

Long run rates and monetary policy

Bayesian Analysis and Modeling (BAM) Summer Workshop 2018

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¹Views expressed here are not those of the ECB or of the FRB

Motivation

- ⇒ "Movements in the [...] yield spread are associated with **movements in risk**" (Atkeson and Kehoe, 2010; Cochrane, 2010)
- In the conventional view, the short rate drops at the beginning of a recession, but it is expected to return the steady state within at least 10 years.

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- In the conventional view, the short rate drops at the beginning of a recession, but it is expected to return the steady state within at least 10 years.
 - In fact, taking account of risk premia, 10 year expected interest rates fall just as fast as the 1 year rate

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- ... and for inflation expectations?

Our paper

- A single model-feature can reconcile the macro and the finance literature: **heteroskedasticity** (in the form of regime switching)
 - Uncertainty shocks also amount to variation in risk: during recessions volatility drives the increase in risk premia. Risk premia are **countercyclical**—as in the finance literature

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 - Uncertainty shocks also amount to variation in risk: during recessions volatility drives the increase in risk premia. Risk premia are **countercyclical**—as in the finance literature
 - "Uncertainty shocks" change precautionary saving: during recessions volatility increases and real rates fall. Nominal 10 year expected interest rates fall together with policy rates—as "observed" in the data

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 - Risk-neutrality (EH holding) an artifact of linearization: we analyse the nonlinear solution of a DSGE model
 - We estimate the nonlinear model on both macro and yields data for the U.S.
 - We show that the model fits both sets of data reasonably well

Literature

- On heteroskedastic shocks in macroeconomic—Sims-Zha (2006), Primiceri (2005), Justiniano-Primiceri (2008) ...

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- Few empirical applications in nonlinear models–van Bindsberger *et al.*(2012), Andreasen (2012) ...

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- Resource constraint

$$Y_t = C_t + G_t + \frac{\zeta}{2} \left(\pi_t - (\pi^*)^{1-\iota} \pi_{t-1}^\iota \right)^2 Y_t$$

The model

- Policy rule

$$i_t = \text{const.} + \psi_{\pi} (\pi_t - \pi^*) + \psi_Y (\tilde{y}_t - \tilde{y}) + \rho_I i_{t-1} + \eta_{t+1}$$

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- Note: constant target π^*

Distinguishing feature: heteroskedasticity

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- Shocks: productivity (stationary and integrated), gov. spending, mark-up, policy
- Two-state, independent Markov switching in the innovation variances:

$$\varepsilon_{i,t+1} \sim N(0, \sigma_{i,s_{i,t}}) \quad \text{for } i = z, G, \eta$$

$$\sigma_{i,s_{i,t}} = \sigma_{i,0}s_{i,t} + \sigma_{i,1}(1 - s_{i,t})$$

with constant transition probabilities

$$p(s_{i,t+1} = k, s_{i,t} = j) = p_{i,jk}$$

Distinguishing feature: preferences

- Epstein-Zin-Weil preferences

$$U [u_t, (\mathbb{E}_t V_{t+1}^{1-\gamma})] = \left\{ (1 - \beta) u_t^{1-\psi} + \beta (\mathbb{E}_t V_{t+1}^{1-\gamma})^{\frac{1-\psi}{1-\gamma}} \right\}^{\frac{1}{1-\psi}}$$

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- γ = risk aversion, ψ = inverse of EIS
- Temporary utility with Trabandt and Uhlig (2011) specification

$$u = (C_t - h\Xi_t C_{t-1}) \left(1 - \eta (1 - \psi) N_t^{1+\frac{1}{\phi}} \right)^{\frac{\psi}{1-\psi}}$$

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- Habits
 - Have first order effects (hump shaped IRFs). High risk aversion makes consumption insensitive to real rate
- Recursive preferences
 - Have no effects to first order – dynamics as in a model with EU. Risk aversion parameter "free" to match yields.

