \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \Sigma_{ii}^{-1} e_t\right)\]

Using the expression for \(\hat{\Sigma}_{jje} - \Sigma_{jje}\), we can show that

\[\frac{1}{N}\sum_{i=1}^{N} \phi_i \hat{\Sigma}_{ii}^{-1} e_t = \frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} e_t + O_p(T^{-1})\]

and

\[\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \hat{\Sigma}_{ii}^{-1} e_t = \frac{1}{N}\sum_{i=1}^{N} \psi_i \Sigma_{ii}^{-1} e_t + O_p(T^{-1})\]

Furthermore, Using the expression for \(\hat{\psi}_i - \psi_i\), we can show

\[\frac{1}{N}\sum_{i=1}^{N} (\hat{\psi}_i - \psi_i) \Sigma_{ii}^{-1} e_t = O_p(N^{-1}) + O_p(T^{-1})\]

So we have

\[\frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} e_t - \left(\frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} \hat{\psi}_i'\right) - \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \Sigma_{ii}^{-1} \hat{\psi}_i'\right)^{-1} \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \Sigma_{ii}^{-1} e_t\right)\]

\[= \frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} e_t - \left(\frac{1}{N}\sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} \hat{\psi}_i'\right) - \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \Sigma_{ii}^{-1} \hat{\psi}_i'\right)^{-1} \left(\frac{1}{N}\sum_{i=1}^{N} \hat{\psi}_i \Sigma_{ii}^{-1} e_t\right)\]
This completes the proof of the second part of Theorem 4.4.

\[ + O_p(N^{-1}) + O_p(T^{-1}) \]

Notice

\[
\frac{1}{N} \sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} \phi_i' - \left( \frac{1}{N} \sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} \psi_i' \right) \left( \frac{1}{N} \sum_{i=1}^{N} \psi_i \Sigma_{ii}^{-1} \psi_i' \right)^{-1} \left( \frac{1}{N} \sum_{i=1}^{N} \psi_i \Sigma_{ii}^{-1} \phi_i' \right)
\]

is equal to \( \frac{1}{N} \Phi' \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi \). Then we have

\[
\sqrt{N}(\hat{h}_t - h_t) = \left( \frac{1}{N} \Phi' \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi \right)^{-1} \times \left( \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} e_{it} - \left( \frac{1}{N} \sum_{i=1}^{N} \phi_i \Sigma_{ii}^{-1} \psi_i' \right) \left( \frac{1}{N} \sum_{i=1}^{N} \psi_i \Sigma_{ii}^{-1} \psi_i' \right)^{-1} \left( \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \psi_i \Sigma_{ii}^{-1} e_{it} \right) \right)
\]

\[ + O_p(N^{-1/2}) + O_p(T^{-1/2}) + O_p(N^{1/2}T^{-1}) \]

If \( N \to \infty, T \to \infty \), and \( \sqrt{N}/T \to 0 \), we have

\[
\sqrt{N}(\hat{h}_t - h_t) \xrightarrow{d} N \left( 0, \lim_{N \to \infty} \frac{1}{N} \Phi' \Sigma_{ee}^{-1/2} \mathcal{M}(\Sigma_{ee}^{-1/2} \Psi) \Sigma_{ee}^{-1/2} \Phi \right)^{-1} \]

This proves the first part of Theorem 4.4.

By \( y_{it} = \sum_{p=1}^{K} x_{itp} \beta_p + \psi_i' g_{it} + \phi_i' h_{it} + \kappa_i d_{it} + e_{it} \), we have

\[
\sqrt{T}(\hat{\kappa}_i - \kappa_i) = - \sum_{p=1}^{K} \left( \frac{1}{T} \sum_{t=1}^{T} d_t d_t' \right)^{-1} \left( \frac{1}{T} \sum_{t=1}^{T} d_t x_{itp} \right) \sqrt{T}(\beta_p - \beta_p)
\]

\[- \left( \frac{1}{T} \sum_{t=1}^{T} d_t d_t' \right)^{-1} \left( \sqrt{T} \frac{1}{T} \sum_{t=1}^{T} d_t \lambda_i'(f_t - f_t) \right) \]

\[- \left( \frac{1}{T} \sum_{t=1}^{T} d_t d_t' \right)^{-1} \left( \sqrt{T} \frac{1}{T} \sum_{t=1}^{T} d_t (\hat{\psi}_i - \psi_i)' \hat{g}_t \right) + \left( \frac{1}{T} \sum_{t=1}^{T} d_t d_t' \right)^{-1} \left( \frac{1}{\sqrt{T}} \sum_{t=1}^{T} d_t e_{it} \right) \]

The first term is \( O_p(N^{-1/2}) + O_p(T^{-1}) \). The second term and the third term can be proved to be \( O_p(T^{1/2}N^{-1}) + O_p(T^{-1/2}) \). So we have

\[
\sqrt{T}(\hat{\kappa}_i - \kappa_i) = \left( \frac{1}{T} \sum_{t=1}^{T} d_t d_t' \right)^{-1} \left( \frac{1}{\sqrt{T}} \sum_{t=1}^{T} d_t e_{it} \right) + O_p(T^{1/2}N^{-1}) + O_p(T^{-1/2})
\]

If \( N \to \infty, T \to \infty \), and \( \sqrt{T}/N \to 0 \), we have

\[
\sqrt{T}(\hat{\kappa}_i - \kappa_i) \xrightarrow{d} N \left( 0, \Sigma_{ii}(\lim_{T \to \infty} \frac{1}{T} \sum_{t=1}^{T} d_t d_t'^{-1}) \right) \]

This completes the proof of the second part of Theorem 4.4. \( \square \)