Solution I

- As usual

$$E_t [f \{ \mathbf{x}_{t+1}, \mathbf{y}_{t+1}, \mathbf{x}_t, \mathbf{y}_t, ; s_{t+1}, s_t \}] = \mathbf{0}$$

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- We seek solutions of the form (Amisano and Tristani, JEDC 2011—a special case of recent Foerster *et al.*, 2016)

$$f(\mathbf{x}_t, \sigma; s_t) = f(\bar{\mathbf{x}}; 0; s_t) + \mathbf{F}_{s_t} (\mathbf{x}_t - \bar{\mathbf{x}}_{s_t}) + \frac{1}{2} (\mathbf{I}_{n_y} \otimes (\mathbf{x}_t - \bar{\mathbf{x}}_{s_t})') \mathbf{E}_{s_t} (\mathbf{x}_t - \bar{\mathbf{x}}_{s_t}) + \mathbf{k}_{y,s_t} \sigma^2$$

Solution II

- Only impact of heteroskedasticity in constant term

$$\hat{y}_t = F\hat{\mathbf{x}}_t + \frac{1}{2} (I_{n_y} \otimes \hat{\mathbf{x}}_t') E\hat{\mathbf{x}}_t + k_{y,s_t}$$

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- Similarly for predetermined variables

Estimation I

- Model is nonlinear

$$\mathbf{y}_{t+1}^o = \mathbf{k}_{y,j} + F\hat{\mathbf{x}}_{t+1} + \frac{1}{2} \left(I_{n_y} \otimes \hat{\mathbf{x}}'_{t+1} \right) E\hat{\mathbf{x}}_{t+1} + D\mathbf{v}_{t+1}$$

$$\mathbf{x}_{t+1} = \mathbf{k}_{x,i} + P\hat{\mathbf{x}}_t + \frac{1}{2} \left(I_{n_x} \otimes \hat{\mathbf{x}}'_t \right) G\hat{\mathbf{x}}_t + \tilde{\sigma}\Sigma_i\mathbf{w}_{t+1}$$

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- but main source of nonlinearity are intercept shifts. Hence extended Kalman filter

$$\mathbf{y}_{t+1}^o = \tilde{k}_{y,t+1}^{(i,j)} + \tilde{F}_{t+1}^{(i,j)}\hat{\mathbf{x}}_{t+1} + D\mathbf{v}_{t+1}$$

$$\hat{\mathbf{x}}_{t+1} = \tilde{k}_{x,t}^{(i)} + \tilde{P}_t^{(i)}\hat{\mathbf{x}}_t + \Sigma_i\mathbf{w}_{t+1}$$

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- Combine the likelihood with a prior and sample using a tuned Metropolis-Hastings algorithm
- tried unscented KF and particle filter without changes in the results

Data

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- Six observables: real per-capita GDP; real personal per-capita consumption; consumption deflator; 3-month nominal rate; 3-year and 10-year zero-coupon yields

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- Six observables: real per-capita GDP; real personal per-capita consumption; consumption deflator; 3-month nominal rate; 3-year and 10-year zero-coupon yields
- "Measurement errors" on all variables

Parameter estimates

- Monetary policy rule:

$$\hat{i}_t = 0.09 [3.09 (\pi_t - \pi^*) + 0.57 (\tilde{y}_t - \tilde{y})] + 0.91 \hat{i}_{t-1} + \eta_{t+1}.$$

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- High inertia

Parameter estimates

| | post mean | post sd | prior mean | prior sd |
|------------|----------------|---------|------------|----------|
| Π | 1.0061 | 0.0007 | 1.0062 | 0.0007 |
| ψ_π | 0.2676 | 0.0241 | 0.1990 | 0.1001 |
| ψ_y | 0.0497 | 0.0075 | 0.0200 | 0.0010 |
| ρ_i | 0.9135 | 0.0169 | 0.8494 | 0.1002 |
| Ξ | 1.0045 | 0.0004 | 1.0050 | 0.0010 |
| ι | 0.7333 | 0.1116 | 0.5003 | 0.1899 |
| ϕ | 0.6156 | 0.0846 | 1.0022 | 0.5049 |
| γ | 11.5185 | 3.6747 | 10.9537 | 6.9730 |
| ψ | 1.3075 | 0.0868 | 1.2035 | 0.2830 |
| ζ | 33.8071 | 3.1344 | 14.9744 | 6.9819 |
| h | 0.8619 | 0.0261 | 0.4996 | 0.1886 |
| β | 0.9984 | 0.0006 | 0.9986 | 0.0014 |

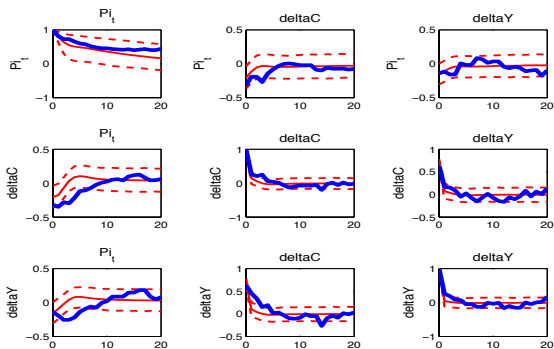
Parameter estimates

| | post mean | post sd | prior mean | prior sd |
|------------------|-----------|---------|------------|----------|
| $\rho_{G,11}$ | 0.8760 | 0.0556 | 0.8997 | 0.0657 |
| $\rho_{G,00}$ | 0.9413 | 0.0351 | 0.8994 | 0.0662 |
| $\rho_{\eta,11}$ | 0.9595 | 0.0196 | 0.8996 | 0.0657 |
| $\rho_{\eta,00}$ | 0.9079 | 0.0447 | 0.8998 | 0.0658 |
| $\rho_{z,11}$ | 0.9728 | 0.0091 | 0.9013 | 0.0651 |
| $\rho_{z,00}$ | 0.9317 | 0.0190 | 0.8993 | 0.0662 |
| ρ_{μ} | 0.5487 | 0.0581 | 0.8552 | 0.0916 |
| ρ_z | 0.9889 | 0.0018 | 0.8582 | 0.0899 |
| ρ_G | 0.9091 | 0.0298 | 0.8559 | 0.0906 |

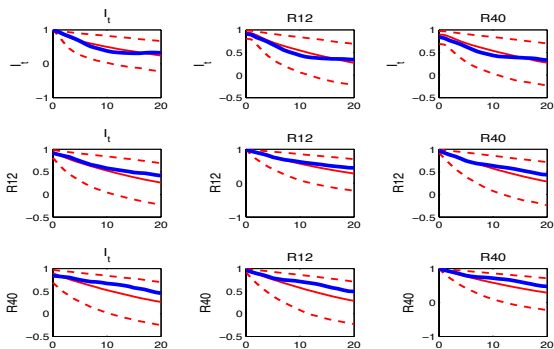
Parameter estimates

| | post mean | post sd | prior mean | prior sd |
|------------------------|-----------|---------|------------|----------|
| $\sigma_{me,\pi}$ | 1.4E-06 | 1.6E-06 | 1.4E-06 | 1.3E-06 |
| $\sigma_{me,\Delta c}$ | 1.3E-06 | 6.8E-07 | 1.4E-06 | 1.1E-06 |
| $\sigma_{me,\Delta y}$ | 0.0036 | 0.0006 | 0.0005 | 0.0003 |
| $\sigma_{me,i}$ | 1.3E-06 | 7.5E-07 | 1.4E-06 | 1.0E-06 |
| $\sigma_{me,i12}$ | 0.0007 | 7.6E-05 | 0.0014 | 0.0010 |
| $\sigma_{me,i40}$ | 0.0004 | 5.0E-05 | 0.0014 | 0.0010 |

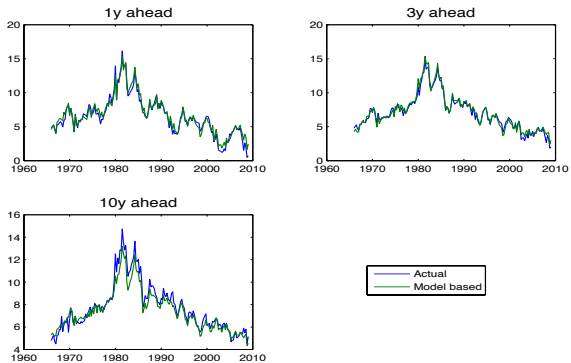
Dynamic correlations: macro variables



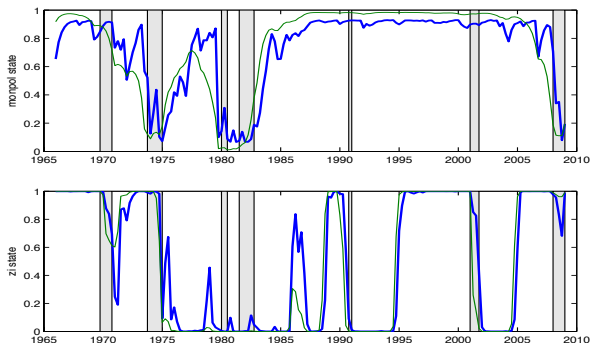
Dynamic correlations: yields



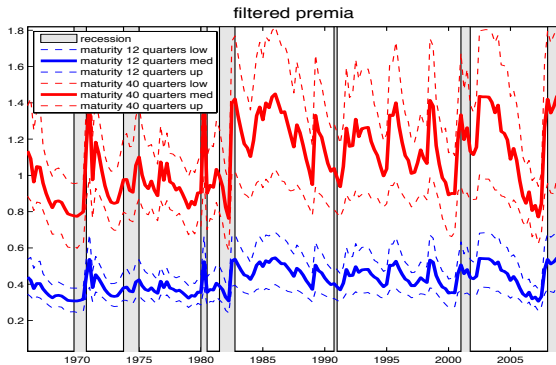
Forward rates



Probability of low-variance regimes



Expected excess holding period returns



Long-term rates over the business cycle

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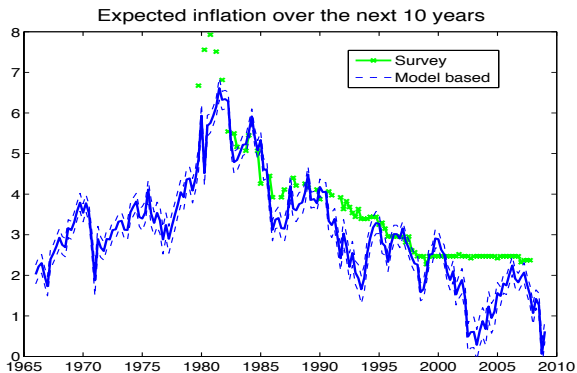
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- Indeed, $Ei \downarrow$ because demand for precautionary saving \uparrow , consumption \downarrow and adds \downarrow pressure on y and π
- After recession "confidence" returns. Uncertainty dynamics are reversed. It becomes clear that i will rise quickly. Risk premia \downarrow and forward rates become closer to Ei

Narrative: expectations



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- A sequence of highly persistent, adverse shocks led to an increase in trend inflation in the 1970s. The shocks were slowly reabsorbed over the 1980s. Long-term inflation expectations moved accordingly
- Inflation was never conquered. Prolonged deviations of inflation from price stability can happen again

Conclusions

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- Estimated model to account for key features of the transmission of monetary policy to long-term rates. Uncertainty/volatility shocks are important to explain observed variations in yields
- In the early parts of recessions, forward spreads are high because uncertainty and risk premia \uparrow not due to Ei . When recession ends, uncertainty and risk premia fall, and Ei rise; changes in forward rate reflect expected future interest rates.

Conclusions (II)

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- Movements in risk affecting spreads are not caused by monetary policy actions. But monetary policy responds to changes in risk, because of changes in precautionary saving
- Changes in real interest rates and in risk premia are important determinants of long term rates
- 10-year inflation expectations are less firmly anchored than one would conclude, based on survey data